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Firm heterogeneity, capital
misallocation and optimal
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Abstract

This paper analyzes the link between monetary policy and capital misallocation in a New Keynesian model with heterogeneous firms and financial frictions. In the model, firms with a high return to capital increase their investment more strongly in response to a monetary policy expansion, thus reducing misallocation. This feature creates a new time-inconsistent incentive for the central bank to engineer an unexpected monetary expansion to temporarily reduce misallocation. However, price stability is the optimal timeless response to demand, financial or TFP shocks. Finally, we present firm-level evidence supporting the theoretical mechanism.

Keywords: Monetary policy, firm heterogeneity, financial frictions, capital misallocation.

JEL classification: E12, E22, E43, E52, L11.

Non-technical summary

An inefficient distribution of capital among firms can significantly diminish overall total factor productivity (TFP). The distribution of capital is determined by the investment choices made by individual firms. The impact of monetary policy on these decisions is substantial. If firm investment responds heterogeneously to changes in monetary policy, then monetary policy can affect capital misallocation and thus TFP. Traditionally, the design of monetary policy has considered aggregate productivity as fixed. But if monetary policy can influence misallocation and TFP, then the central bank should contemplate the broader implications of its decisions on the supply side of the economy. This is especially the case in periods of active monetary policy, such as the current inflationary environment. The aim of this paper is to explore the relationship between monetary policy and capital misallocation, and its implications for optimal monetary policy.

To this end, we propose a model that combines the workhorse New Keynesian model with a model of firm heterogeneity, where financial frictions lead to a suboptimal allocation of capital. The economy is populated by a continuum of firms owned by entrepreneurs. Entrepreneurs are heterogeneous in their net worth and productivity. They face a borrowing constraint, which causes capital to be misallocated. This misallocation arises because the borrowing constraint binds for the most productive entrepreneurs: Their marginal return on capital exceeds the cost of capital and they do not employ as much capital as desirable. Instead, less productive entrepreneurs with a lower marginal return on capital employ some capital as well. This economy resembles the standard New Keynesian model with capital, except that in this case TFP is endogenous. The dynamics of TFP are determined by the evolution of the distribution of capital among entrepreneurs.

We show that the allocation of capital, and hence TFP, improve in the medium run in response to an expansionary monetary policy shock. This is because the policy expansion alleviates financial frictions, such that high-return firms increase their investment more than the low-return firms. As a result the share of the aggregate capital stock used by high-return firms increases and thus TFP goes up. We call this effect the capital misallocation channel of monetary policy.

Turning to optimal monetary policy, we show that the misallocation channel introduces a new source of *time inconsistency*. In the short run, the central bank is tempted

to exploit the capital misallocation channel. It wants to lower the interest rates temporarily, creating an expansionary policy shock, which raises TFP as just explained. But in the long run such a policy is undesirable because its benign effects crucially rely on this policy being unanticipated. We find this source of time inconsistency to be quantitatively more relevant than the standard time-inconsistency motive discussed in the New Keynesian literature, namely the desire to exploit the short-run trade-off between inflation and output gap.

We then analyze optimal monetary policy from a 'timeless perspective'. That is we consider a central bank that sticks to an optimal policy rule and does not give in to time-inconsistent temptations of the type just described. The optimal response to demand, financial and TFP shocks always features price stability. There is thus no meaningful trade-off between price stability and managing misallocation, just as the standard New Keynesian model features no trade-off between price stability and aggregate demand management (this is commonly known as the "divine coincidence"). This implies that, even if the central bank can affect the supply side of the economy through its monetary policy, it finds optimal *not* to do so in a systematic way and it rather sticks to price stability. Strict inflation targeting thus is optimal, as it would be in the absence of capital misallocation. Implementing such a strict inflation targeting policy, however, requires larger swings in the policy rate, if monetary policy affects misallocation. Thus, the inability to set negative policy rates, (the so-called zero lower bound) becomes more of a problem. Under optimal policy, the policymaker thus has to commit to a longer period of zero rates than he would, if monetary policy had no impact on misallocation.

We conclude by presenting empirical evidence on the capital misallocation channel of monetary policy. We combine panel data on the quasi-universe of Spanish firms with identified monetary policy shocks. We show that the mechanism at play in the model is supported by the data: firms with high marginal return on capital increase their investment relatively more than low-return firms after an expansionary monetary policy shock. This implies a shift in the distribution of capital towards more productive firms, improving the capital allocation. Using a simple model-derived measure of aggregate misallocation, we find that the positive impact of expansionary monetary policy is quantitatively in line with the data.

1 Introduction

An inefficient allocation of capital across firms may significantly reduce aggregate total factor productivity (TFP). In a market economy, capital is allocated according to the investment decisions of individual firms. Monetary policy is an important driver of these investment decisions, as it directly affects firms' funding costs and indirectly influences firms' revenues and other costs such as wages. If firm investment responds heterogeneously to changes in monetary policy, then monetary policy can affect capital misallocation and thus TFP.

Monetary policy design has traditionally taken aggregate productivity as given. In the workhorse model of monetary policy – the New Keynesian model – the central bank faces a trade-off between stabilizing inflation and reducing the short-term deviations of output from its potential level. If monetary policy can affect misallocation and TFP, the central bank should also ponder how its decisions will impact the supply side of the economy in the medium term. Such considerations may be of particular relevance in phases of very active monetary policy, such as in the current inflationary environment.

The objective of this paper is to shed light on the interaction between monetary policy and capital misallocation and its implications for optimal monetary policy. To this end, we develop a tractable framework that combines the workhorse New Keynesian model with a model of firm heterogeneity, based on [Moll \(2014\)](#), in which capital misallocation arises from financial frictions. The economy is populated by a continuum of firms owned by entrepreneurs, who have access to a constant-returns-to-scale technology. Entrepreneurs are heterogeneous in their net worth and receive idiosyncratic productivity shocks. They face financial frictions, as their borrowing cannot exceed a multiple of their net worth. This results in endogenous capital misallocation: entrepreneurs with productivity above a certain threshold are constrained, and borrow as much capital as possible since their marginal revenue product of capital (MRPK) is higher than their cost of capital. Entrepreneurs below the threshold are unconstrained: their optimal size is zero and they choose to lend their net worth to other entrepreneurs.¹ This economy allows for an aggregate representation akin to that in the standard New Keynesian model with capital, except that in this case TFP is endogenous, and the dynamics of TFP are determined by the evolution of the distribution of capital among

¹This is the tractable limit case of an economy with decreasing returns to scale at the firm level, in which unconstrained firms are optimally very small and the bulk of production is carried out by productive and constrained firms.

entrepreneurs. We calibrate the model to replicate key firm-level moments of Spanish firms.

We start by analyzing the interactions between monetary policy and misallocation. We show that the allocation of capital, and hence TFP, improve in response to an expansionary monetary policy shock. This is because the policy expansion alleviates financial frictions, such that high-MRPK firms increase their investment more than the low-MRPK firms, raising the share of the aggregate capital stock used by high-MRPK firms. We call this the *capital misallocation channel* of monetary policy.

Next, we investigate the implications of this capital misallocation channel for the optimal conduct of monetary policy by analyzing the Ramsey problem of a benevolent central bank. This is technical challenge, as the space state of the model includes the distribution of net worth across firms, an infinite-dimensional object. We overcome it by introducing a new algorithm to compute optimal policy problems in the presence of non-trivial heterogeneity.

We study first optimal policy in the absence of shocks. The steady state of the Ramsey problem features zero-inflation, as in the standard complete markets New Keynesian model, which is nested.² However, the misallocation channel introduces a new source of *time inconsistency* as the central bank attempts to exploit the capital misallocation channel: When starting from the zero-inflation steady state, the central bank engineers a temporary monetary expansion in the short run while committing to price stability in the long run. This strategy allows the central bank to temporarily improve capital allocation and increase TFP, even if it means tolerating positive inflation during a certain period of time. We find this source of time inconsistency to be quantitatively more relevant than the standard time-inconsistency motive in the New Keynesian literature, namely the desire to exploit the short-run trade-off between inflation and output gap.

We analyze next the optimal monetary policy from a 'timeless perspective' (Woodford, 2003), in which the central bank respects the commitments that it has optimally made at a date far in the past. This allows us to study systematic changes in monetary policy in response to shocks. We consider demand, financial and TFP shocks. The optimal response in all these cases features price stability. There is thus no meaningful trade-off between price stability and managing misallocation, just as the standard New Keynesian model features no trade-off between price stability and aggregate demand

²Our model nests the complete-market model as a particular case in which the borrowing constraint is arbitrarily loose, or in which entrepreneurs productivity levels are arbitrarily similar.

management (this is commonly known as the “divine coincidence”, [Blanchard and Gali, 2007](#)) . This implies that, even if the central bank can affect the supply side of the economy through its monetary policy, it finds optimal *not* to do so in a systematic way and it rather sticks to price stability.

The implementation of the optimal policy policy, however, differs with respect to the complete-market case. Under incomplete markets, a negative demand shock leads to an endogenous fall in TFP through the increase of capital misallocation. The fall in TFP, in turn, amplifies the reduction of the natural rate brought about by the demand shock itself, such that the natural rate drops more than in the case with complete markets.³ As the interest rate mimics the natural rate, it declines more, and more persistently, than in the standard New Keynesian model. This difference in implementation matters when nominal interest rates are constrained by the zero lower bound (ZLB). The optimal policy in this case, originally proposed by [Eggertsson and Woodford \(2004\)](#), is “low for longer”: nominal rates should remain at the ZLB longer than what would be prescribed if the ZLB were not present.⁴ In the case with incomplete markets, the larger and more persistent decline in natural rates due to the endogenous fall in TFP leads to what we call a “low for even longer” optimal policy: nominal rates should remain at the ZLB for significantly longer than they should under complete markets. Doing so reduces shortfalls not only of inflation and output, but also of TFP.

Finally, we present empirical evidence on the capital misallocation channel of monetary policy. We combine micro-level panel data on the quasi-universe of Spanish firms with monetary policy shocks identified using the high-frequency and sign-restrictions approach of [Jarociński and Karadi \(2020\)](#). The mechanism at play in the model is supported by the data: firms with high MRPK increase their investment relatively more than low-MRPK firms in response to an expansionary monetary policy shock. This implies a shift in the distribution of capital towards high-MRPK firms, improving the capital allocation. Using a simple model-derived measure of aggregate misallocation, we find that the positive impact of expansionary monetary policy is quantitatively in line with the data.⁵

³The natural rate is the rate that would pertain in the absence of nominal rigidities.

⁴See also [Eggertsson et al. \(2003\)](#); [Adam and Billi \(2006\)](#), and [Nakov et al. \(2008\)](#).

⁵The fact that TFP increases after a positive monetary policy shock has been previously documented by [Evans \(1992\)](#), [Christiano et al. \(2005\)](#), [Garga and Singh \(2021\)](#), [Jordà et al. \(2020\)](#), [Moran and Queraltó \(2018\)](#), [Meier and Reinelt \(2020\)](#) or [Baqaee et al. \(2021\)](#), among others. While these authors propose complementary mechanisms such as R&D, hysteresis effects, or markup heterogeneity to account for it, our findings suggest that capital misallocation also plays a significant role.

This paper contributes to three strands of the literature. First, our model is related to the extensive literature on capital misallocation, and the different channels that may affect it, such as [Hsieh and Klenow \(2009\)](#) or [Midrigan and Xu \(2014\)](#) – see [Restuccia and Rogerson \(2017\)](#) for a review on this literature. Our paper builds on [Moll \(2014\)](#), whose tractable heterogeneous producer model we augment with a New Keynesian monetary block to understand how monetary policy affects misallocation.⁶ The fact that monetary policy expansions reduce misallocation and increase TFP might seem in conflict with previous papers, such as [Reis \(2013\)](#), [Gopinath et al. \(2017\)](#), or [Asriyan et al. \(2021\)](#), who argue that a decline in real interest rates may fuel capital misallocation in real economies. We show that there is no such a conflict: our model also delivers a decline in TFP in response to a fall in real rates due to a negative demand shock when prices are flexible. The difference in the behavior of misallocation compared to a monetary policy shock is due to the different natural rate dynamics: though in response to both shocks the *real* rate drops, the *natural* rate falls only for the demand shock, remaining constant for the monetary policy shock. Just observing the dynamics of real interest rates is not sufficient to infer whether misallocation will improve or worsen. In the presence of nominal rigidities, it is the joint dynamics of real and natural rates that matter.

Second, an emerging literature analyzes how financial frictions and firm heterogeneity affect the transmission of monetary policy.⁷ Our paper is especially related to those papers analyzing the heterogeneous investment sensitivity to monetary policy. This strand of literature finds that firms' investment is more responsive to monetary policy shocks when their default risk is low ([Ottonello and Winberry, 2020](#)), when they have high leverage or fewer liquid assets ([Jeenas, 2020](#)), when they are young and do not pay dividends ([Cloyne et al., 2018](#)), when their excess bond premia is low ([Ferreira et al., 2022](#)), or when a higher fraction of their debt matures ([Jungherr et al., 2022](#)). We contribute to this literature by showing that firms with high-MRPK are more responsive to monetary policy shocks, and by analyzing the implications of this on the optimal

⁶[Buera and Nicolini \(2020\)](#) employ a discrete-time version of [Moll \(2014\)](#) with cash-in-advance constraints to analyze the impact of different monetary and fiscal policies after a credit crunch.

⁷One strand of this literature analyzes the links between monetary policy, firm heterogeneity and the allocation of resources through heterogeneity in markups and entry-exit (e.g. [Meier and Reinelt, 2020](#), [Bilbiie et al., 2014](#) or [Baqaee et al., 2021](#)), risk-taking ([David and Zeke, 2021](#)), firm-level productivity trends ([Adam and Weber, 2019](#)), or the importance of the price elasticity of investment ([Koby and Wolf, 2020](#)).

conduct of monetary policy.⁸

Finally, this paper contributes to the literature analyzing optimal monetary policy problems in heterogeneous-agent economies. A number of recent papers, such as [Bhandari et al. \(2021\)](#), [Acharya et al. \(2020\)](#), [Bilbiie and Ragot \(2020\)](#), [Le Grand et al. \(2022\)](#), [Mckay and Wolf \(2022\)](#), or [Auclert et al. \(2022\)](#) propose different approaches to solve these problems. Similar to [Nuño and Thomas \(2022\)](#), [Bigio and Sannikov \(2021\)](#), [Smirnov \(2022\)](#), or [Dávila and Schaab \(2022\)](#), we set up the problem as one of infinite-dimensional optimal control. We propose a new, simple and broadly applicable algorithm to solve these kinds of problems, which leverages the computational advantages of continuous time. The key novelty of our algorithm is to discretize the Ramsey planners' continuous-time, continuous-state problem using finite differences (as in [Achdou et al., 2021](#)), and then to use symbolic differentiation to obtain the planner's first-order conditions. This produces a high-dimensional nonlinear dynamic system, which is efficiently solved in the sequence space using a Newton algorithm.⁹ Furthermore, this paper is, to the best of our knowledge, the first to analyze optimal monetary policy in a model with heterogeneous firms.

The structure of the paper is as follows. Section 2 presents the model, which we calibrate in Section 3. Section 4 analyzes the drivers and dynamics of misallocation. Section 5 studies optimal monetary policy. Section 6 provides supporting empirical evidence for the main mechanism of the model. Finally, Section 7 concludes.

2 Model

We propose a New Keynesian closed economy model with financial frictions and heterogeneous firms. Time is continuous and there is no aggregate uncertainty. The economy is populated by five types of agents: households, the central bank, entrepreneurs that operate input-good firms, retail, and final goods producers. Entrepreneurs are heterogeneous in their net worth and productivity. They combine capital and labor to produce the input good. The input good is differentiated by imperfectly competitive

⁸In complementary empirical work, [Albrizio et al. \(2023\)](#) analyze the impact of monetary policy on capital misallocation in more detail, both from an intensive and extensive margin. They find that the intensive margin is the main reallocation channel and that firms' investment sensitivity to monetary policy is driven by their MRPK rather than their age, leverage, or cash.

⁹The algorithm can be implemented using several available software packages. To make our method accessible to a large audience, we employ Dynare.

retail goods producers facing sticky prices, whose output is aggregated by the final good producer. The latter two types of firms are standard in New Keynesian models.

2.1 Heterogeneous input-good firms

The heterogeneous-firm block is based on [Moll \(2014\)](#). There is a continuum of entrepreneurs. Each entrepreneur owns some net worth, which they hold in units of capital. They can use this capital for production in their own input-good producing firm – firm for short – or lend it to other entrepreneurs. Similar to [Gertler and Karadi \(2011\)](#), we assume that entrepreneurs are atomistic members of the representative household, to whom they may transfer dividends.¹⁰

Entrepreneurs are heterogeneous in two dimensions: their net worth a_t and their idiosyncratic productivity z_t .¹¹ Each entrepreneur owns a constant returns to scale (CRS) technology which uses capital k_t and labor l_t to produce the homogeneous input good y_t :

$$y_t = f_t(z_t, k_t, l_t) = (z_t k_t)^\alpha (l_t)^{1-\alpha}. \quad (1)$$

The capital share $\alpha \in (0, 1)$ is the same across entrepreneurs. Idiosyncratic productivity z_t follows a diffusion process,

$$dz_t = \mu(z_t)dt + \sigma(z_t)dW_t, \quad (2)$$

where $\mu(z)$ is the drift and $\sigma(z)$ the diffusion of the process.¹²

Entrepreneurs can use their net worth to produce in their firm with their own technology, or lend it to firms owned by other entrepreneurs. Firms employ labor l_t , which they hire at the real wage w_t , and capital k_t , which is the sum of the entrepreneur's net worth (a_t) and what the firm borrows ($b_t = k_t - a_t$) at the real cost of capital R_t . Capital is borrowed from the agents which save, i.e. both households and lending entrepreneurs.¹³

¹⁰This assumption is the only relevant difference between the real side of our model and the model of [Moll \(2014\)](#). We consider it to avoid having to deal with redistributive issues between households and entrepreneurs when analyzing optimal monetary policy. Both models produce linear dividend policies, so they can be seen as equivalent from a positive perspective.

¹¹For notational simplicity, we use x_t instead of $x(t)$ for the variables depending on time. Furthermore, we suppress the input goods firm's index.

¹²The process is bounded with reflective barriers.

¹³Since debt contracts are instantaneous and in units of capital, firms balance sheets are not exposed

Firms sell the input good at the real price $m_t = p_t^y/P_t$, which is the inverse of the gross markup associated to retail products over input goods, being p_t^y the nominal price of the input good and P_t the price of the final good, i.e. the numeraire. Entrepreneurs use the return on their activities to distribute (non-negative) dividends d_t to the household and to invest in additional capital at the real price q_t . Capital depreciates at rate δ . An entrepreneur's flow budget constraint can be expressed as follows

$$\dot{a}_t = \frac{1}{q_t} \left[\underbrace{m_t f_t(z_t, k_t, l_t) - w_t l_t - R_t k_t}_{\text{Firm's profits}} + \underbrace{(R_t/q_t - \delta)}_{\text{Return on net worth}} q_t a_t - \underbrace{d_t}_{\text{Dividends}} \right]. \quad (3)$$

Note that we have rearranged the budget constraint to yield the law of motion of net worth *in units of capital*.

Firms face a collateral constraint, such that the value of capital used in production cannot exceed $\gamma > 1$ of their net worth,¹⁴

$$q_t k_t \leq \gamma q_t a_t. \quad (4)$$

Entrepreneurs retire and return to the household according to an exogenous Poisson process with arrival rate η . Upon retirement they pay all their net worth, valued $q_t a_t$, to the household, and they are replaced by a new entrepreneur with the same productivity level. Entrepreneurs maximize the discounted flow of dividends, which is given by

$$V_0(z, a) = \max_{k_t, l_t, d_t} \mathbb{E}_0 \int_0^\infty e^{-\eta t} \Lambda_{0,t} \left(\underbrace{d_t}_{\text{Dividends}} + \underbrace{\eta q_t a_t}_{\text{Terminal value}} \right) dt, \quad (5)$$

subject to the budget constraint (3), the collateral constraint (4), and the productivity process (2). Future profits are discounted by the household's stochastic discount factor $\Lambda_{0,t}$. Below we show that $\Lambda_{0,t} = e^{-\int_0^t r_s ds}$, where r_t is the real interest rate.

We can split the entrepreneurs' problem into two parts: a static profit maximization

to Fisherian debt deflation or financial accelerator effects (Bernanke et al., 1999). Asriyan et al. (2021) include the latter. This assumption keeps the model tractable. Allowing for such effects would amplify the effect of monetary policy shocks discussed later.

¹⁴Assuming alternatively that firms' borrowing is constrained to a multiple of their earnings (Lian and Ma, 2020) would amplify the effect of monetary policy shocks discussed later on. The increase in high-productivity firms' profits, that (as we explain below) drives the positive impact of expansionary monetary policy on TFP, would relax the constraint of high-productivity firms and improve the allocation of capital further.

problem and a dynamic dividend-distribution problem. First, entrepreneurs maximize firm profits given their productivity and net worth,

$$\max_{k_t, l_t} \{m_t f_t(z_t, k_t, l_t) - w_t l_t - R_t k_t\}, \quad (6)$$

subject to the collateral constraint (4).

Since the production function has constant returns to scale, entrepreneurs find it optimal to operate a firm at the maximum scale defined by the collateral constraint whenever their idiosyncratic productivity is high enough, that is whenever z exceeds a certain threshold level z_t^* . Else the optimal size of the firm is $k^*(z) = 0$, because they cannot run a profitable firm given their low productivity. In that case the borrowing constraint does not bind and the entrepreneur rents out its capital. From now on, we refer to the two groups of entrepreneurs as 'constrained' and 'unconstrained'. That is, firm's demand for capital and labor is:

$$k_t(z_t, a_t) = \begin{cases} \gamma a_t, & \text{if } z_t \geq z_t^*, \\ 0, & \text{if } z_t < z_t^*, \end{cases} \quad (7)$$

$$l_t(z_t, a_t) = \left(\frac{(1 - \alpha)m_t}{w_t} \right)^{1/\alpha} z_t k_t(z_t, a_t). \quad (8)$$

Firm's profits are then given by

$$\Phi_t(z_t, a_t) = \max \{z_t \varphi_t - R_t, 0\} \gamma a_t, \quad \text{where} \quad \varphi_t = \alpha \left(\frac{1 - \alpha}{w_t} \right)^{\frac{1-\alpha}{\alpha}} m_t^{\frac{1}{\alpha}}. \quad (9)$$

Note that the term $z_t \varphi_t$ is simply the marginal revenue product of capital (MRPK) of the firm with productivity z_t . The productivity cut-off, above which firms are profitable, is given by

$$z_t^* \varphi_t = R_t, \quad (10)$$

This expression tells us that the MRPK of the marginal firm equals the marginal cost of capital. Any firm with $z > z^*$ thus makes profits over and above the cost of capital. These profits arise despite perfect competition, because the borrowing constraint binds.

Note furthermore, that factor demands and profits are linear in net worth. This is a consequence of the CRS technology and makes the model significantly more tractable.

As discussed by Moll (2014), the assumption of CRS in firms' production function (1) can be seen as the limiting case of decreasing returns to scale (DRS), $y_t = [(z_t k_t)^\alpha (l_t)^{1-\alpha}]^\nu$, $\nu < 1$, when $\nu \rightarrow 1$. In the case with DRS, there is a threshold $z^*(a)$ which depends on net-worth, such that the firms with $z \leq z^*(a)$ are unconstrained and produce at their optimal level ($k^*(z)$), whereas those with $z > z^*(a)$ are constrained. When this threshold increases, previously marginally constrained firms become unconstrained and reduce their capital stock below the maximum implied by the constraint. When $\nu \rightarrow 1$, the optimal size of low-productivity firms, and hence its production, are very small, $k^*(z), y^*(z) \rightarrow 0$. Therefore our model should be understood as the tractable limit of the more realistic DRS case.¹⁵ This highlights a crucial point regarding the interpretation of the model: the threshold z^* does not capture entry/exit, but rather the limiting case of expansions and contractions of the optimal size of active firms. Entry and exit of firms in this model is exogenous, and given by the exogenous retirement rate η .

Second, entrepreneurs choose the dividends d_t that they pay to the household. The solution to this problem is derived in Appendix B.1. There we show how entrepreneurs never distribute dividends ($d_t = 0$) until retirement, when they bring all their net worth home to the household. The intuition is the following. The return on one unit of capital in the hands of the entrepreneur is *at least* ($R_t - \delta q_t$), while for the household the return of this unit of capital is *exactly* ($R_t - \delta q_t$). It is thus always worthwhile for entrepreneurs to keep their funds in the firm. To keep things simple, as in Gertler and Karadi, 2011 we assume the representative household uses a fraction $\psi \in (0, 1)$ of these dividends to finance the net worth of the new entrepreneurs, so net dividends are $(1 - \psi)$ of the net worth of retiring entrepreneurs.

Using (9), the law of motion of an entrepreneur's net worth (3) can thus be rewritten as¹⁶

$$\dot{a}_t = \frac{1}{q_t} [(\gamma \max \{z_t \varphi_t - R_t, 0\} + R_t - \delta q_t) a_t]. \quad (11)$$

¹⁵Ferreira et al., 2023 find that financially constrained firms are found across the entire firm-size distribution.

¹⁶In the model, there are no firm-level capital adjustment costs. Furthermore, due to CRS, changes in z^* imply that firms at the margin disinvest or reinvest fully instantaneously. In a DRS model the changes in the capital stock of a firm that switches from being constrained to being unconstrained and thus crossing the threshold $z^*(a)$ would be smaller. This is so since the optimal capital stock of unconstrained firms would be positive (and not zero, as in the limiting CRS case). These smaller changes in capital can be achieved by reductions in the gross investment rate, without requiring the reselling of capital, provided the depreciation rate is high enough.

2.2 Households

There is a representative household, composed of workers and entrepreneurs, that saves in capital D_t or in nominal instantaneous bonds B_t^N . Nominal bonds are in zero net supply. Workers supply labor L_t . The household maximizes

$$W_t = \max_{C_t, L_t, B_t^N, D_t} \int_0^\infty e^{-\rho_t^h t} u(C_t, L_t) dt. \quad (12)$$

$$\begin{aligned} s.t. \quad \dot{D}_t q_t &= (R_t - \delta q_t) D_t - S_t^N - C_t + w_t L_t + T_t, \\ \dot{B}_t^N &= (i_t - \pi_t) B_t^N + S_t^N, \end{aligned} \quad (13)$$

where S_t^N is the investment into nominal bonds and T_t are the profits received by the household, which is the sum of the profits of the capital and retail-goods producers (discussed below) and net dividends received from entrepreneurs $((1 - \psi)\eta q_t A_t)$.

We assume separable utility of CRRA form, i.e., $u(C_t, L_t) = \frac{C_t^{1-\zeta}}{1-\zeta} - \Upsilon \frac{L_t^{1+\vartheta}}{1+\vartheta}$. Solving this problem (see Appendix B.2 for details), we obtain the Euler equation,

$$\frac{\dot{C}_t}{C_t} = \frac{r_t - \rho_t^h}{\zeta}, \quad (14)$$

the labor supply condition

$$w_t = \frac{\Upsilon L_t^\vartheta}{C_t^{-\zeta}}, \quad (15)$$

and the Fisher equation

$$r_t = i_t - \pi_t, \quad (16)$$

where, for convenience, we have made use of the following definition of the real interest rate

$$r_t \equiv \frac{R_t - \delta q_t + \dot{q}_t}{q_t}, \quad (17)$$

which equals the real return on capital adjusted by capital gains and depreciation. Integrating the Euler equation (14), we can define the stochastic discount factor $\Lambda_{0,t}$ as

$$\Lambda_{0,t} \equiv e^{-\int_0^t \rho_s^h ds} \frac{u'_c(C_t)}{u'_c(C_0)} = e^{-\int_0^t r_s ds}.$$

2.3 Final good producers

As usual in New Keynesian models, a competitive representative final goods producer aggregates a continuum of output produced by retailer $j \in [0, 1]$,

$$Y_t = \left(\int_0^1 y_{j,t}^{\frac{\varepsilon-1}{\varepsilon}} dj \right)^{\frac{\varepsilon}{\varepsilon-1}}, \quad (18)$$

where $\varepsilon > 0$ is the elasticity of substitution across goods. Cost minimization implies

$$y_{j,t}(p_{j,t}) = \left(\frac{p_{j,t}}{P_t} \right)^{-\varepsilon} Y_t, \text{ where } P_t = \left(\int_0^1 p_{j,t}^{1-\varepsilon} dj \right)^{\frac{1}{1-\varepsilon}}.$$

2.4 Retailers

We differentiate between heterogeneous input-good firms and retailers.¹⁷ We assume that monopolistic competition occurs at the retail level. Retailers purchase input goods from the input-good firms, differentiate them and sell them to final good producers. Each retailer j chooses the sales price $p_{j,t}$ to maximize profits subject to price adjustment costs as in [Rotemberg \(1982\)](#), taking as given the demand curve $y_{j,t}(p_{j,t})$ and the price of input goods, p_t^y . We assume the government pays a proportional constant subsidy $\tau = \frac{\varepsilon}{\varepsilon-1}$ on the input good, so that the net real price for the retailer is $\tilde{m}_t = m_t(1 - \tau)$. This subsidy is financed by a lump-sum tax on the retailers Ψ_t . This fiscal scheme is introduced to eliminate the distortions caused by imperfect competition in steady state, as common in the optimal policy literature. The adjustment costs are quadratic in the rate of price change $\dot{p}_{j,t}/p_{j,t}$ and expressed as a fraction of output Y_t ,

$$\Theta_t \left(\frac{\dot{p}_{j,t}}{p_{j,t}} \right) = \frac{\theta}{2} \left(\frac{\dot{p}_{j,t}}{p_{j,t}} \right)^2 Y_t, \quad (19)$$

where $\theta > 0$. Suppressing notational dependence on j , each retailer chooses $\{p_t\}_{t \geq 0}$ to maximize the expected profit stream, discounted at the stochastic discount factor of

¹⁷Distinguishing between heterogeneous input-good firms and retailers is standard practice in previous New Keynesian models featuring firm heterogeneity and nominal rigidities, such as [Ottonello and Winberry \(2020\)](#) or [Jeenas \(2020\)](#). Besides greater tractability, it avoids the possibly implausible countercyclical behaviour of New Keynesian markups interfering with our mechanism, which we see as an important advantage. It also does justice to the fact that retail prices are significantly more sticky than intermediate goods prices (for Europe see [Alvarez et al. \(2010\)](#), [Alvarez et al. \(2006\)](#), [Gautier et al. \(2023\)](#)).

the household,

$$\int_0^\infty \Lambda_{0,t} \left[\left(\frac{p_t}{P_t} - \tilde{m}_t \right) \left(\frac{p_t}{P_t} \right)^{-\varepsilon} Y_t - \Psi_t - \Theta_t \left(\frac{\dot{p}_t}{p_t} \right) \right] dt. \quad (20)$$

The symmetric solution to the pricing problem yields the New Keynesian Phillips curve (see Appendix B.3), which is given by

$$\left(r_t - \frac{\dot{Y}_t}{Y_t} \right) \pi_t = \frac{\varepsilon}{\theta} (\tilde{m}_t - m^*) + \dot{\pi}_t, \quad m^* = \frac{\varepsilon - 1}{\varepsilon}. \quad (21)$$

where π_t denotes the inflation rate $\pi_t = \dot{P}_t/P_t$.

2.5 Capital producers

A representative capital producer owned by the representative household produces capital and sells it to the household and the firms at price q_t , which she takes as given. Her cost function is given by $(\iota_t + \Xi(\iota_t)) K_t$ where ι_t is the investment rate and $\Xi(\iota_t)$ is a capital adjustment cost function. She maximizes the expected profit stream, discounted at the stochastic discount factor of the household. Profits are paid in a lump-sum fashion to the household.

$$\begin{aligned} W_t &= \max_{\iota_t, K_t} \int_0^\infty \Lambda_{0,t} (q_t \iota_t - \iota_t - \Xi(\iota_t)) K_t dt. \\ \text{s.t. } \dot{K}_t &= (\iota_t - \delta) K_t. \end{aligned} \quad (22)$$

The optimality conditions imply (see Appendix B.4)

$$r_t = (\iota_t - \delta) + \frac{\dot{q}_t - \Xi''(\iota_t) \dot{\iota}_t}{q_t - 1 - \Xi'(\iota_t)} - \frac{q_t \iota_t - \iota_t - \Xi(\iota_t)}{q_t - 1 - \Xi'(\iota_t)}. \quad (23)$$

2.6 Distribution

As previously explained, we assume that, for each entrepreneur retiring to the household, a new entrepreneur enters operating the same technology, that is, with the same productivity level. This new entrepreneur receives a startup capital stock from the household in a lump-sum fashion, equal to a fraction $\psi < 1$ of the net worth of the entrepreneur she replaces. Let $G_t(z, a)$ be the joint distribution of net worth and pro-

ductivity. The evolution of its density $g_t(z, a)$ is given by the Kolmogorov Forward equation

$$\frac{\partial g_t(z, a)}{\partial t} = \underbrace{-\frac{\partial}{\partial a}[g_t(z, a)s_t(z)a]}_{\text{Retained earnings}} - \underbrace{\frac{\partial}{\partial z}[g_t(z, a)\mu(z)] + \frac{1}{2}\frac{\partial^2}{\partial z^2}[g_t(z, a)\sigma^2(z)]}_{\text{Productivity changing randomly}} + \underbrace{-\eta g_t(z, a)}_{\text{Entrepreneurs retiring}} + \underbrace{\frac{\eta}{\psi}g_t(z, \frac{a}{\psi})}_{\text{Entrepreneurs entering}}, \quad (24)$$

where $s_t(z)$ is the entrepreneurs' investment rate from (11)

$$s_t(z) \equiv \frac{\dot{a}_t}{a_t} = \frac{1}{q_t} \left(\underbrace{\gamma \max\{z\varphi_t - R_t, 0\}}_{\text{Profit rate from operating the firm}} + R_t - \delta q_t \right), \quad (25)$$

and $1/\psi g_t(z, a/\psi)$ is the density of new entrepreneurs entering.

Using this two-dimensional distribution we can define the one-dimensional distribution of *net-worth shares* as $\omega_t(z) \equiv \frac{1}{A_t} \int_0^\infty a g_t(z, a) da$. This distribution measures the share of net worth held by entrepreneurs with productivity z . Due to the linearity of the firms choices in a , it contains all the relevant information in a more compact form. Given this definition and the structure of the problem, net-worth shares are non-negative, continuous, once differentiable everywhere and they integrate up to 1. The law of motion of net worth shares is given by (see Appendix B.5)

$$\frac{\partial \omega_t(z)}{\partial t} = \left[s_t(z) - \frac{\dot{A}_t}{A_t} - (1 - \psi)\eta \right] \omega_t(z) - \frac{\partial}{\partial z} \mu(z) \omega_t(z) + \frac{1}{2} \frac{\partial^2}{\partial z^2} \sigma^2(z) \omega_t(z). \quad (26)$$

2.7 Market Clearing and Aggregation

Market clearing. Define aggregate capital used in production as $K_t = \int k_t(z, a) dG_t(z, a)$, aggregate firm net worth as $A_t = \int a dG_t(z, a)$, and aggregate net debt as $B_t = \int b_t(z, a) dG_t(z, a)$. Since the capital borrowed by an individual entrepreneur equals that used in production minus its net worth $b_t = k_t - a_t$, we have that

$$K_t = A_t + B_t, \quad (27)$$

Asset market clearing requires that net borrowing of entrepreneurs B_t equals net savings of the household D_t ,

$$B_t = D_t. \quad (28)$$

Let $\Omega(z)$ be the cumulative distribution of net-worth shares, i.e. $\Omega_t(z) = \int_0^z \omega_t(x) dx$. By combining equations (27), (28), aggregating capital used by firms (7), and solving for A_t , we obtain

$$A_t = \frac{D_t}{\gamma(1 - \Omega_t(z_t^*)) - 1}. \quad (29)$$

Labor market clearing implies

$$L_t = \int_0^\infty l_t(z, a) dG_t(z, a). \quad (30)$$

Aggregation. Aggregating up, one can express output as a function of aggregate factors and aggregate TFP

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha}, \quad (31)$$

where aggregate TFP Z_t is an endogenous variable given by

$$Z_t = \left(\mathbb{E}_{\omega_t(\cdot)} [z \mid z > z_t^*] \right)^\alpha = \left(\frac{\int_{z_t^*}^\infty x \omega_t(x) dx}{1 - \Omega_t(z_t^*)} \right)^\alpha. \quad (32)$$

This highlights that, in terms of output, the model is isomorphic to a standard representative-agent New Keynesian model with capital and TFP Z_t . TFP is endogenous and evolves over time and, as we discuss below, its evolution depends on factor prices.

Note that TFP Z_t serves as a measure of misallocation. The financial frictions faced by entrepreneurs imply that capital is not optimally allocated. The entrepreneur operating the most productive firm does not have enough net worth to operate the whole capital stock, hence less productive firms operate as well, which is suboptimal and reduces TFP. Thus the more misallocated capital is, the lower is TFP.

Factor prices are

$$w_t = (1 - \alpha) m_t Z_t K_t^\alpha L_t^{-\alpha}, \quad (33)$$

$$R_t = \alpha m_t Z_t K_t^{\alpha-1} L_t^{1-\alpha} \frac{z_t^*}{\mathbb{E}_{\omega_t(\cdot)} [z \mid z > z_t^*]}. \quad (34)$$

Finally, the law of motion of the aggregate net-worth of entrepreneurs is given by

$$\frac{\dot{A}_t}{A_t} = \frac{1}{q_t} [\gamma(1 - \Omega_t(z_t^*)) (\alpha m_t Z_t K_t^{\alpha-1} L_t^{1-\alpha} - R_t) + R_t - \delta q_t - q_t(1 - \psi)\eta]. \quad (35)$$

Appendix B.6 derives these aggregate formulae step by step.

2.8 Central Bank

The central bank controls the nominal interest rate i_t on nominal bonds held by households. For the positive analysis in Section 4 we assume that the central bank sets the nominal rate according to a standard Taylor rule of the form

$$di = -v (i_t - (\rho^h + \phi(\pi_t - \bar{\pi}) + \bar{\pi})) dt, \quad (36)$$

where $\bar{\pi}$ is the inflation target, ϕ is the sensitivity to inflation deviations and v determines the persistence of the policy rule. For the normative analysis in Section 5 we assume that the central bank implements the Ramsey-optimal policy.

3 Numerical solution and calibration

Numerical algorithm. We solve the model numerically using a new method, described in Appendix C. It combines a discretization of the model using an upwind finite-difference method similar to the one in Achdou et al. (2021) with a Newton algorithm that computes non-linear transitional dynamics in a single loop. This can be easily implemented using Dynare’s perfect foresight solver.¹⁸

Our solution approach is different from the one in Winberry (2018) or Ahn et al. (2018). These papers analyze heterogeneous-agent models with aggregate shocks building on the seminal contribution by Reiter (2009). To this end, they linearize the model around the deterministic steady state. Winberry (2018) illustrates how this can be also implemented using Dynare, and Ahn et al. (2018) extend the methodology to continuous-time problems. Here, instead, we compute the *nonlinear* transitional dynamics in the the sequence space, as Boppart et al. (2018) or Auclert et al. (2021).

¹⁸Notice that the variables of the model include the distribution $\omega(z)$, which is an infinite-dimensional object. The finite-difference discretization turns this continuous variable into a finite dimensional vector.

Boppart et al. (2018) show how the perfect-foresight transitional dynamics to a (small) MIT shock, such as the ones we compute here, coincide with the impulse responses obtained by a first-order perturbation approach in the model with aggregate uncertainty.

Our method solves for the same approximate solution as the nonlinear version of the sequence space Jacobian approach by Auclert et al. (2021). An important technical difference is that we solve simultaneously for all variables (prices, aggregates and distributions) in a single loop without decomposing the model in blocks.

Calibration. Table 1 summarizes our calibration. We calibrate the parameters of the heterogeneous firms block to match data on Spanish firms. The entrepreneur’s exit rate (η) is set to 10 percent, in line with the average exit rate 2007-2020 in the data from the Spanish Statistical Institute (INE).¹⁹ The other parameters of the heterogeneous-firm block are disciplined by detailed firm-level panel data on the quasi-universe of Spanish firms. We postpone the description of this data to the empirical Section 6, with further details in Appendix A.1. The fraction of assets of exiting entrepreneurs reinvested (ψ) is 0.1, so that new entrant’s account for 1 percent of the total capital stock, in line with the data. The borrowing constraint parameter γ is 1.56, implying that entrepreneurs can borrow up to 56% of their net worth, or 36% of their total assets, which targets the aggregate debt to total assets ratio in the data. We assume that individual productivity z follows an Ornstein-Uhlenbeck process in logs,²⁰ with a reflective lower (upper) barriers at some value close to 0 (very high value).²¹

$$d \log(z) = -\varsigma_z \log(z)dt + \sigma_z dW_t. \quad (37)$$

We estimate this process using our firm panel data set, as explained in Appendix A.5. The estimate for ς_z corresponds to an annual persistence of 0.83, and the annual volatility of the shock is estimated to be 0.73.

The conventional macro parameters are set to standard values. The rate of time preference of the household ρ^h is 0.01, which targets an average real rate of return of 1 percent. The capital share α is set at 0.35 and the capital depreciation rate δ at

¹⁹Specifically, the data comes from the *Directorio Central de Empresas*, which is a dataset maintained by INE, and it contains aggregate data on all firms operating in Spain, and its status (incumbent, entrant or exiter). The dataset can be accessed [here](#).

²⁰By Ito’s lemma, this implies that z in levels follows the diffusion process $dz = \mu(z)dt + \sigma(z)dW_t$, where $\mu(z) = z \left(-\varsigma_z \log z + \frac{\sigma_z^2}{2} \right)$ and $\sigma(z) = \sigma_z z$.

²¹We truncate the process for z at 48. This corresponds to truncating the MRPK distribution at the same level as in the data.

Table 1: **Calibration**

Parameter	Value	Source/target	
η	Firms' death rate	0.1	Average exit rate
ψ	Fraction firms' assets at entry	0.1	Capital of firms younger than 1 year / All firms capital
γ	Borrowing constraint parameter	1.56	Debt / Total Assets firms
ς_z	Mean reverting parameter	0.19	Estimate based on firm level data
σ_z	Volatility of the shock	0.73	Estimate based on firm level data
ρ^h	Household's discount factor	0.01	1%
α	Capital share in production function	0.35	Gopinath et al. (2017)
δ	Capital depreciation rate	0.06	Gopinath et al. (2017)
ζ	Intertemporal elasticity of substitution HH	1	Log utility in consumption
ϑ	Inverse Frisch Elasticity	1	Kaplan et al. (2018)
Υ	Constant in disutility of labor	0.71	Normalization $L = 1$
ϕ^k	Capital adjustment costs	8	VAR evidence Christiano et al. (2016)
ϵ	Elasticity of substitution retail goods	10	Mark-up of 11%
θ	Price adjustment costs	100	Slope of Phillips curve of 0.1 as in Kaplan et al. (2018)
$\bar{\pi}$	Inflation target	0	Standard
ϕ	Slope Taylor rule	1.5	Standard
v	Persistence Taylor rule	0.2	Standard

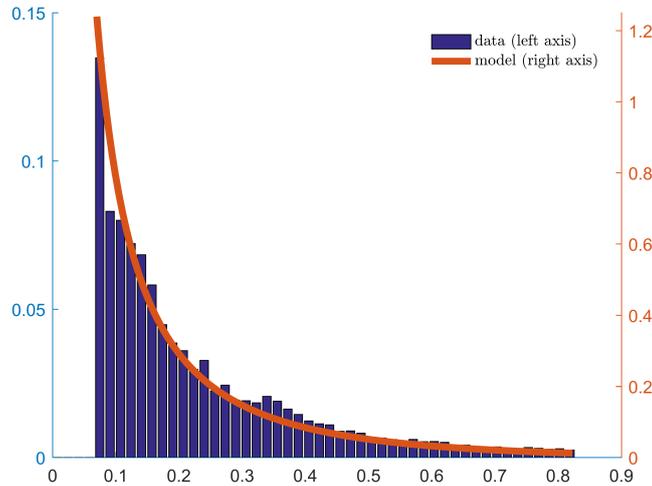
0.06, as in [Gopinath et al. 2017](#). We assume log-utility in consumption ($\zeta = 1$) and the inverse Frisch elasticity ϑ is also set to 1, standard values in the literature. We set the constant multiplying the disutility of labor Υ such that aggregate labor supply in steady state is normalized to one. We assume capital adjustment costs are quadratic, i.e. $\Xi(\iota_t) = \frac{\phi^k}{2}(\iota_t - \delta)^2$, and set ϕ^k to 8, such that the peak response of investment to output after a monetary policy shock is around 2, in line with the VAR evidence of [Christiano et al. \(2016\)](#).

Regarding the New Keynesian block, the elasticity of substitution of retail goods ϵ is set to 10, so that the steady state mark-up is $1/(\epsilon - 1) = 0.11$. The Rotemberg cost parameter θ is set to 100, so that the slope of the Phillips curve is $\epsilon/\theta = 0.1$ as in [Kaplan et al. \(2018\)](#). The Taylor rule parameters take the following standard values: $\bar{\pi} = 0$, $\phi = 1.5$ and $v = 0.2$.

The model generates the steady-state MRPK distribution shown by the orange line in [Figure 3](#). The dark blue bars show the MRPK distribution in the data, where a firm's MRPK is proxied by the product of a firm's value added over capital and its

capital share, as explained in Appendix A.2. Despite its simplicity, the model matches this distribution well, thus predicting a plausible amount of capital misallocation.²² Note that matching this distribution is key, since our results exploit the heterogeneous response of investment along the MRPK distribution.

Figure 1: MRPK distribution



Notes: The figure shows the steady state distribution of firms MRPK in the model (orange solid line) and compares it to the data (histogram with blue bars). See Appendices A.1 and A.2 for more details on the measurement of firm level MRPK in the data and robustness to differences in sectoral capital shares. We drop observations above an MRPK of 0.82, which implies dropping firms in the 5% upper tail of the capital-weighted MRPK distribution. Note that, by construction, the model cannot explain firms with an MRPK below the cost of capital ($R = r + \delta = 0.07$ in steady state), which we also drop in this figure.

4 Monetary policy and capital misallocation

In this Section, we analyze the links between monetary policy and capital misallocation. To this end, we first delve into the theoretical mechanisms through which changes in equilibrium prices affect misallocation, to then analyze the response of misallocation in general equilibrium to a monetary policy and a time-preference shock. Understanding these mechanisms is key to understand the optimal monetary policy explained in the following Section 5.

²²In Appendix A.2, we reproduce the MRPK distribution of Figure 3, but computing MRPK using sectoral alphas. The fit of the model worsens only slightly in the direction of under-predicting the measured misallocation.

4.1 The capital misallocation channel of monetary policy

As Section 2 highlighted, misallocation and thus TFP is endogenous and evolves over time. Misallocation is driven by the investment dynamics within the heterogeneous input-goods firms block of the model, which in turn depend on the factor prices determined in general equilibrium. One can thus think about the heterogeneous firm sector as a model block that translates (sequences of) prices into (a sequence of) TFP.

As discussed above, by equation (32) which we reproduce here, TFP depends on the allocation of capital across entrepreneurs:

$$Z_t = \left(\mathbb{E}_{\omega_t(\cdot)} [z \mid z > z_t^*] \right)^\alpha, \quad (38)$$

That is, TFP is the capital-weighted average of firms' idiosyncratic productivity. TFP thus depends on the mass of the net-worth distribution, $\omega_t(\cdot)$, above the productivity threshold, z_t^* (the shaded area in Figure 4.1). Entrepreneurs below the cut-off z^* are unconstrained, operate at their optimal size $k(z) = 0$, and lend their net worth to constrained entrepreneurs above the cut-off. Equation (38) allows us to identify how changes in equilibrium prices affect aggregate TFP in this economy (i) by changing the *net-worth distribution*, $\omega_t(\cdot)$; and (ii) by changing the *productivity-threshold* z_t^* . We now explore these two margins in isolation.

We start analyzing the case in which the dynamics of TFP are driven purely by changes in the net worth distribution, which happens when the cut-off z_t^* is constant. In this case, the *excess investment rate* is key for the dynamics of TFP. We define the excess investment rate as the ratio of profits over net worth

$$\tilde{\Phi}_t(z) \equiv \frac{\Phi_t}{q_t a_t} = \max \left\{ \frac{\gamma \varphi_t}{q_t} (z - z_t^*), 0 \right\}, \quad (39)$$

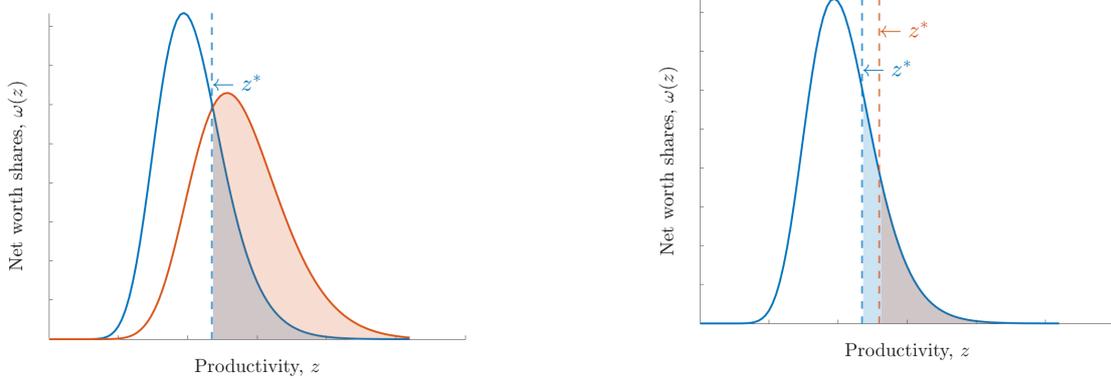
where $\tilde{\Phi}_t(z)$ is the return on equity that a firm with MRPK $\varphi_t z$ makes *in excess of* the cost of capital R_t . Since entrepreneurs reinvest all profits, $\tilde{\Phi}(z)$ also describes the speed at which the net worth of an entrepreneur with productivity z grows *in excess of* the growth rate of the unconstrained entrepreneurs with productivity $z \leq z_t^*$.²³

²³ $\tilde{\Phi}(z)$ is also a measure of how constrained a firm is, since $\tilde{\Phi}(z)$ is the Lagrange multiplier of the borrowing constraint in the firm's maximization problem. From the first order conditions of the firm, we get that $MRPK_t = R_t + q_t \lambda^{BC}$, where λ^{BC} is the multiplier on the borrowing constraint. Hence, $\lambda^{BC} = \tilde{\Phi}_t(z)$.

Figure 2: The capital misallocation channel.

(a) Change in the net-worth distribution.

(b) Shift in the productivity threshold.



Notes: The figure illustrates the net-worth share distribution $\omega(z)$ and the productivity-threshold z^* (blue). The light blue area is the initial mass of constrained firms. Panel (a) shows the impact of a change in the net-worth distribution. Panel (b) shows the impact of an increase in the threshold (from blue dashed line to orange dashed line). The new mass of constrained firms after the change is depicted by the shaded orange area in both panels.

Proposition 1. (TFP response to the slope). *Conditional on a constant cutoff z^* , the dynamics of Z_t are fully determined by the slope of the excess investment rate, $\frac{\gamma\varphi_t}{q_t}$. An increase in $\frac{\gamma\varphi_t}{q_t}$ leads to an increase in the growth rate of TFP through changes in the net worth distribution:*

$$\frac{\partial}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)} \frac{d \log Z_t}{dt} \Big|_{z^*} = \frac{\int_{z^*}^{\infty} z^2 \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} z \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz} > 0.$$

All the proofs in the Section can be found in Appendix B.8. This proposition states that the slope $\frac{\gamma\varphi_t}{q_t}$ determines how, conditional on a constant cut-off z^* , the net-worth share distribution moves, and hence in which direction TFP evolves. $\frac{\gamma\varphi_t}{q_t}$ is a function of prices. If the slope increases, then high-productivity firms' profitability advantage widens, such that they grow faster than low-productivity firms, the net worth distribution shifts rightwards, the allocation of capital improves, and TFP increases, as represented in Panel a of Figure 4.1. Note that, in the model, high-productivity firms have a high MRPK, which is given by $\varphi_t z$. So we can equivalently say that an increase in the relative growth rate of high-MRPK firms improves TFP.

We turn next to the case when the distribution remains constant and the cut-off changes in response to price changes. This happens in the limit of iid shocks, that is,

the limit as $\varsigma_z \rightarrow \infty$, as discussed in [Itskhoki and Moll \(2019\)](#). In this case the net worth distribution $\omega(\cdot)$ is constant, and the response of TFP growth depends exclusively on the changes in the cutoff, according to the following proposition

Proposition 2. (TFP response to the cutoff). *Conditional on a constant density $\omega(\cdot)$, the dynamics of Z_t are fully determined by the threshold z_t^* . The partial derivative of TFP growth with respect to the growth rate of the threshold $\frac{dz_t^*}{dt}$ is positive:*

$$\frac{\partial}{\partial \left(\frac{dz_t^*}{dt} \right)} \frac{d \log Z_t}{dt} \Big|_{\omega(\cdot)} = \frac{\alpha \omega(z_t^*) \int_{z_t^*}^{\infty} (z - z_t^*) \omega(z) dz}{\int_{z_t^*}^{\infty} \omega(z) dz \int_{z_t^*}^{\infty} z \omega(z) dz} > 0.$$

This proposition implies that a change in prices that raises the threshold $z_t^* = R_t/\varphi_t$ increases TFP. Panel (b) in [Figure 4.1](#) illustrates how an increase in the threshold decreases the share of constrained firms by crowding out low-productivity entrepreneurs. The intuition is simple: low-MRPK constrained firms that were close to the threshold become unconstrained and reduce their capital optimally to 0, which implies that these entrepreneurs stop using their net worth for production, and instead they lend it to more productive firms, decreasing misallocation. Changes in the productivity-threshold thus capture changes in the share of constrained versus unconstrained firms. This mechanism is different from the extensive margin: it is not meant to capture firm entry and exit, which in our model is exogenously given by the probability of retiring η . Rather, it captures the idea that previously constrained firms become unconstrained and vice versa.

In general equilibrium these two margins simultaneously determine the evolution of TFP. However, rather than being two independent mechanisms, they are tightly connected through general equilibrium forces.

Corollary 1. *An increase in the slope of the excess investment rate, $\frac{\gamma\varphi_t}{q_t}$, is associated to an increase in the growth rate of the threshold $\frac{dz_t^*}{dt}$, iff it is associated to a sufficiently small increase in the growth rate of the ratio between households' net savings and firms' net worth: $\partial \left(\frac{d \frac{D_t}{A_t}}{dt} \right) / \partial \left(\frac{\gamma\varphi_t}{q_t} \right) < \hat{\Delta}_t$ where $\hat{\Delta}_t \equiv \frac{(\mathbb{E}[z|z > z_t^*] - z_t^*)(1 - \Omega(z_t^*))}{\omega_t(z_t^*)} > 0$:*

$$\frac{\partial \left(\frac{dz_t^*}{dt} \right)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)} > 0 \iff \frac{\partial \left(\frac{d \frac{D_t}{A_t}}{dt} \right)}{\partial \frac{\gamma\varphi_t}{q_t}} < \hat{\Delta}_t.$$

That is, the two margins work in the same direction unless the supply of capital from the household strongly counteracts the dynamics in the firm sector. This is intuitive: an increase in the slope of the excess investment function $\frac{\gamma\varphi_t}{q_t}$ implies a rightwards shift of the net worth share distribution. *Ceteris paribus*, this increases the demand for capital and reduces its supply. Thus, the threshold z_t^* has to increase to clear the capital market, unless the household's supply of capital simultaneously increases by a lot.

This corollary allows us to characterize the capital misallocation channel of monetary policy, that is, the response of misallocation and TFP to monetary policy. If expansionary monetary policy shocks increases the slope $\frac{\gamma\varphi_t}{q_t}$ – which is the case in a wide range of New Keynesian models – while not increasing the supply of capital by the household relative to entrepreneurial net worth too much, then misallocation decreases and TFP increases. To see if this is indeed the case, we now turn to general equilibrium analysis.

4.2 Misallocation in general equilibrium

Households' time preference shock without nominal rigidities. While our primary focus in this Section is on monetary policy shocks, it is instructive to start with the response to a time-preference shock, that is a temporary fall in the household's discount factor ρ^h , abstracting for the moment from nominal rigidities, or equivalently assuming strict inflation targeting (dashed orange line in Figure 4.2).²⁴ This shock causes a decrease of real interest rates and a hump shaped decline of TFP. The drop in TFP results from a deterioration of the capital allocation due to the two interconnected effects discussed above.

First, the share of net-worth held by high-MRPK firms decreases because their profitability advantage over low-MRPK firms is squeezed, which lowers their relative growth rate. Said differently, the slope of the excess investment rate $\frac{\gamma\varphi_t}{q_t}$ drops (panel i). This decrease is largely brought about by the increase in capital prices (panel e), since φ_t is largely unaffected by the shock, which barely moves wages and input good prices (panels b and c). These price movements are similar to those in the complete-market limit of our model, that is in the standard representative agent model (RA), as Figure

²⁴For ease of comparison, we pick the time path for ρ^h such that the implied real rate coincides with that resulting from a monetary policy shock analyzed next. This implies a drop of ρ^h of 2.3 b.p., and a gradual return to the steady state level.

13 in Appendix B.10 shows.²⁵ The price dynamics that determine the evolution of the net-worth share distribution are thus standard.

Second, some previously unconstrained low-MPRK firms increase their scale and produce at full capacity, becoming constrained. Formally speaking, the cut-off $z_t^* = \frac{r_t q_t + \delta q_t - q_t}{\varphi_t}$ (eq. 10) falls. The drop in z_t^* is the result of the drop in the real rate r_t , which is only partially compensated by the increase in capital prices q_t . The intuition follows directly from the corollary above, with inverted sign: the decrease in the slope of the excess investment function implies a leftwards shift of the net worth share distribution while, at the same time, the rise in the household's patience increases the relative supply of households' savings. Thus, the threshold z_t^* has to decrease to clear the capital market.

Panel a of Figure 4 decomposes the evolution of TFP into the contributions of the individual price changes. Changes in intermediate prices m_t and wages w_t barely change TFP. The increase in capital prices q_t raises TFP, but this effect is trumped by the negative impact on TFP of the fall in real rates r_t .

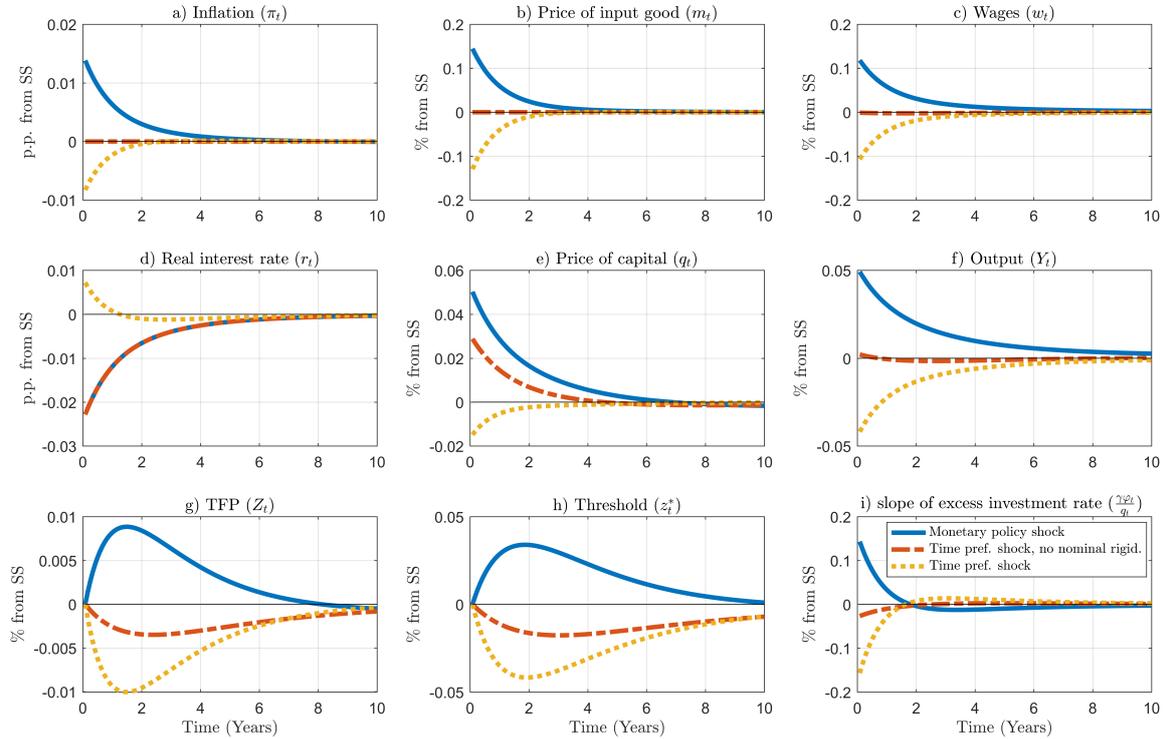
Monetary policy shock. Now we turn to the baseline model with nominal rigidities and analyze a monetary policy shock, that is, a 1 b.p. reduction in the nominal rate (blue lines in Figure 4.2) when the central bank follows the Taylor rule in equation (36). The natural rate, defined as the real interest rate in the counterfactual flexible price economy, remains constant as changes in the Taylor rule do not affect real variables under flexible prices. The decline in nominal rates causes an increase in output and inflation and a drop in the real rate (panels a, d, f) through the standard New Keynesian transmission mechanism.²⁶

Furthermore, TFP now increases (panel g). This is so, first, because the distribution of net worth shifts towards more productive firms as the slope of the excess investment rate $\frac{\gamma \varphi_t}{q_t}$ increases (panel i). This increase happens because now the changes in input-good prices m_t and wages w_t cause φ_t to rise by more than the price of capital, q_t

²⁵The complete-market economy is the standard representative agent New Keynesian model with capital. It represents a limit case of the baseline economy where either the borrowing constraint is infinitely loose, so that the net-worth distribution becomes irrelevant and only the most productive entrepreneur operates, or where the variance in entrepreneurial productivity z is 0. In this case, capital allocation is efficient (no misallocation) and TFP is exogenous. Appendix B.9 compares the baseline and complete-market models.

²⁶As before, factor prices respond largely as in the complete-market RA model (see Figure 13 in the appendix).

Figure 3: Impulse responses.

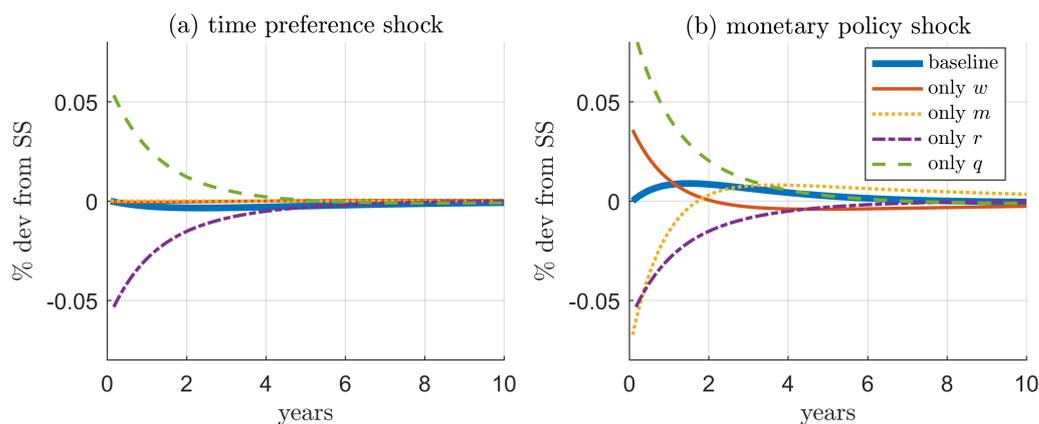


Notes: The figure shows the deviations from steady state of the economy. The solid blue line is the response of the baseline economy to an expansionary monetary policy shock of 1 basis point. The dashed orange line is the response of the economy to a time preference shock in the absence of nominal rigidities, where the path for ρ^h is chosen so as to reproduce the path of the real rate for the monetary policy shock. The yellow dotted line is the response to the same shock with sticky prices.

(panel e).²⁷ Second, the cut-off $z_t^* = \frac{r_t q_t + \delta q_t - \dot{q}_t}{\varphi_t}$ also increases (panel h), because the increase in the price of capital q_t overcompensates the decrease in the real rate r_t and the increase in φ_t . Contrary to the time-preference shock above, the monetary policy shock increases the relative profitability of high-MRPK firms, such that the net worth of constrained high-MRPK firms grows faster than the supply of capital from unconstrained entrepreneurs and households – which, unlike before, have no additional desire to save. Hence, the threshold z_t^* must thus increase to clear the capital market (see the corollary above), which means that constrained low-MRPK firms reduce their scale and

²⁷Input-good prices m_t and wages w_t affect φ_t in opposite directions, as the higher prices (panel b) increases excess profits whereas higher wages (panel c) reduces them. However, the elasticity with respect to both variables is different, being larger ($\frac{1}{\alpha}$) for prices than for wages ($\frac{1-\alpha}{\alpha}$). As the increase of wages and input-good prices is roughly of the same magnitude, the different elasticities explain why φ_t increases.

Figure 4: Decomposing the effect of a monetary policy shock on TFP



Notes: The figure decomposes the effect of a monetary policy shock on TFP (bold blue line) into the effect of the individual factor price changes. This is done by computing how TFP would have evolved if all prices but one would have remained at steady state.

become unconstrained. We refer to the impact of monetary policy on misallocation as the *capital misallocation channel of monetary policy*.

Panel b of Figure 4 shows how each factor price contributes to TFP. The increase in input-good prices m_t , which positively affects φ_t , has a positive effect on the excess investment rate and a negative effect on the threshold. The latter effect prevails in the short run, having a negative net effect of TFP, while the former prevails in the long run, with a positive net effect. The opposite happens with the increase in wages w_t , but to a smaller scale. As for the time preference shock, the decrease in the real rate r_t contributes negatively to TFP, while the increase in capital prices q_t contributes positively, with the latter dominating in this case.

Households' time preference shock with nominal rigidities. For completeness we also report the response to the time preference shock with nominal rigidities under a Taylor rule. The response (dotted yellow lines in Figure 4.2) is a combination of the response to the time preference shock absent price rigidities, or equivalently under strict inflation targeting, and the response to a negative monetary policy shock. To see this, note that the central bank does not reduce nominal rate as fast as under strict inflation targeting, and hence the real interest rate exceeds the natural rate (compare the dotted yellow and the dashed red line in panel d). The Taylor rule thus leads to a contractionary temporary deviation from the strict inflation targeting policy. As a

result TFP drops even more than with flexible prices.²⁸

Robustness and empirical support. Note that the responses of capital misallocation and TFP described above directly depend on the relative response of factor prices, which we can only solve for numerically in general equilibrium. So the direction of the response may depend on parameters. We show in Appendix B.11 that our results are robust to a wide range of parameters.

One may wonder if the positive response of TFP through a reduction in misallocation to a monetary expansion is empirically plausible. It has been widely documented that expansionary monetary policy shocks indeed raise TFP (see Evans, 1992; Christiano et al., 2005; Garga and Singh, 2021; Jordà et al., 2020; Moran and Queralto, 2018; Meier and Reinelt, 2020 or Baqaee et al., 2021). The peak effect on TFP predicted by our model (0.87 p.p. increase in TFP to a 1 p.p. decrease in rates) falls within the range 0.4-1.7 p.p. of medium-run peak responses of TFP to monetary policy shocks estimated by those papers. Furthermore, their estimated responses are also hump-shaped over the medium-run. The model is thus consistent with this evidence.

In Section 6 we go a step further and provide direct evidence that misallocation indeed contributes to the rise in TFP by confirming two testable implications of our theory: (i) that after an expansionary monetary policy shock high-MRPK firms increase their investment relatively more; and (ii) that overall misallocation decreases. Furthermore, we show that the quantitative response of the model is in the ballpark of that estimates. Taken together these results suggest that the capital misallocation channel may indeed contribute to the rise in TFP, and can hence complement the alternative mechanisms proposed by this literature such as R&D, hysteresis effects, or markup heterogeneity.

Relation to the misallocation literature. Our results resolve the apparent tension between the literature showing how monetary policy shocks affect TFP on the one hand, and the literature finding that low real rates may fuel misallocation on the other hand (e.g. Gopinath et al., 2017 or Asriyan et al., 2021).²⁹ We show that there is no such a conflict: our model delivers both an increase in TFP in response to an expansionary monetary policy shock and a decline in TFP in response to a negative demand shock when prices are flexible. The difference in the behavior of misallocation

²⁸The responses to a permanent time preference shock are qualitatively similar. See Appendix B.9.

²⁹They find that misallocation increases and TFP decreases in response to a decline in real interest rates in flexible-price economies.

is due to the different natural rate dynamics: though in response to both shocks the *real* rate drops, the *natural* rate falls only for the demand shock, remaining constant for the monetary policy shock.

5 Optimal monetary policy

Having analyzed the interaction of monetary policy and capital misallocation, we now turn to optimal policy.

5.1 Central bank objective and numerical approach

Ramsey problem. We assume that the central bank sets its policy instrument – the nominal interest rate i_t – such as to maximize household utility under full commitment. That is, the central bank solves the following Ramsey problem:

$$\max_{\{\omega(z), w, r, q, \varphi, R, K, A, L, C, D, Z, \mathbb{E}[z|z > z_t^*], \Omega, z^*, \iota, \pi, m, \tilde{m}, i, Y, T\}_{t \geq 0}} \mathbb{E}_0 \int_0^\infty e^{-\rho^h t} u(C_t, L_t) dt \quad (40)$$

subject to the all the private equilibrium conditions derived above and listed in Appendix B.7 and the initial conditions $\{\omega_0(z), K_0, D_0, A_0\}$. Relative to the standard New Keynesian model – i.e. the complete market version of our model – the problem of the central bank is richer by one dimension: the central bank understands that its policy affects TFP through the capital misallocation channel, and has to account for that.

Algorithm. This additional richness also makes the problem harder to solve computationally. The central bank’s controls include the net-worth distribution $\omega_t(z)$, as the central bank internalizes the impact of its decisions on it. Notice that the density $\omega_t(z)$ not only depends on time, but also on individual productivity. This poses a challenge when solving optimal monetary policy, as we need to compute the first order conditions (FOCs) with respect to this infinite-dimensional object. There are a number of proposals in the literature to deal with this problem. [Bhandari et al. \(2021\)](#) make the continuous cross-sectional distribution finite-dimensional by assuming that there are N agents instead of a continuum. They then derive standard FOCs for the planner. In order to cope with the large dimensionality of their problem, they employ a perturbation technique. [Le Grand et al. \(2022\)](#) employ the finite-memory algorithm proposed by [Ragot \(2019\)](#). It requires changing the original problem such that, after

K periods, the state of each agent is reset. This way the cross-sectional distribution becomes finite-dimensional. Nuño and Thomas (2022), Bigio and Sannikov (2021), Smirnov (2022) and Dávila and Schaab (2022) deal with the full infinite-dimensional planner’s problem. This implies that the continuous Kolmogorov forward (KF) and the Hamilton-Jacobi-Bellman (HJB) equations are constraints faced by the central bank. They derive the planner’s FOCs using calculus of variations, thus expanding the original problem to also include the Lagrange multipliers, which in this case are also infinite-dimensional. These papers solve the resulting differential equation system using the upwind finite-difference method of Achdou et al. (2021).

Here we propose a new algorithm, detailed in Appendix D. Instead of determining the FOCs for the planner’s continuous space problem, we first discretize the planner’s objective and constraints (the private equilibrium conditions) using finite differences. This transforms the original infinite-dimensional problem into a high-dimensional problem, in which the value function and the state density are replaced by large vectors with a dimensionality equal to the number of grid points (200 in our application) used to approximate the individual state space. In this discretized model the dynamics of the (now finite-dimensional) distribution ω_t are given by $(\mathbf{I} - \Delta t \mathbf{A}_t^T) \omega_t = \omega_{t-1}$, where Δt is the time step and \mathbf{A}_t is a matrix whose entries depend *nonlinearly* and in *closed form* on the idiosyncratic and aggregate variables in period t .³⁰

Second, we find the planner’s FOCs *by symbolic differentiation*. This delivers a large-dimensional system of difference equations. Third, we find the Ramsey steady state by solving this system at steady state. To do so, we compute the steady state of the model conditional on the steady-state level of the policy instrument with a conventional iterative method, and then use this function to find the Ramsey steady state using the Newton method. Fourth, we solve the system of difference equations non-linearly in the sequence space using the Newton method, as already described in Section 3 and Appendix C. The symbolic differentiation and the two applications of the Newton algorithm can conveniently be automated using several available software packages. In our case, we employ Dynare, but the approach is also compatible with the nonlinear sequence space Jacobian toolbox. This algorithm can be employed to compute optimal policies in a large class of heterogeneous agent models. Compared to other techniques, it stands out for being easy to implement. In appendix D we compare

³⁰Technically, this matrix results from the discretization of the *infinitesimal generator* of the idiosyncratic states.

our method conceptually to the ones cited above. We also present a proposition showing that our algorithm delivers the same results as computing the FOCs by hand using calculus of variations and then discretizing the model, as the time step gets smaller. Finally we apply our the algorithm to solve the model in [Nuño and Thomas \(2022\)](#) in order to illustrate its generality, demonstrating that results coincide.

5.2 Optimal Ramsey policy

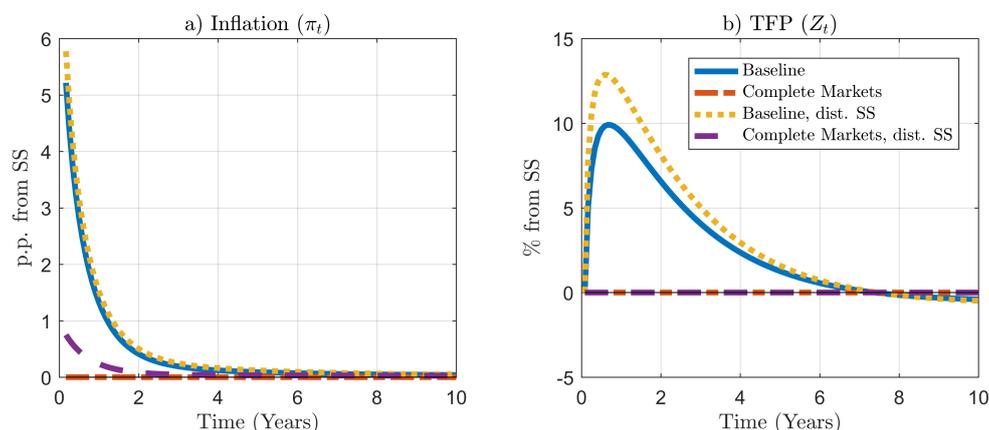
Steady state. Let us focus first on the steady state of the Ramsey problem. It is well known that in the standard (complete-market) New Keynesian economy without steady state distortions inflation is zero in the Ramsey steady state. Due to capital misallocation, our baseline (incomplete-market) economy does not feature steady state efficiency. Yet, inflation is still zero in the steady state of the Ramsey problem.³¹ This result mirrors a similar result from the textbook New Keynesian model with a distorted steady state ([Woodford, 2003](#); [Gali, 2008](#)). Though the long-run Phillips curve allows monetary policy to affect misallocation in the long run through positive trend inflation, the benefits of this policy are compensated for by the cost of the anticipation of this policy.

Time-0 optimal policy. We turn next to the deterministic dynamics under the Ramsey optimal plan. We solve for the Ramsey plan when the initial state of the economy coincides with the steady state under the optimal policy, i.e., that with zero inflation. The Ramsey planner faces no pre-commitments. This is commonly referred to as the “time-0 optimal policy” ([Woodford, 2003](#)).

We compare our baseline incomplete-market economy with a complete-market economy. The Ramsey plan in the model with complete markets is time-consistent. Hence, inflation and the rest of variables remain constant at their steady state values (see the dashed red lines in [Figure 5](#)). Market incompleteness, however, introduces a new source of time inconsistency, inducing the central bank to temporarily deviate from the zero-inflation policy. The solid blue lines in [Figure 5](#) show how the central bank engineers a sizable surprise monetary expansion, increasing inflation (panel a). The resulting dynamics are precisely those caused by an expansionary monetary policy shock, which were described in detail in [Section 4.2](#). As a result TFP increases (panel b). The central bank thus engineers a monetary expansion, tolerating a temporary increase in inflation,

³¹This is a numerical result that holds at close to machine precision for a wide range of parameters.

Figure 5: Time 0 optimal monetary policy.



Notes: The figure shows the deviations from steady state of the economy when the planner solves the Ramsey problem without pre-commitments and in the absence of shocks. The baseline economy is the solid blue line, and the complete market economy the dashed orange line. The dotted yellow line and the purple dashed line repeat the same exercise in the absence of the subsidy that undoes the markup distortion.

in order to achieve a persistent rise in TFP, brought about by a more efficient allocation of capital.

It is well known that the Ramsey policy in the complete market economy with a steady-state mark-up distortion also features inflationary time inconsistency. Comparing the optimal policy above with the optimal policy when there is no subsidy to correct for the mark-up distortion reveals that the time inconsistency problem caused by the incomplete market distortion is much larger: the optimal inflation level due to market incompleteness is more than six times higher than that due to the mark-up (dashed purple line). We hence conclude that the time inconsistency problem is not only large in absolute terms (with an average inflation of 3% during the first year), but also dwarfs the one resulting from markups in the standard New Keynesian model.

The desire of the central bank to redistribute resources towards high-MRPK entrepreneurs is reminiscent of the case with optimal fiscal policy analyzed by [Itskhoki and Moll \(2019\)](#). They find that optimal fiscal policy in economies starting at below steady-state net-worth levels initially redistributes from households towards entrepreneurs in order to speed up net worth accumulation, and thus increase TFP growth. In our case, and given the lack of fiscal instruments, it is the central bank who engineers this redistribution through an expansion in aggregate demand.

5.3 Timeless optimal policy response

Next, we analyze the optimal policy response when an unexpected shock hits the economy that was previously in its zero-inflation steady state. In this case, we adopt a “timeless perspective” (Woodford, 2003). Timelessly optimal Ramsey policy implies that the central bank sticks to pre-commitments, implementing the policy that it would have chosen to implement if it had been optimizing from a time period far in the past.³² This allows us to study systematic changes in monetary policy in response to shocks under the, ex-ante optimal, time-invariant state-contingent policy rule.³³

Households’ time preference shock. We analyze the optimal response to a time preference shock from a timeless perspective. Figure 6 shows that the optimal inflation response in the baseline economy (blue solid line) mimics that under complete markets (orange dashed line): the central bank stabilizes inflation at its steady state value of zero (panel a). This is what is usually known as “divine coincidence” (Blanchard and Gali, 2007): the real rate follows the natural rate (panel c) and output is at its natural level (panel b).³⁴ This result has important implications: despite the fact that the central bank can use monetary policy to cushion the impact of shocks by exploiting the misallocation channel of monetary policy, it chooses not to do so and sticks instead to strict price stability. As we show in Appendix B.12, this result applies also to other shocks, such as a TFP or a financial shock.

In order to implement the optimal policy in response to the time-preference shock, the central bank should lower the real and nominal rates. However, in the baseline model with incomplete markets, this requires that the central bank acts more forcefully than under complete markets (panel c). The reason is that the original demand shock leads to a negative ‘supply shock’ through its impact on aggregate TFP (panel d), which depresses output and natural rates relative to the complete market case, as discussed in the previous Section.

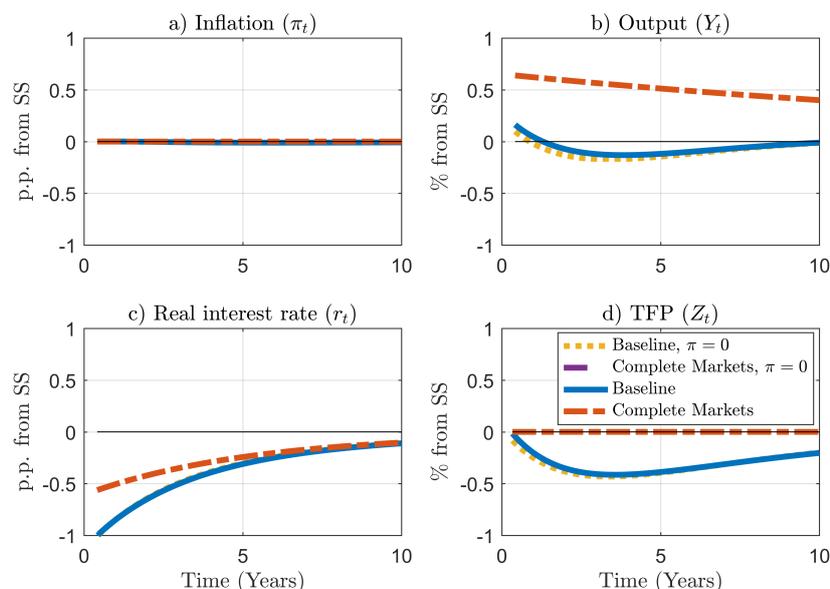
Zero lower bound. The fact that the central bank responds more persistently to the demand shock under the optimal policy has important implications when the zero lower bound constrains its room for maneuver. Figure 7 displays the optimal response

³²The Lagrange multipliers associated to forward-looking equations in the planner’s FOCs in this case are initially set to their steady state values.

³³As discussed in Section 3, building on the argument by Boppart et al. (2018) one can reinterpret the timeless response to MIT shocks as a first-order approximation to the response under uncertainty.

³⁴The natural level corresponds to the case with flexible prices (yellow and purple dashed lines). The divine coincidence holds only in an approximate sense, but the deviation is negligibly small.

Figure 6: Optimal monetary policy response to a household discount factor shock.

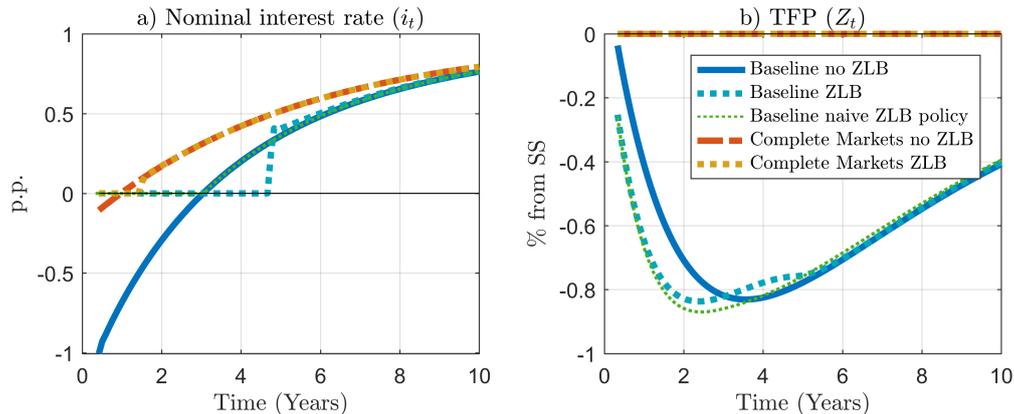


Notes: The figure shows the optimal response from a timeless perspective (in deviations from steady state) to a 1 p.p. decrease in the rate of time preference of the household ρ^h that is mean reverting with a yearly persistence of 0.8. The baseline economy is the solid blue line, and the complete market economy the dashed orange line. The figure also shows the paths of the variables under strict inflation targeting (yellow and purple lines), though they are barely visible since they are overlaid by the optimal policy paths.

from a timeless perspective to a large negative demand shock that drives the natural rate below the zero lower bound (ZLB). The optimal policy under complete markets, as shown by Eggertsson et al. (2003), is to adopt a “low for longer” strategy: The nominal rate (dotted yellow line, panel a) should remain at the ZLB for a longer period than it would in the absence of the ZLB (dashed orange line).

In the baseline economy with incomplete markets, optimal policy is also characterized by a low for longer strategy (dotted light blue line, panel a). However, the lift-off date is now delayed relative to the complete markets model. We call this a “low for *even* longer” policy. The reason is simple. As discussed above, natural rates fall more persistently in the case with incomplete markets, and so do nominal rates under the optimal policy without the ZLB (solid blue line). To compensate for the inability to move rates into negative territory, the central bank commits to stay low for even longer. Incomplete markets make this commitment even more important, since otherwise not only output and inflation, but also TFP would fall more. To see this, we compare the baseline (light blue dotted line) with a policy that sets the rate equal to the maximum of its optimal value in the absence of the ZLB and zero (green dotted line) and verify

Figure 7: Optimal monetary policy response to a demand shock with the zero lower bound.



Notes: The figure shows the optimal response from a timeless perspective (in deviations from steady state) to a 4 p.p. decrease in the rate of time preference of the household ρ^h that is mean reverting with a yearly persistence of 0.8. The baseline economy without the zero lower bound is the solid blue line, and the complete market economy without the zero lower bound is the dashed orange line. The dotted light blue line is the optimal response in the baseline economy with the zero lower bound, while the dotted green line is the response under a policy that sets the rate equal to the maximum of its optimal value in the absence of the ZLB and zero. The yellow dotted line is the optimal response in the complete market economy with the zero lower bound.

that TFP declines by more in the latter case.

6 Testable implications

A key prediction of our theory is that an expansionary monetary policy shock increases TFP by reducing misallocation. While the main focus of our paper is conceptual, we conclude by evaluating this testable implication. As already discussed in Section 4, the literature has repeatedly confirmed that expansionary policy shocks increase TFP. In this section, we show that a reduction in capital misallocation may contribute to this increase. First, we use firm level data to test the mechanism suggested by the model: after an expansionary monetary policy shock, high-MRPK firms increase their investment relatively more. Second, with a simple model-derived measure of misallocation, we quantify the impact of monetary policy shocks on misallocation and TFP in the data. In both cases we compare the results to the model predictions.

Data. For our empirical analysis we combine granular Spanish firm-level panel data with Jarociński and Karadi monetary policy shocks. We use yearly balance-sheet and cash-flow data from the quasi-universe of Spanish firms from 2000 to 2016 from the *Central de Balances Integrada*. The main advantage of this dataset is that it covers the

quasi-universe of Spanish firms, including not only large firms with access to stock and bond markets, but also medium and small firms more reliant on bank credit and internal financing. This contrasts with most papers in this literature, which use data from publicly traded firms (e.g. Compustat). These are generally large firms with access to the equity market, which can potentially behave very differently from the rest of firms in the economy, as documented for example by [Caglio et al. \(2021\)](#). Appendix [A.1](#) details the data definition and the cleaning process and reports descriptive statistics. Our key variable of interest is firm’s MRPK, which we proxy by value added over capital following the literature (see for instance [Bau and Matray, 2023](#)). Appendix [A.2](#) explains the rationale for this using proxy in more detail and explains that all the empirical results in this paper are robust to sectoral differences in capital shares.

The monetary policy shock is taken from [Jarociński and Karadi \(2020\)](#), who use sign restrictions to decompose unexpected high frequency movements of interest rates around policy announcements into an information surprise and a monetary policy surprise component. We use the latter component, and we aggregate these shocks to yearly frequency following the methodology employed by [Ottonello and Winberry \(2020\)](#). Appendix [A.3](#) provides more details on the identification and aggregation of the monetary policy shock.

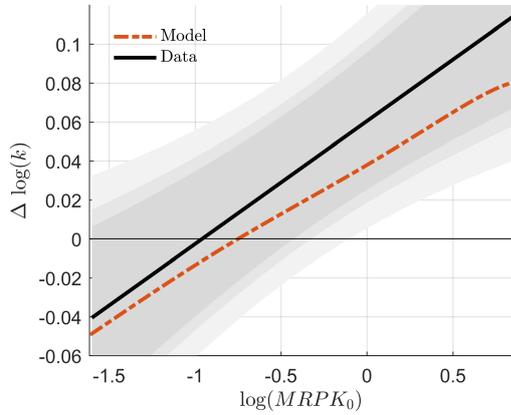
Firm level responses. The model predicts that TFP increases in response to a monetary shock, because the capital stock of firms with a higher MRPK grows relatively more. This happens both because constrained high-MRPK firms’ net worth increases in relative terms, which thus invest relatively more, and because low-MRPK firms’ optimal size is reduced, which thus invest less. To illustrate this, we simulate a monetary policy shock in the model and calculate the average response of firms’ capital stock as a function of their initial MRPK. The model predicts that this response is *near-linearly increasing* in the logarithm of pre-shock MRPK, as the dashed orange line in Figure [8](#) illustrates.

To test this prediction in the Spanish data, we estimate the following relationship, which is linear in log MRPK, as suggested by the model:

$$\log k_{j,t} - \log k_{j,t-1} = \beta_0 + \beta_1 \log (MRPK_{j,t-1}) + \beta_2 \log (MRPK_{j,t-1}) \varepsilon_t + \beta_3 \varepsilon_t + \Gamma' Y_{t-1} + \gamma_s + u_{j,t}. \quad (41)$$

where $k_{j,t}$ is the tangible capital of an individual firm j at time t , $MRPK_{j,t-1}$ is lagged MRPK, and ε_t is the monetary policy shock. While these controls would be sufficient

Figure 8: Response of investment to an expansionary monetary policy shock as a function of initial MRPK



Notes: The figure displays the average effect of an 1 p.p. expansionary monetary policy shock on the growth rate of the capital stock in the year after the shock in p.p. $-100 * (\log k_{j,1} - \log k_{j,0})$ – as a function of the firms' log MRPK before the shock $\log(MRPK_{j,0})$. For the model (orange), the relationship is calculated analytically. See Appendix B.13 for more detail. Estimating the regression (41) on simulated data would recover a linear approximation of it. We compare the model prediction to the estimated relationship (41) (black). The shaded areas mark the 90, 95 and 99% confidence intervals.

Table 2: Response of firm-level investment to an expansionary monetary policy shock

	(1)	(2)
	$\Delta \log k_{j,t-1,t}$	$\Delta \log k_{j,t-1,t}$
$\varepsilon_t \log(MRPK_{j,t-1})$	0.0470*** (0.02)	0.0286*** (0.01)
ε_t	0.0605*** (0.02)	
Obs	3,692,188	3,692,188
R^2	0.01	0.02
γ_{st}	No	Yes
γ_s	Yes	No

Notes: Column (1) reports the estimated differential effect (β_2) and average effect (β_3) from regression (41), that is, including sector fixed effects, aggregate controls (lagged GDP growth, inflation and unemployment). Standard errors clustered at the sector-year level. Column (2) reports the differential effect (β_2) estimated including sector-year fixed effects.

in the context of the model, we add further controls to account for non-modelled forces. Y_{t-1} are macroeconomic controls to account for the business cycle, that include GDP growth, inflation and unemployment and γ_s is a vector of sector fixed effects, which control for potential sector-specific confounders. Finally, $u_{j,t}$ is the residual.

The main coefficient of interest is β_2 , which is the empirical counterpart to the slope of the orange function in Figure 8: a positive value indicates that high-MRPK firms increase their investment more than low-MRPK firms after an expansionary monetary policy surprise. Table 2, column 1 shows that after a 1 p.p. expansionary monetary policy shock, a firm with an MRPK that is 1% higher than that of another firm increases its capital stock by 0.0470 p.p. more.

The coefficient β_3 corresponds to the intercept in Figure 8. Since the mean and standard deviation of $\log(MRPK_{j,t})$ are -0.87 and 1.4, these estimates document both a

substantially positive average effect of monetary policy on investment,³⁵ and – more importantly – an economically significant amount of heterogeneity in the firms’ responses. Being positive, these estimates thus support the models qualitative predictions. What is more, as Figure 8 shows, they are also quantitatively close to the model.

We perform a battery of robustness tests, by subsequently enriching the baseline specification in (41). First, we add sector-time fixed effects. This accounts for sectoral differences in capital shares, as further explained in Appendix A.2, and for other potential year-sector specific confounders. While more robust, the average effect of the monetary policy shock is absorbed in this specification. The estimate of the slope remains significant and of similar order of magnitude (see column 2 in Table 2).

Furthermore, we add firm-level fixed effects, firm-level controls, including measures of size, leverage and liquidity, introduce aggregate controls interacted with $\log(MRPK)$, and also interact the firm-level controls with the monetary policy shock. In all cases the coefficient β_2 remains positive and statistically significant, and of similar magnitude. We also demean the MRPK at the firm level, to ensure that the results are not driven by permanent heterogeneity in responsiveness across firms or systematic bias in our empirical proxy for the MRPK – in the spirit of Ottonello and Winberry (2020) – and we still find a positive and significant coefficient. Furthermore, these results are not driven by the smaller firms in our sample: if we restrict the analysis to large firms, the coefficient is still positive and significant, and even of greater magnitude. All of these results are reported in Appendix A.4.

In the model, the share of capital held by high-MRPK firms increases in response to a monetary expansion because the monetary expansion increases the profits of high-MRPK firms disproportionately, which are then invested. To test for this particular mechanism, in column (11) of Table 4 in Appendix A.4 we estimate the baseline regression (41) with sector-year fixed effects, but with the log change in profits on the left-hand side. We find the coefficient to be positive and significant, providing further support for the mechanism presented in the paper.

A follow-up empirical paper by Albrizio et al. (2023) explores these results further. They show that debt holdings increase relatively more for high-MRPK firms, which would also be in line with our model predictions. Furthermore, they document that

³⁵If we assume a capital depreciation rate δ of 10% as in Ottonello and Winberry (2020), our estimates imply that after a 1 p.p. expansionary monetary policy shock, average investment increases 19%. This is very close to the 20% found by these authors.

common proxies for tighter financial frictions, such as firm age, leverage, or liquidity, matter for investment sensitivity to monetary policy only as long as firms have a high-MRPK.

Aggregate productivity. In the model, the individual investment decisions aggregate up to changes in misallocation, such that aggregate TFP increases after an expansionary monetary policy shock. To test this prediction quantitatively, we need an empirical measure of TFP that abstracts from any changes in TFP that are brought about by anything but changes in the allocation of capital. For this purpose we define *dynamic weighted average MRPK*, $WAM_{t,\tau}$, as

$$WAM_{t,\tau} \equiv \sum_{j=0}^J MRPK_t^j \frac{k_{j,t+\tau}}{K_{t+\tau}},$$

where j indexes the firm, J is the number of firms, and $K_{t+\tau}$ is the aggregate capital. We approximate the growth rate of $WAM_{t,\tau}$ from time t to $t + \tau$ by the log difference $\Delta \log WAM_{t,\tau} \equiv \log WAM_{t,\tau} - \log WAM_{t,0}$. $\Delta WAM_{t,\tau}$ tells us how much the economy-wide average MRPK has changed from period t to $t + \tau$ *only* due to changes in the distribution of capital across firms, holding constant the MRPK of the firm at the initial level. As we show in Appendix A.6, in our model $\Delta \log WAM_{t,\tau}$ is approximately proportional to the growth rate of TFP Z_t :

$$\Delta \log WAM_{t,\tau} \approx 1/\alpha \Delta \log Z_{t,\tau}$$

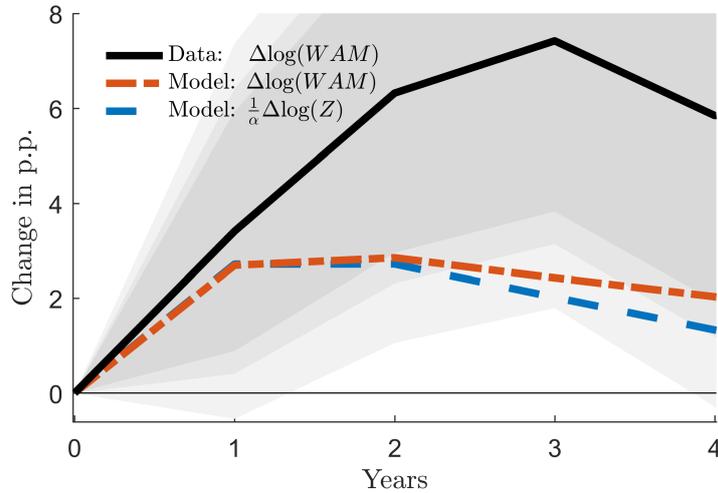
Through the lens of the model, our empirical measure $\Delta \log WAM_{t,\tau}$ can hence be interpreted as a measure of changes in TFP that are brought about purely through changes in the allocation of capital, muting any other channels through which a monetary policy shock may simultaneously affect standard measures of TFP.

We use this variable as dependent variable in the following simple local projection:

$$\Delta \log WAM_{t-1,\tau,s} = \alpha_{s,\tau} + \beta_\tau \varepsilon_t + u_{t,\tau,s} \text{ for } \tau = 1, 2, 3, 4.$$

We estimate this regression at the sector level s to account for potential sectoral differences. This also accounts for differences in capital shares across sectors. $\alpha_{s,\tau}$ denotes horizon specific sector fixed effects, ε_t is the monetary policy shock, and $u_{t,\tau,s}$ is the residual. The regression coefficient β_τ thus tells us the cumulative change in our mea-

Figure 9: Response of average MRPK to an expansionary monetary policy shock.



Notes: The Figure shows the estimated impulse response function after a 1 p.p. expansionary monetary policy shock of $\Delta \log WAM_{t,\tau}$ on the data (black line), and the shaded area marks the 90%, 95% and 99% confidence intervals of the data estimates. It also shows response after a 1 p.p. expansionary monetary policy shock in the model of $\Delta \log WAM_{t,\tau}$ (orange broken line), and the log changes of model TFP (scaled by $1/\alpha$) (blue dashed line).

sure of capital misallocation at different horizons τ after a 1 p.p. monetary policy easing surprise.

Figure 6 reports our estimates for β_τ at different horizons τ (black line), with confidence intervals shaded in gray. Standard error are clustered at sector level. A 1 p.p. expansionary monetary policy shock causes an *increase* of the dynamic weighted average MRPK of 3 p.p. at impact, and of 7 p.p. at peak after 3 years. The effect is significant throughout 4 years at the 95% level.³⁶

The dashed-dotted orange line in Figure 6 shows that the model produces a similar path for the dynamic weighted average MRPK, albeit of smaller magnitude.³⁷ The model explains about half of the observed increase in $\Delta \log WAM_{t,\tau}$ in the data. The model can hence be interpreted as being conservative with regards to the strength of the capital misallocation channel. The peak increase in $WAM_{t,\tau}$ of almost 2.5 p.p. predicted by the model (dashed-dotted orange line) corresponds to an increase of TFP of 0.87 p.p. (dashed blue line).

³⁶Albrizio et al. (2023), using the sector-level variance of MRPK (see Hsieh and Klenow 2009) as a measure of misallocation for Spanish data, also find that expansionary monetary policy shocks decrease misallocation.

³⁷The model counterpart is constructed by feeding a 1 p.p. monetary policy easing surprise into the model as a temporary deviation from the Taylor Rule, and then computing $\Delta WAM_{t,\tau}$.

7 Conclusions

This paper introduces a tractable model with heterogeneous firms, financial frictions, and nominal rigidities in order to understand the link between monetary policy and capital misallocation, and its policy implications. We calibrate this economy using Spanish firm-level data, and show that it can reproduce fairly well the MRPK distribution in the data. Our model predicts that an expansionary monetary policy shock improves the allocation of capital and thus raises TFP. We call this effect the *capital misallocation channel*. We present empirical evidence supporting this prediction: expansionary policy induces high-MRPK firms to increase their investment relatively more than low-MRPK firms. We analyze optimal monetary policy for a benevolent central bank with commitment. The central bank has a strong time-inconsistent incentive to exploit the capital misallocation channel, engineering a temporary economic expansion to increase TFP at the cost of some inflation. When commitment to a timeless policy rules out this time-inconsistent policy, we find that the optimal policy is price stability.

The paper also makes a methodological contribution. It introduces a new algorithm to compute optimal policies in heterogeneous-agent models. The algorithm leverages the numerical advantages of continuous time and will allow researchers to solve optimal policy in heterogeneous-agent models in an efficient and simple way using Dynare.

The model presented in this paper abstracts from several relevant mechanisms driving firm dynamics, such as endogenous default, size-varying capital constraints, frictions in the labor market, or decreasing returns to scale, among many others. This allows us to provide a clear understanding of the forces linking monetary policy with capital misallocation, as well as highlighting the similarities and differences with the standard New Keynesian model. A natural extension would be to add more of these features to study their impact on the optimal conduct of monetary policy.

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Online appendix

A Empirical Appendix

A.1 Firm level data

The empirical exercise relies on annual firm balance-sheet data from the *Central de Balances Integrada* database (Integrated Central Balance Sheet Data Office Survey). We use an unbalanced panel of firms from 1999 to 2016, since these are the years for which the monetary policy shocks are available. Being a detailed administrative dataset, the main advantage is that it covers the quasi-universe of Spanish firms (see [Almunia et al., 2018](#) for further details on the representativeness of this dataset). We use for our analysis only high quality observations, as defined by the Integrated Central Balance Sheet Data Office.

Our main variable of interest, firm's marginal revenue product of capital (MRPK), is proxied by the log of the ratio of value added over tangible capital.³⁸ We drop firms in the 5% upper tail of the capital-weighted MRPK distribution, so as to focus on firms holding a non-negligible capital share. Variables are deflated using industry price levels to preserve the firms' price-level changes and consider a revenue-based measure of MRPK ([Foster et al., 2008](#)). The capital-weighted MRPK distribution in the data is shown in blue bars in [Figure 3](#).

Our dependent variable, the investment rate (or capital growth), is defined as the difference of firm's tangible capital, in logarithm, between periods t and $t - 1$. We also use other firm-level information as controls in the robustness Section below. Profits are computed as net ordinary profits, defined as value added minus personnel costs, net financial revenue and depreciation. Leverage is computed as total debt (short-term plus long-term debt) divided by total assets, and it is trimmed below 0 and over 10. Net financial assets are constructed as the log difference between financial assets and financial liabilities, where financial assets include short-term financial investment, trade receivables, inventories and cash holdings; and financial liabilities include short-term debt, trade payables and long-term debt. We trim this variable below at -10 and above at 10. This variable controls for firms' savings, following [Armenter and Hnatkovska \(2017\)](#). We proxy for size using the logarithm of total assets. Real sales growth is

³⁸Implicitly, this restricts our sample to observations with positive value added.

defined as the log-difference of sales in two consecutive years, the previous and the current one.³⁹ We use the production deflator for value added, sales, financial assets and liabilities and the investment price deflator for capital and total assets. Profits are also deflated using the investment price deflator (see more on this in Appendix A.4). The variables used in the regressions are winsorized at 0.5%. Descriptive statistics are reported in Table 3.

Table 3: **Descriptive statistics**

	mean	sd	p5	p95
Capital growth (1 period)	-0.00	0.29	-0.30	0.49
Net operating profit growth (1 period)	-0.01	1.11	-1.89	1.84
MRPK (logs)	-0.87	1.39	-3.57	0.73
MRPK (levels)	0.77	0.66	0.03	2.07
Total Assets	6.02	1.57	3.56	8.64
Leverage	0.31	0.37	0.00	0.97
Net financial Assets	0.07	0.50	-0.71	0.68
Sales growth	0.13	1.41	-0.50	0.82
Observations	5184233			

Notes: The table shows the mean (column 1), standard deviation (column 2), 5th and 95th percentile value (column 3 and 4 respectively) of the main variables used in the calibration and empirical analysis. MRPK is shown in logs and in levels. The table also displays total assets (logs), leverage, net financial assets; and the log difference of the capital stock (capital growth), output (sales growth) and profits (profit growth). The number of observations are those for which the variable MRPK is available.

A.2 Proxying MRPK and accounting for sectoral differences in the capital share

A firm's MRPK is not directly observable. However, in a Cobb-Douglas production framework, such as the one presented in Section 2, a firm's MRPK is proportional to its average revenue product of capital (ARPK):

$$MRPK_t \equiv \alpha k_t^{\alpha-1} l_t^{1-a} \propto ARPK_t \equiv y_t/k_t = k_t^{\alpha-1} l_t^{1-a}$$

We thus use the easily measurable ARPK as an empirical measure for the unobservable MRPK, following the literature (see for instance [Bau and Matray, 2023](#)). To

³⁹Both sales growth and capital growth are winsorized at 0.5%.

account for the use of intermediate inputs, which we don't model explicitly, we use value added (sales minus intermediate inputs) instead of sales.

However, capital shares α may differ across sectors. This would imply that the ARPK is no longer a valid proxy for the MRPK in cross sectoral comparisons. In the following we explain that all our results are robust to this concern. We use our MRPK proxy on four occasions. We discuss each of them in turn.

Steady state MRPK distribution In Section 3 we show that the model predicts a steady state MRPK distribution that is in line with the MRPK distribution implied by the data, assuming a uniform capital share for all sectors. Here we relax this assumption and allow for sectoral difference in the capital shares. Following Hsieh and Klenow (2009) and Gopinath et al. (2017), we take the sectoral capital shares of a relatively undistorted economy such as the United States. As Figure 10 shows, the fit of the model worsens only slightly in the direction of underpredicting the measured misallocation. The baseline calibration can hence be considered conservative.

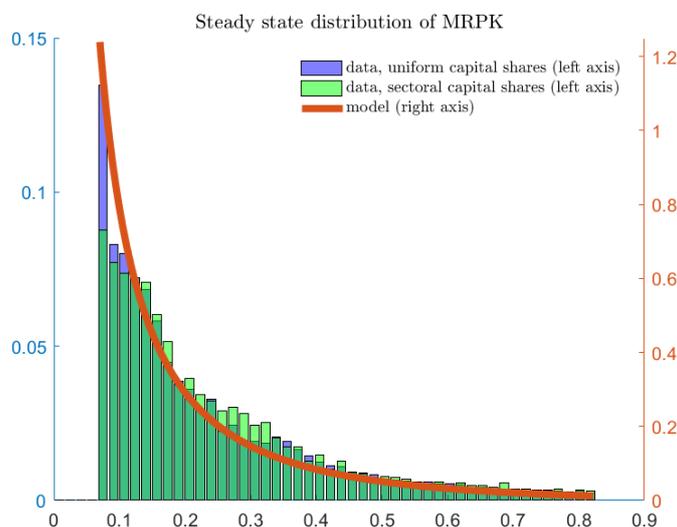


Figure 10: MRPK distribution

Firm level capital growth In Section 6 we show that high MRPK firms respond more strongly to monetary policy shocks. The first robustness we perform (column 2 of Table 2) includes sector-time fixed effects γ_{st} . These soak up any differences between the MRPK and the ARPK that sectoral differences in factor shares could introduce.

To show this, we reproduce our specification here, acknowledging explicitly that we use the ARPK as a proxy for the MRPK and adding a sector index

$$\log k_{j,t} - \log k_{j,t-1} = \beta_0 + \beta_1 \log (ARPK_{j,s,t-1}) + \beta_2 \log (ARPK_{j,s,t-1}) \varepsilon_t + \gamma_{s,t} + u_{j,t},$$

Using the above relationship between the MRPK and the ARPK $ARPK_{j,s,t-1} = \frac{MRPK_{j,t-1}}{\alpha_s}$, where α_s denotes the sectoral capital share we can rewrite this equation as

$$\begin{aligned} \log k_{j,t} - \log k_{j,t-1} &= \beta_0 + \beta_1 \log \left(\frac{MRPK_{j,t-1}}{\alpha_s} \right) + \beta_2 \log \left(\frac{MRPK_{j,t-1}}{\alpha_s} \right) \varepsilon_t + \gamma_{s,t} + u_{j,t} \\ &= \beta_0 + \beta_1 \log (MRPK_{j,t-1}) + \beta_2 \log (MRPK_{j,t-1}) \varepsilon_t \\ &\quad + \{ \beta_1 \log (\alpha_s) + \beta_2 \log (\alpha_s) \varepsilon_t \} + \gamma_{s,t} + u_{j,t}. \end{aligned}$$

Thus, the sector time-fixed effects γ_{st} absorb the differences in sectoral capital shares (the term in curly brackets) and the coefficients β_1 (β_2) can be interpreted as the association of MRPK (the interaction of MRPK with the monetary policy shock) with capital growth, as we do in the main text. ⁴⁰

Furthermore, note that in Appendix A.2 we include a specification with firm-level fixed effects and MRPK demeaned at the firm level (following [Ottonello and Winberry \(2020\)](#)), which accounts for differences in the capital shares even at the firm level.

Dynamic weighted average MRPK In Section 6 we show that the log difference of dynamic weighted average MRPK ($\Delta \log WAM_{t-1,\tau,s}$) increases in response to a monetary policy shock. This variable is constructed by aggregating our MPRK-proxy at the sectoral level. Since we consider its log difference, sectoral differences in capital share α_s wash out. Thus, sectoral differences in capital shares do not affect our results or their interpretation.

Productivity process In Appendix A.5 we use our MRPK proxy to estimate the productivity process. To account for potential sectoral differences in capital shares, here we redo the estimation adding sector fixed effects (which absorb sectoral capital shares since the specification is in logs) or accounting for US sectoral capital shares as previously explained. The estimates change only marginally from $\rho_z = 0.83$ and

⁴⁰Following the same reasoning, even time variation in the sectoral capital shares is accounted for.

$\sigma = 0.73$ to $\rho_z = 0.77$ and $\sigma = 0.71$, and $\rho_z = 0.80$ and $\sigma = 0.72$ respectively.⁴¹

A.3 Monetary policy shocks

We use the monetary policy shocks constructed by [Jarociński and Karadi \(2020\)](#). The key idea behind their identification strategy is that movements of interest rates and stock markets within a narrow window around monetary policy announcements can help disentangle monetary policy shocks from information surprises. While an unexpected policy tightening raises interest rates and reduces stock prices, a positive central bank information shock (i.e. unexpected positive assessment of the economic outlook) raises both. Their identification of monetary policy relies hence on sign restrictions: an unexpected monetary policy tightening raises interest rates and reduces stock prices, while an unanticipated positive information shock increases both. Each surprise change in interest rates is hence decomposed into a combination of '*central bank information shocks*' and '*monetary policy shocks*'. We use the latter, as provided by the authors with the published paper.

We use their monetary policy shocks at monthly frequency. Since our firm-level panel is at annual frequency, we aggregate the monthly monetary policy shocks following the scheme of [Ottonello and Winberry \(2020\)](#). However, instead of aggregating daily shocks into quarterly series, we apply a monthly-to-yearly transformation. This scheme accounts for the fact that firms have less time to react to shocks happening at the end of the year than to shocks happening earlier on. In particular, a monthly shock enters both the current year and the following year's annual shock, with the split between the current and the next year depending on the timing of the monthly shock within the current year.⁴² Concretely, we construct the monetary policy shock as

$$\varepsilon_t = \sum_m \omega_{\text{past}}(m) \varepsilon_{m,t-1} + \sum_m \omega_{\text{current}}(m) \varepsilon_{m,t} \quad \omega_{\text{past}}(m) = \frac{m-1}{12}, \quad \omega_{\text{current}}(m) = \frac{12-(m-1)}{12}$$

where ε_t is the aggregated annual monetary policy shock in year t , and $\varepsilon_{m,t}$ is the high-frequency shock in month $m = 1, \dots, 12$ of year t . See [\(Albrizio et al., 2023\)](#) for a formal derivation of this weighting scheme. Note that we multiply the original shocks

⁴¹Even firm level fixed effects do not change the estimates much.

⁴²For instance, a high-frequency surprise happening in January is entirely attributed to the current year, while the one occurring in December mainly contributes to the following year's annual shock.

by (-1), so that positive monetary policy shocks corresponds to expansionary monetary policy. Figure 11 shows the time series of the shock.

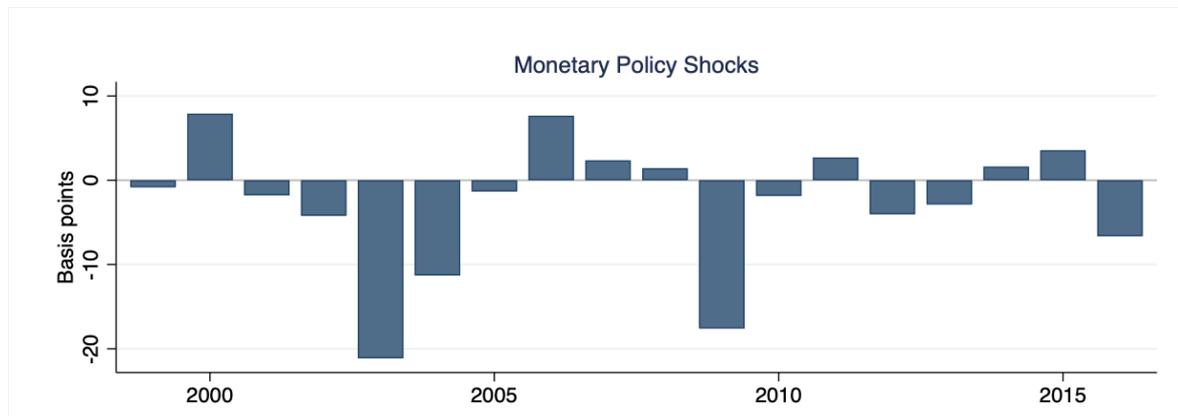


Figure 11: Monetary policy shocks at annual frequency.

Source: Jarociński and Karadi (2020) and own calculations.

A.4 Robustness of the firm level regression

In this Section we perform several robustness of our finding that high-MRPK firms' investment responds more to monetary shocks. We consider variations of the main empirical specification explained in the main text, equation (41), which we repeat here expanded to include the robustness specifications we perform below

$$\log k_{j,t} - \log k_{j,t-1} = \beta_0 + \beta_1 \log(MRPK_{j,t-1}) + \beta_2 \log(MRPK_{j,t-1}) \varepsilon_t + \lambda' Z_{j,t-1} + \gamma_{st} + \kappa_j + u_{j,t}, \quad (42)$$

where $Z_{j,t-1}$ includes a vector of lagged firm-level controls (total assets, sales growth, leverage, capital growth and net short term financial assets), and in some specification it also includes their interaction with the monetary policy shock; γ_{st} are sector-year fixed effects; and κ_j are firm-level fixed effects.

Column (1) in Table 4 reproduces the results of Column 2 of Table 2 of the main text. It does not include firm-level nor aggregate controls, and it only includes sector-time fixed effects. Column (2) includes firm fixed-effects, and Column (3) also adds firm-level controls (the lag of total assets, sales growth, leverage, capital growth and net short term financial assets). The results remain positive, significant and of similar magnitude. Column (4) reports results for the same specifications as Column (3), but adding the interaction of $\log(MRPK_{j,t-1})$ with lagged GDP growth, to rule out any heterogeneity in response to business cycle movements. Column (5) adds to the

specification of column (4) the interaction of the monetary policy shocks with the firm level controls. Results remain significant and quantitatively similar. Column (6) runs the same specification of Column (5), just replacing the main variable of interest $\log(MRPK_{j,t-1})$ with its demeaned value at the firm level, to make sure that results are not driven by permanent heterogeneity in MRPK levels, in the spirit of [Ottonello and Winberry, 2020](#). Our results also survive to this strict specification.

One of the advantages of our dataset is that it also includes small and privately held firms. But precisely because of this, it could be the case that small firms are the ones driving these results and one may wonder if the same empirical pattern holds for large firms. To address this concern, we replicate the analysis of our baseline specification with sector-year fixed effects, Column (1), but keeping only firms below the 90th percentile of employment (Column 7), and keeping only firms above the 90th percentile (Column 8). The coefficient on the slope of MRPK is positive, significant, and even quantitatively larger for larger firms. The 90th percentile of employment in the Spanish distribution of firms is relatively low (15 employees), so we repeat the same regression but keeping only firms with at least 100 employees in Column (9), and reach the same conclusion. Since our panel is highly unbalanced, we run our baseline specification, but only for firms that we observe for at least for 6 consecutive years (from $t - 1$ to $t + 4$) (Column 10), hence restricting the sample as in the aggregate analysis performed in Section 4.2. The coefficient is nearly twice as large as that of Column (1). Summing up, all these exercises point at the robustness of the empirical result of a higher heterogeneous response of investment for high-MRPK firms to a monetary policy shock.

Finally, we want to test directly whether high-MRPK firms' profits are indeed increasing, in line with the theoretical predictions. We use as data counterpart net ordinary net profits deflated by the capital investment deflator.⁴³ We estimate the same equation as that of Column 1 of Table 4, but with the log change in profits on the left-hand side. The results are depicted in Column (11). As predicted by the model, the coefficient is positive and significant, providing further support for the mechanism presented.

⁴³Results are robust to deflate profits using the production deflator rather than the capital investment deflator

Table 4: Robustness

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log k$	$\Delta \log n_{op}$
$\varepsilon * mrpk$	0.0286*** (0.01)	0.0217** (0.01)	0.0262*** (0.01)	0.0270*** (0.01)	0.0344*** (0.01)	0.080** (0.04)	0.0275** (0.01)	0.0607*** (0.02)	0.0679*** (0.02)	0.0450*** (0.02)	0.108*** (0.04)
$\varepsilon * mrpk$											
Observations	3692188	3538432	2641657	2641657	2641657	2641657	3290824	401359	40932	1253505	1970415
R^2	0.020	0.255	0.295	0.295	0.295	0.295	0.020	0.035	0.063	0.033	0.026
Time-sector FE	YES	YES	YES	YES	YES	YES	YES	YES	YES	YES	YES
Firm FE	NO	YES	YES	YES	YES	YES	NO	NO	NO	NO	NO
Firm Controls	NO	NO	YES	YES	YES	YES	NO	NO	NO	NO	NO
Agg. Control	NO	NO	NO	YES	YES	YES	NO	NO	NO	NO	NO
MP shock*Firm Control	NO	NO	NO	NO	YES	YES	NO	NO	NO	NO	NO
MRPK demeaning	NO	NO	NO	NO	NO	YES	NO	NO	NO	NO	NO
Panel	FULL	FULL	FULL	FULL	FULL	FULL	EMP<p90	EMP>p90	LARGE	N > 5	FULL

Notes: This table reports the results of estimating equation (42), departing from some of the specifications of the estimation in the main text. Column (1) includes sector-time fixed effects. Column (2) includes firm fixed-effects, and Column (3) also includes lagged firm-level controls (total assets, sales growth, capital growth and net short term financial assets). Column (4) runs the same specification as Column (3), but adding the interaction of $\log(MRPK_{j,t-1})$ and lagged GDP growth. Column (5) adds to the specification of column (4) the interaction of lagged firm level controls and the monetary policy shock. Column (6) runs the same specification as Column (5), just replacing the main variable of interest $\log(MRPK_{j,t-1})$ with its demeaned value at the firm level. Columns (7), (8) and (9) run the baseline specification, but including only firms with less employees than the 90th percentile (15 employees), firms with more employees than 90th percentile, and firms with more than 100 employees, respectively. Column (10) shows the baseline specification, but only for firms that we observe at least for 6 consecutive years (from $t-1$ to $t+4$). Column (11) runs the same specification as that of Column (1), but using the log change of profits as dependent variable. Standard errors are clustered at the sector-year level.

A.5 Estimating the process for idiosyncratic productivity z

We assume that individual productivity z in logs follows an Ornstein-Uhlenbeck process

$$d \log(z) = -\varsigma_z \log(z)dt + \sigma_z dW_t.$$

To estimate this continuous time process on discrete data, we approximate it by an AR(1) process using an Euler-Maruyama approximation

$$\log(z_t^j) = \rho_z \log(z_{t-1}^j) + \varepsilon_t, \quad \varepsilon_t \sim N(0, \sigma_z \sqrt{\Delta t}),$$

where $\rho_z \approx 1 - \varsigma_z \Delta t \approx \exp(-\varsigma_z \Delta t)$.

In the model, firm level productivity z is proportional to firm level MRPK

$$MRPK_t(z) = z\varphi_t$$

Using this we can rewrite the discretized process for z as

$$\log(MRPK_t(z_t^j)/\varphi_t) = \rho_z \log(MRPK_{t-1}(z_{t-1}^j)/\varphi_{t-1}) + \varepsilon_t, \quad \varepsilon_t \sim N(0, \sigma_z)$$

$$\log(MRPK_t(z_t^j)) = \rho_z \log(MRPK_{t-1}(z_{t-1}^j)) + f(\varphi_t, \varphi_{t-1}) + \varepsilon_t,$$

We estimate this equation using OLS on our panel data specified above, capturing the term $f(\varphi_t, \varphi_{t-1})$ by using year fixed effects. We find $\rho_z = 0.83$, and the standard deviation of the shock is $\sigma = 0.73$. This estimate is robust to including sector fixed effects to account for sectoral differences in capital shares (see Appendix A.2). As the data frequency is annual, $\Delta t = 1$, we back out the implied to continuous time parameter $\varsigma_z = -\log(\rho_z) = 0.189$.

A.6 Derivation of the approximate correspondence of ΔZ_t and $\Delta WAM_{t,s}$

We define $k_t(z) = \int_0^\infty k(z, a) g_t(z, a) da$, and $a_t(z) = \int_0^\infty a g_t(z, a) da$. Manipulating the definition of TFP (32) by subsequently using the definitions of $\Omega(z_t^*)$, $\omega_t(z)$ and the

linearity of $k_t(z)$ in $a_t(z)$ when $z > z_t^*$, we get

$$\begin{aligned}
Z_t^{1/\alpha} &= \frac{\int_{z_t^*}^{\infty} z \omega_t(z) dz}{(1 - \Omega(z_t^*))} dz \\
&= \int_{z_t^*}^{\infty} z \frac{\omega_t(z)}{\int_{z_t^*}^{\infty} \omega_t(z) dz} dz \\
&= \int_{z_t^*}^{\infty} z \frac{a_t(z)/A_t}{\int_{z_t^*}^{\infty} a_t(z)/A_t dz} dz \\
&= \int_0^{\infty} z \frac{k_t(z)}{\int_0^{\infty} k_t(z) dz} dz \\
&= \int_0^{\infty} z \frac{k_t(z)}{K_t} dz.
\end{aligned}$$

Now consider two points in time t and $t + \tau$, where $t < t + \tau$. Since z follows a persistent process we can approximate a firm's j productivity level at $t + \tau$ by its productivity at t , $z_{j,t} \approx z_{j,t+\tau}$. This approximation holds exactly in the limit as the process slows down ($\varsigma_z \rightarrow 1$ and $\sigma_z \rightarrow 0$) or as the time difference shrinks ($\tau \rightarrow 0$). We can thus write

$$Z_{t+\tau}^{1/\alpha} \approx \int_0^1 z_{j,t} \frac{k_{j,t+\tau}}{K_{t+\tau}} dj$$

where $k_{j,t+\tau}$ denotes the period $t + \tau$ capital of an active firm with initial productivity level $z_{j,t}$ in period t . Using the definition of the MRPK,

$$MRPK_t(z_{j,t}) \equiv \varphi_t z_{j,t}$$

we arrive at

$$Z_{t+\tau}^{1/\alpha} \approx \frac{1}{\varphi_t} \int_0^1 MRPK_t(z_{j,t}) \frac{k_{j,t+\tau}}{K_{t+\tau}} dj.$$

To understand how a monetary policy shock affects $Z_t^{1/\alpha}$, we are interested in the evolution of $(\log Z_{t+\tau}^{1/\alpha} - \log Z_t^{1/\alpha}) = \frac{1}{\alpha} (\log Z_{t+\tau} - \log Z_t)$ where t now denotes the period of the shock arrival. Using the above relationship and defining $WAM_{t,\tau} \equiv \int_0^1 MRPK_t(z_{j,t}) \frac{k_{j,t+\tau}}{K_{t+\tau}} dj$

we can write:

$$\begin{aligned} \frac{1}{\alpha} (\log Z_{t+\tau} - \log Z_t) &\approx \log \int_0^1 MRPK_t(z_{j,t}) \frac{k_{j,t+\tau}}{K_{t+\tau}} dj - \log \int_0^1 MRPK_t(z_{j,t}) \frac{k_{j,t}}{K_t} dj \\ &= \log WAM_{t,\tau} - \log WAM_{t,0} \end{aligned}$$

The empirical counterpart of $\int_0^1 MRPK_t(z_{j,t}) \frac{k_{j,t+\tau}}{K_{t+\tau}} dj$ is the expression in the main text $\sum_0^J MRPK_t^j \frac{k_{t+\tau}^j}{K_{t+\tau}}$. In the main text we report both the $\frac{1}{\alpha} (\log Z_{t+\tau} - \log Z_t)$ and $\log WAM_{t,\tau} - \log WAM_{t,0}$.⁴⁴ The approximation is indeed good, especially for the first years, as we show in Figure 6.

⁴⁴ $WAM_{t,\tau}$ can be computed analytically given the transitional dynamics of equilibrium prices and z_t^* .

B Further details on the model

B.1 Entrepreneur's intertemporal problem

The Hamilton-Jacobi-Bellman (HJB) equation of the entrepreneur is given by

$$r_t V_t(z, a) = \max_{d_t \geq 0} d_t + s_t^a(z, a, d) \frac{\partial V}{\partial a} + \mu(z) \frac{\partial V}{\partial z} + \frac{\sigma^2(z)}{2} \frac{\partial^2 V}{\partial z^2} + \eta(q_t a_t - V_t(z, a)) + \frac{\partial V}{\partial t}.$$

We guess and verify a value function of the form $V_t(z, a) = \kappa_t(z) q_t a$. The first order condition is

$$\kappa_t(z) - 1 = \lambda_d \text{ and } \min\{\lambda_d, d_t\} = 0,$$

where $\lambda_d = 0$ if $\kappa_t(z) = 1$. If $\kappa_t(z) > 1 \forall z, t$, then $d_t = 0$ and the firm does not pay dividends until it closes down. If this is the case, then the value of $\kappa_t(z)$ can be obtained from

$$\begin{aligned} (r_t + \eta) \kappa_t(z) q_t = \\ \eta q_t + (\gamma \max\{z_t \varphi_t - R_t, 0\} + R_t - \delta q_t) \kappa_t(z) + \mu(z) q_t \frac{\partial \kappa_t}{\partial z} + \frac{\sigma^2(z)}{2} q_t \frac{\partial^2 \kappa_t}{\partial z^2} + \frac{\partial (q_t \kappa_t)}{\partial t}. \end{aligned} \quad (43)$$

Lemma. $\kappa_t(z) > 1 \forall z, t$

Proof. The drift of the entrepreneur's capital holdings is

$$s_t^a = \frac{1}{q_t} [(\gamma \max\{z_t \varphi_t - R_t, 0\} + R_t - \delta q_t)] \geq \frac{R_t - \delta q_t}{q_t}$$

which is expected to hold with strict inequality eventually if $\exists \mathbb{P}(z_t \geq z_t^*) > 0$ (which is satisfied in equilibrium since z is unbounded), and hence

$$\mathbb{E}_0 a_t = \mathbb{E}_0 a_0 e^{\int_0^t s_u^a du} > a_0 e^{\int_0^t \frac{R_s - \delta q_s}{q_s} ds}. \quad (44)$$

The value function is then

$$\begin{aligned} \kappa_{t_0}(z) q_{t_0} a_{t_0} &= V_{t_0}(z, a_{t_0}) \\ &= \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t (r_s + \eta) ds} (d_t + \eta q_t a_t) dt \end{aligned}$$

$$\begin{aligned}
&\geq \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t (r_s + \eta) ds} \eta q_t a_t dt = \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t \left(\overbrace{R_s - \delta q_s + \dot{q}_s}^{r_s} + \eta \right) ds} \eta q_t a_t dt \\
&= \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t \left(\frac{R_s - \delta q_s}{q_s} + \eta \right) ds - \log \frac{q_t}{q_{t_0}}} \eta q_t a_t dt = \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t \left(\frac{R_s - \delta q_s}{q_s} + \eta \right) ds} \eta q_{t_0} a_t dt \\
&> \mathbb{E}_{t_0} \int_0^\infty e^{-\int_{t_0}^t \left(\frac{R_s - \delta q_s}{q_s} + \eta \right) ds} \eta q_{t_0} a_{t_0} e^{\int_0^t \frac{R_s - \delta q_s}{q_s} ds} dt = \int_0^\infty e^{-\eta t} \eta q_{t_0} a_{t_0} dt = q_{t_0} a_{t_0},
\end{aligned}$$

where in the first equality we have employed the linear expression of the value function, in the second equation (5), in the third the fact that dividends are non-negative, in the fourth the definition of the real rate (17) and in the last line the inequality (44). Hence $\kappa_{t_0}(z) > 1$ for any t_0 .

B.2 Household's problem

We can rewrite the household's problem as

$$W_t = \max_{C_t, L_t, D_t, B_t^N, S_t^N} \mathbb{E}_0 \int_0^\infty e^{-\rho_t^h t} \left(\frac{C_t^{1-\zeta}}{1-\zeta} - \Upsilon \frac{L_t^{1+\vartheta}}{1+\vartheta} \right) dt. \quad (45)$$

$$s.t. \quad \dot{D}_t = [(R_t - \delta q_t) D_t + w_t L_t - C_t - S_t^N + \Pi_t] / q_t, \quad (46)$$

$$\dot{B}_t^N = S_t^N + (i_t - \pi_t) B_t^N, \quad (47)$$

where S_t^N is the investment into nominal bonds.

The Hamiltonian is

$$\begin{aligned}
H &= \left(\frac{C_t^{1-\zeta}}{1-\zeta} - \Upsilon \frac{L_t^{1+\vartheta}}{1+\vartheta} \right) \\
&+ \varrho_t [((R_t - \delta q_t) D_t + w_t L_t - C_t - S_t^N + (q_t \nu_t - \iota_t - \Phi(\nu_t)) K_t + \Pi_t) / q_t] + \eta_t [S_t^N + (i_t - \pi_t) B_t^N]
\end{aligned}$$

The first order conditions are

$$C_t^{-\zeta} - \varrho_t / q_t = 0 \quad (48)$$

$$-\Upsilon L_t^\vartheta + \varrho_t w_t / q_t = 0 \quad (49)$$

$$-\varrho_t / q_t + \eta_t = 0 \quad (50)$$

$$\dot{\varrho}_t = \rho_t^h \varrho_t - \varrho_t (R_t - \delta q_t) / q_t \quad (51)$$

$$\dot{\eta}_t = \rho_t^h \eta_t - \eta_t [(i_t - \pi_t)] \quad (52)$$

(48) and (49) combine to the optimality condition for labor

$$w_t = \frac{L_t^\vartheta}{C_t^{1-\eta}},$$

(48) can be rewritten as

$$\varrho_t = C_t^\eta q_t$$

Now take derivative with respect to time

$$\dot{\varrho}_t = -\eta C_t^{\eta-1} \dot{C}_t q_t + C_t^\eta \dot{q}_t$$

and plug this into (51) and rearrange to get the first Euler equation

$$\frac{\dot{C}_t}{C_t} = \frac{R_t - \delta q_t + \dot{q}_t}{q_t} - \rho_t^h$$

(50) can be rewritten as

$$\eta_t = \varrho_t / q_t$$

Now take derivative with respect to time

$$\dot{\eta}_t = \frac{\dot{\varrho}_t q_t - \varrho_t \dot{q}_t}{q_t^2}$$

Use these two expressions and the definition of $\dot{\varrho}_t$ in (52) to get the second Euler equation

$$\frac{\dot{C}_t}{C_t} = \frac{(i_t - \pi_t) - \rho_t^h}{\eta}$$

Combining the two Euler equations, we get the Fisher equation

$$\frac{R_t - \delta q_t + \dot{q}_t}{q_t} = (i_t - \pi_t)$$

Finally using the definition of $r_t \equiv \frac{R_t - \delta q_t + \dot{q}_t}{q_t}$ we can rewrite the first Euler equation and the Fisher equation as in the main text.

B.3 New Keynesian Philips curve

The proof is similar to that of Lemma 1 in [Kaplan et al. \(2018\)](#). The Hamilton-Jacobi-Bellman (HJB) equation of the retailer's problem is

$$r_t V_t^r(p) = \max_{\pi} \left(\frac{p - p_t^y(1 - \tau)}{P_t} \right) \left(\frac{p}{P_t} \right)^{-\varepsilon} Y_t - \frac{\theta}{2} \pi^2 Y_t + \pi p \frac{\partial V^r}{\partial p} + \frac{\partial V^r}{\partial t},$$

where where $V_t^r(p)$ is the real value of a retailer with price p . The first order and envelope conditions for the retailer are

$$\begin{aligned} \theta \pi Y_t &= p \frac{\partial V^r}{\partial p}, \\ (r - \pi) \frac{\partial V^r}{\partial p} &= \left(\frac{p}{P_t} \right)^{-\varepsilon} \frac{Y_t}{P_t} - \varepsilon \left(\frac{p - p_t^y(1 - \tau)}{P_t} \right) \left(\frac{p}{P_t} \right)^{-\varepsilon - 1} \frac{Y_t}{P_t} + \pi p \frac{\partial^2 V^r}{\partial p^2} + \frac{\partial^2 V^r}{\partial t \partial p}. \end{aligned}$$

In a symmetric equilibrium we will have $p = P$, and hence

$$\begin{aligned} \frac{\partial V^r}{\partial p} &= \frac{\theta \pi Y_t}{p}, \\ (r - \pi) \frac{\partial V^r}{\partial p} &= \frac{Y_t}{p} - \varepsilon \left(\frac{p - p_t^y(1 - \tau)}{p} \right) \frac{Y_t}{p} + \pi p \frac{\partial^2 V^r}{\partial p^2} + \frac{\partial^2 V^r}{\partial t \partial p}. \end{aligned} \tag{53}$$

Deriving (53) with respect to time gives

$$\pi p \frac{\partial^2 V^r}{\partial p^2} + \frac{\partial^2 V^r}{\partial t \partial p} = \frac{\theta \pi \dot{Y}}{p} + \frac{\theta \dot{\pi} Y}{p} - \frac{\theta \pi^2 \dot{Y}}{p},$$

and substituting into the envelope condition and dividing by $\frac{\theta Y}{p}$ we obtain

$$\left(r - \frac{\dot{Y}}{Y} \right) \pi = \frac{1}{\theta} \left(1 - \varepsilon \left(1 - \frac{p_t^y(1 - \tau)}{p} \right) \right) + \dot{\pi}.$$

Finally, rearranging we obtain the New Keynesian Phillips curve

$$\left(r - \frac{\dot{Y}}{Y}\right) \pi = \frac{\varepsilon}{\theta} \left(\frac{1 - \varepsilon}{\varepsilon} + \tilde{m}\right) + \dot{\pi}.$$

The total profit of retailers, net of the lump-sum tax, which is transferred to the households lump sum, is

$$\Pi_t = (1 - m_t) Y_t - \frac{\theta}{2} \pi_t^2 Y_t. \quad (54)$$

B.4 Capital producers' problem

The problem of the capital producer is

$$W_t = \max_{\iota_t, K_t} \mathbb{E}_0 \int_0^\infty e^{-\int_0^t r_s ds} (q_t \iota_t - \iota_t - \Xi(\iota_t)) K_t dt. \quad (55)$$

$$\dot{K}_t = (\iota_t - \delta) K_t, \quad (56)$$

We construct the Hamiltonian

$$H = (q_t \iota_t - \iota_t - \Xi(\iota_t)) K_t + \lambda_t (\iota_t - \delta) K_t$$

with first-order conditions

$$(q_t - 1 - \Xi'(\iota_t)) + \lambda_t = 0 \quad (57)$$

$$(q_t \iota_t - \iota_t - \Xi(\iota_t)) + \lambda_t (\iota_t - \delta) = r_t \lambda_t - \dot{\lambda}_t \quad (58)$$

Taking the time derivative of equation (57)

$$\dot{\lambda}_t = -(\dot{q}_t - \Xi''(\iota_t) \dot{\iota}_t)$$

which, combined with (58), yields

$$(q_t \iota_t - \iota_t - \Xi(\iota_t)) - (q_t - 1 - \Xi'(\iota_t)) (\iota_t - \delta - r_t) = (\dot{q}_t - \Xi''(\iota_t) \dot{\iota}_t)$$

Rearranging we get

$$r_t = (\iota_t - \delta) + \frac{\dot{q}_t - \Xi''(\iota_t) \iota_t}{q_t - 1 - \Xi'(\iota_t)} - \frac{q_t \iota_t - \iota_t - \Xi(\iota_t)}{q_t - 1 - \Xi'(\iota_t)}.$$

B.5 Distribution

The joint distribution of net worth and productivity is given by the Kolmogorov Forward equation

$$\frac{\partial g_t(z, a)}{\partial t} = -\frac{\partial}{\partial a}[g_t(z, a)s_t(z)a] - \frac{\partial}{\partial z}[g_t(z, a)\mu(z)] + \frac{1}{2}\frac{\partial^2}{\partial z^2}[g_t(z, a)\sigma^2(z)] - \eta g_t(z, a) + \eta/\psi g_t(z, a/\psi), \quad (59)$$

where $1/\psi g_t(z, a/\psi)$ is the distribution of entry firms.

To characterize the law of motion of net-worth shares, defined as $\omega_t(z) = \frac{1}{A_t} \int_0^\infty a g_t(z, a) da$, first we take the derivative of $\omega_t(z)$ wrt time

$$\frac{\partial \omega_t(z)}{\partial t} = -\frac{\dot{A}_t}{A_t^2} \int_0^\infty a g_t(z, a) da + \frac{1}{A_t} \int_0^\infty a \frac{\partial g_t(z, a)}{\partial t} da. \quad (60)$$

Next, we plug in the derivative of $g_t(z, a)$ wrt time from equation(59) into equation (60),

$$\begin{aligned} \frac{\partial \omega_t(z)}{\partial t} &= -\frac{\dot{A}_t}{A_t^2} \int_0^\infty a g_t(z, a) da + \frac{1}{A_t} \int_0^\infty a \left(-\frac{\partial}{\partial a}[g_t(z, a)s_t(z)a] \right) da \\ &\quad - \frac{\partial}{\partial z} \mu(z) \frac{1}{A_t} \int_0^\infty a g_t(z, a) da + \frac{1}{2} \frac{\partial^2}{\partial z^2} \sigma^2(z) \frac{1}{A_t} \int_0^\infty a g_t(z, a) da \\ &\quad - \frac{1}{A_t} \int_0^\infty \eta a g_t(z, a) da + \frac{1}{A_t} \int_0^\infty \eta a/\psi g_t(z, a/\psi) da. \end{aligned}$$

Using integration by parts and the definition of net worth shares, we obtain the second order partial differential equation that characterizes the law of motion of net-worth shares,

$$\frac{\partial \omega_t(z)}{\partial t} = \left[s_t(z) - \frac{\dot{A}_t}{A_t} - (1 - \psi)\eta \right] \omega_t(z) - \frac{\partial}{\partial z} \mu(z) \omega_t(z) + \frac{1}{2} \frac{\partial^2}{\partial z^2} \sigma^2(z) \omega_t(z). \quad (61)$$

The stationary distribution is therefore given by the following second order partial

differential equation,

$$0 = (s(z) - (1 - \psi)\eta)\omega(z) - \frac{\partial}{\partial z}\mu(z)\omega(z) + \frac{1}{2}\frac{\partial^2}{\partial z^2}\sigma^2(z)\omega(z). \quad (62)$$

B.6 Market clearing and aggregation

Define the cumulative function of net-worth shares as

$$\Omega_t(z) = \int_0^z \omega_t(z)dz. \quad (63)$$

Using the optimal choice for k_t from equation (7), we obtain

$$K_t = \int k_t(z, a)dG_t(z, a) = \int_{z_t^*}^{\infty} \int \gamma a \frac{1}{A_t} g_t(z, a) da dz A_t = \gamma(1 - \Omega(z_t^*))A_t. \quad (64)$$

By combining equations (27), (28) and (64), and solving for A_t , we obtain

$$A_t = \frac{D_t}{\gamma(1 - \Omega(z_t^*)) - 1}, \quad (65)$$

Labor market clearing implies

$$L_t = \int_0^{\infty} l_t(z, a)dG_t(z, a). \quad (66)$$

Define the following auxiliary variable,

$$X_t \equiv \int_{z_t^*}^{\infty} z\omega_t(z)dz = \mathbb{E}[z \mid z > z_t^*](1 - \Omega(z_t^*)). \quad (67)$$

Using labor demand from (8), X_t and using the definition of φ_t , we obtain

$$L_t = \int_0^{\infty} \left(\frac{\varphi_t}{\alpha m_t}\right)^{\frac{1}{1-\alpha}} z_t \gamma a_t dG_t(z, a) = \left(\frac{\varphi_t}{\alpha m_t}\right)^{\frac{1}{1-\alpha}} \gamma A_t X_t. \quad (68)$$

Plugging in (8) into production function (1), and using again the definition of shares, we obtain

$$Y_t = \int \underbrace{\frac{z_t \varphi_t}{\alpha m_t}}_{y_t(z, a)} \gamma a dG_t(z, a) = \frac{\varphi_t}{\alpha m_t} X_t \gamma A_t = Z_t A_t^\alpha L_t^{1-\alpha}, \quad (69)$$

where in the last equality we have used equation (68), and we have defined

$$Z_t = (\gamma X_t)^\alpha. \quad (70)$$

Aggregate profits of retailers are given by

$$\Phi_t^{Agg} = \int \gamma \max \{z_t \varphi_t - R_t, 0\} a_t dG_t(z, a) = [\varphi_t X_t - R_t (1 - \Omega(z^*))] \gamma A_t. \quad (71)$$

We can also write the aggregate production in terms of physical capital,

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha}, \quad (72)$$

where the TFP term Z_t is defined as

$$Z_t = \left(\frac{X_t}{(1 - \Omega(z_t^*))} \right)^\alpha = (\mathbb{E}[z \mid z > z_t^*])^\alpha. \quad (73)$$

$$Z_t^{1/\alpha} (1 - \Omega(z_t^*)) = (X_t) = (\mathbb{E}[z \mid z > z_t^*])^\alpha. \quad (74)$$

Aggregating the budget constraint of all input good firms, using the linearity of savings policy (11) and using (65), we obtain

$$\begin{aligned} \dot{A}_t &= \int \dot{a} dG(z, a, t) - \eta \int (1 - \psi) a_t dG(z, a, t) = \\ &= \int_0^\infty \frac{1}{q_t} (\gamma \max \{z_t \varphi_t - R_t, 0\} + R_t - \delta q_t - q_t (1 - \psi) \eta) a_t dG(z, a), \end{aligned}$$

Dividing by A_t both sides of this equation, using the definition of net worth shares and the fact that these integrate up to one, we obtain

$$\frac{\dot{A}_t}{A_t} = \frac{1}{q_t} (\gamma \varphi_t X_t - R_t \gamma (1 - \Omega(z_t^*)) + R_t - \delta q_t - q_t (1 - \psi) \eta). \quad (75)$$

Using the definition of X_t , and substituting φ_t using equation (68), we can simplify equation (75) as

$$\frac{\dot{A}_t}{A_t} = \frac{1}{q_t} (1 - \Omega(z_t^*)) \gamma (\alpha m_t Z_t L_t^{1-\alpha} ((1 - \Omega(z_t^*)) \gamma A_t)^{\alpha-1} - R_t) + R_t - \delta q_t - q_t (1 - \psi) \eta. \quad (76)$$

Using (64) and (65) we can replace $(1 - \Omega(z_t^*))\gamma A_t$ by K_t , which delivers equation (35).

Finally, we can obtain factor prices

$$w_t = (1 - \alpha)m_t Z_t A_t^\alpha L_t^{-\alpha} \quad (77)$$

$$R_t = \alpha m_t Z_t A_t^{\alpha-1} L_t^{1-\alpha} \frac{z_t^*}{\gamma X_t} \quad (78)$$

where wages come from substituting the definition of φ_t into equation (68); and interest rates come from plugging in the wage expression (77) into the cut-off rule (10) and using equation (65). We could equivalently write equation (78) in terms of real rate of return r_t :

$$r_t = \frac{1}{q_t} \left(\alpha m_t Z_t A_t^{\alpha-1} L_t^{1-\alpha} \frac{z_t^*}{\gamma X_t} \right) - \delta + \frac{\dot{q}}{q_t} \quad (79)$$

We can easily get these equations in terms of capital instead of net worth by simply using equation (64), i.e. $A_t = \frac{K_t}{\gamma(1-\Omega(z_t^*))}$, and using that $\mathbb{E}[z | z > z_t^*] = \frac{X_t}{(1-\Omega(z_t^*))} = \frac{\int_{z_t^*}^{\infty} z \omega_t(z) dz}{(1-\Omega(z_t^*))}$ (see equation (70) and (73)).

B.7 Full set of equations

The competitive equilibrium economy is described by the following 22 equations, for the 22 variables $\{\omega(z), w, r, q, \varphi, R, K, A, L, C, D, Z, \mathbb{E}[z | z > z_t^*], \Omega, z^*, \iota, \pi, m, \tilde{m}, i, Y, T\}$. Remember that $\mu(z) = z \left(-\varsigma_z \log z + \frac{\sigma^2}{2} \right)$ and $\sigma(z) = \sigma_z z$, and that government bonds are in zero net supply ($B_t^N = 0$, hence $S_t^N = 0$). Except from the last equation (Taylor rule), the other 21 equations are the constraints of the Ramsey problem described in Section 2.8.

$$\frac{\partial \omega_t(z)}{\partial t} = \left(s_t(z) - (1 - \psi)\eta - \frac{\dot{A}_t}{A_t} \right) \omega_t(z) - \frac{\partial}{\partial z} [\mu(z)\omega_t(z)] + \frac{1}{2} \frac{\partial^2}{\partial z^2} [\sigma^2(z)\omega_t(z)]$$

$$\text{where } s_t(z) \equiv \frac{1}{q_t} (\gamma \max \{z_t \varphi_t - R_t, 0\} + R_t - \delta q_t)$$

$$\Omega_t(z^*) = \int_0^{z^*} \omega_t(z) dz$$

$$\varphi_t = \alpha \left(\frac{(1 - \alpha)}{w_t w_t} \right)^{(1-\alpha)/\alpha} m_t^{\frac{1}{\alpha}}$$

$$\tilde{m}_t = m_t(1 - \tau)$$

$$w_t = (1 - \alpha)m_t Z_t K_t^\alpha L_t^{-\alpha}$$

$$R_t = \alpha m_t Z_t K_t^{\alpha-1} L_t^{1-\alpha} \frac{z_t^*}{\mathbb{E}[z | z > z_t^*]}$$

$$\frac{\dot{A}_t}{A_t} = \frac{1}{q_t} [\gamma(1 - \Omega(z_t^*)) (\alpha m_t Z_t K_t^{\alpha-1} L_t^{1-\alpha} - R_t) + R_t - \delta q_t - q_t(1 - \psi)\eta]$$

$$K_t = A_t + D_t$$

$$\dot{K}_t = (\iota_t - \delta)K_t$$

$$A_t = \frac{D_t}{\gamma(1 - \Omega(z_t^*)) - 1}$$

$$Z_t = (\mathbb{E}[z | z > z_t^*])^\alpha$$

$$\mathbb{E}[z | z > z_t^*] = \frac{\int_{z_t^*}^{\infty} z \omega_t(z) dz}{(1 - \Omega(z_t^*))}$$

$$\frac{\dot{C}_t}{C_t} = \frac{r_t - \rho_t^h}{\eta}$$

$$w_t = \frac{\Upsilon L_t^\vartheta}{C_t^{-\eta}}$$

$$\dot{D}_t = [(R_t - \delta q_t) D_t + w_t L_t - C_t + T_t] / q_t$$

$$r_t = i_t - \pi_t$$

$$r_t = \frac{R_t - \delta q_t + \dot{q}_t}{q_t}$$

$$(q_t - 1 - \Phi'(\iota_t))(r_t - (\iota_t - \delta)) = \dot{q}_t - \Phi''(\iota_t) i_t - (q_t \iota_t - \iota_t - \Phi(\iota_t))$$

$$\left(r_t - \frac{\dot{Y}_t}{Y_t} \right) \pi_t = \frac{\varepsilon}{\theta} (\tilde{m}_t - m^*) + \dot{\pi}_t, \quad m^* = \frac{\varepsilon - 1}{\varepsilon}$$

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha}$$

$$T_t = (1 - m_t) Y_t - \frac{\theta}{2} \pi_t^2 Y_t + q_t (1 - \psi) \eta A_t + \left[\iota_t q_t - \iota_t - \frac{\phi^k}{2} (\iota_t - \delta)^2 \right] K_t$$

$$di = -v \left(i_t - (\rho_t^h + \phi (\pi_t - \bar{\pi}) + \bar{\pi}) \right) dt.$$

B.8 Proofs of subsection 4.2

TFP is given by equation (32)

$$Z_t = \left(\frac{\int_{z_t^*}^{\infty} z \omega_t(z) dz}{\int_{z_t^*}^{\infty} \omega_t(z) dz} \right)^\alpha.$$

We compute the growth rate of TFP

$$\begin{aligned} \frac{1}{Z_t} \frac{dZ_t}{dt} &= \frac{d \log Z_t}{dt} = \alpha \left[\frac{d}{dt} \left(\log \int_{z_t^*}^{\infty} z \omega_t(z) dz \right) - \frac{d}{dt} \left(\log \int_{z_t^*}^{\infty} \omega_t(z) dz \right) \right] \\ &= \alpha \left[\frac{\int_{z_t^*}^{\infty} z \frac{\partial \omega_t(z)}{\partial t} dz - z_t^* \omega_t(z_t^*) \frac{dz_t^*}{dt}}{\int_{z_t^*}^{\infty} z \omega_t(z) dz} + \frac{- \int_{z_t^*}^{\infty} \frac{\partial \omega_t(z)}{\partial t} dz + \omega_t(z_t^*) \frac{dz_t^*}{dt}}{\int_{z_t^*}^{\infty} \omega_t(z) dz} \right], \end{aligned}$$

where the dynamics of the density are

$$\begin{aligned} \frac{\partial \omega_t(z)}{\partial t} &= \left[\underbrace{\frac{\gamma \varphi_t}{q_t} \max \{ (z - z^*), 0 \}}_{\equiv \Phi_t(z)} + \underbrace{\frac{R_t - \delta q_t}{q_t} - \frac{\dot{A}_t}{A_t} - (1 - \psi)\eta}_{\equiv \Xi_t} \right] \omega_t(z) \\ &\quad + \varsigma_z \frac{\partial}{\partial z} (\log(z) \omega_t(z)) + \frac{\sigma_z^2}{2} \frac{\partial^2}{\partial z^2} \omega_t(z). \end{aligned}$$

From there we can analyze two limit cases.

Constant cutoff First, we analyze the case in which the cut-off remains approximately constant. In this case, the growth rate is

$$\frac{1}{Z_t} \frac{dZ_t}{dt} \Big|_{z^*} = \frac{\int_{z^*}^{\infty} z \frac{\partial \omega_t(z)}{\partial t} dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} \frac{\partial \omega_t(z)}{\partial t} dz}{\int_{z^*}^{\infty} \omega_t(z) dz}.$$

We now show that, in this case, (i) prices only influence TFP through changes in the slope of the excess investment rate, $\frac{\gamma \varphi_t}{q_t}$; and (ii) that this response is positive. The derivative of the TFP growth rate with respect to a price or a function of prices x_t is

$$\frac{\partial}{\partial x_t} \frac{d \log Z_t}{dt} \Big|_{z^*} = \frac{\int_{z^*}^{\infty} z \frac{\partial \omega_t(z)}{\partial x_t} dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} \frac{\partial \omega_t(z)}{\partial x_t} dz}{\int_{z^*}^{\infty} \omega_t(z) dz},$$

where

$$\left. \frac{\partial \dot{\omega}_t(z)}{\partial x_t} \right|_{z^*} = \left. \frac{\partial}{\partial x_t} \left(\tilde{\Phi}_t(z) + \tilde{\Xi}_t \right) \right|_{z^*} \omega(z),$$

given the definitions of $\tilde{\Phi}_t(z)$ and $\tilde{\Xi}_t$ above. Then we have:

$$\begin{aligned} \left. \frac{\partial}{\partial x_t} \frac{d \log Z_t}{dt} \right|_{z^*} &= \frac{\int_{z^*}^{\infty} z \frac{\partial \tilde{\Phi}_t(z)}{\partial x_t} \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} \frac{\partial \tilde{\Phi}_t(z)}{\partial x_t} \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz} \\ &+ \underbrace{\frac{\partial \tilde{\Xi}_t}{\partial x_t} \left(\frac{\int_{z^*}^{\infty} z \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz} \right)}_0. \end{aligned}$$

This expression shows how only the excess investment rate $\tilde{\Phi}(z)$ matters to understand the impact of changes in prices on the growth rate of TFP. Conditional on z^* , price changes affect the excess investment rate by affecting its slope $\frac{\gamma \varphi_t}{q_t}$. So the effect of a shock on TFP growth is determined by its effect on $\frac{\gamma \varphi_t}{q_t}$. This proves claim (i).

To prove that an increase in the slope $\frac{\gamma \varphi_t}{q_t}$ increases TFP growth, we compute

$$\begin{aligned} \left. \frac{\partial}{\partial \left(\frac{\gamma \varphi_t}{q_t} \right)} \frac{d \log Z_t}{dt} \right|_{z^*} &= \frac{\int_{z^*}^{\infty} z \frac{\partial \tilde{\Phi}_t(z)}{\partial \left(\frac{\gamma \varphi_t}{q_t} \right)} \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} \frac{\partial \tilde{\Phi}_t(z)}{\partial \left(\frac{\gamma \varphi_t}{q_t} \right)} \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz} \\ &= \frac{\int_{z^*}^{\infty} z(z - z^*) \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} z(z - z^*) \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz}, \end{aligned}$$

To uncover the sign, we analyze the term

$$\frac{\int_{z^*}^{\infty} (z - z^*) z \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} (z - z^*) \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz} = \frac{\int_{z^*}^{\infty} z^2 \omega_t(z) dz}{\int_{z^*}^{\infty} z \omega_t(z) dz} - \frac{\int_{z^*}^{\infty} z \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz}. \quad (80)$$

We define $\bar{\omega}_t(z) \equiv \frac{\omega_t(z)}{\int_{z^*}^{\infty} \omega_t(z) dz} \mathbb{I}_{z > z^*}$ and $\tilde{\omega}_t(z) \equiv \frac{z \omega_t(z)}{\int_{z^*}^{\infty} \omega_t(z) z dz} \mathbb{I}_{z > z^*}$. These are continuous probability density functions over the domain $[z^*, \infty)$, as they are non-negative and sum up to 1. They satisfy the monotone likelihood ratio condition as

$$I(z) = \frac{\tilde{\omega}_t(z)}{\bar{\omega}_t(z)} = z \frac{\int_{z^*}^{\infty} z \omega_t(z) dz}{\int_{z^*}^{\infty} \omega_t(z) dz}$$

is non decreasing. This implies that function $\tilde{\omega}_t(z)$ dominates $\bar{\omega}_t(z)$ first-order stochastically. Hence

$$\frac{\int_{z_t^*}^{\infty} z\omega_t(z)}{\int_{z_t^*}^{\infty} \omega_t(z) dz} = \mathbb{E}_{\bar{\omega}_t(z)}[z] = \int_{z^*}^{\infty} z\bar{\omega}_t(z) dz < \int_{z_t^*}^{\infty} z\tilde{\omega}_t(z) dz = \mathbb{E}_{\tilde{\omega}_t(z)}[z] = \frac{\int_{z_t^*}^{\infty} z^2\omega_t(z)}{\int_{z_t^*}^{\infty} z\omega_t(z) dz}.$$

Therefore, equation (80) is positive. An increase in the slope of the excess investment rate, $\frac{\gamma\varphi_t}{q_t}$, thus increases TFP growth, which proves claim (ii):

$$\frac{\partial}{\partial \left(\frac{\gamma\varphi_t}{q_t}\right)} \cdot \frac{d \log Z_t}{dt} \Big|_{z_t^*} = \frac{\int_{z_t^*}^{\infty} z^2\omega_t(z)}{\int_{z_t^*}^{\infty} z\omega_t(z) dz} - \frac{\int_{z_t^*}^{\infty} z\omega_t(z)}{\int_{z_t^*}^{\infty} \omega_t(z) dz} > 0.$$

Iid shocks Next, we consider the limit of iid shocks, that is, the limit as $\varsigma_z \rightarrow \infty$. In this case, as discussed in [Itskhoki and Moll \(2019\)](#), the distribution $\omega(z)$ is constant and the growth rate of TFP simplifies to

$$\frac{1}{Z_t} \frac{dZ_t}{dt} = \alpha\omega(z_t^*) \frac{\int_{z_t^*}^{\infty} (z - z_t^*)\omega(z) dz}{\int_{z_t^*}^{\infty} \omega(z) dz} \frac{dz_t^*}{\int_{z_t^*}^{\infty} z\omega(z) dz} \frac{dz_t^*}{dt}.$$

Notice that $\alpha\omega(z_t^*) \frac{\int_{z_t^*}^{\infty} (z - z_t^*)\omega(z) dz}{\int_{z_t^*}^{\infty} \omega(z) dz} > 0$ for any value of the cut-off. In this case, the growth rate of TFP depends linearly with the growth rate of the cut-off: if the later increases, so does the former.

Corollary The market clearing condition for capital (29) can be rewritten as

$$\frac{D_t}{A_t} = \gamma \left(\int_{z_t^*}^{\infty} \omega_t(x) dx \right) - 1$$

Take the time derivative on both sides and then manipulate the RHS

$$\frac{\partial \frac{D_t}{A_t}}{\partial t} = \frac{\partial \gamma \int_{z_t^*}^{\infty} \omega_t(x) dx}{\partial t} = \int_{z_t^*}^{\infty} \frac{\partial \omega_t(z)}{\partial t} dz - \omega_t(z_t^*) \frac{\partial z_t^*}{\partial t}.$$

Now take the derivative with respect to the slope $\frac{\gamma\varphi_t}{q_t}$, drawing on the above in the second line, and finally rearrange.

$$\begin{aligned}
\left(\frac{\partial^2 \frac{D_t}{A_t}}{\partial t \partial \frac{\gamma\varphi_t}{q_t}}\right) &= \partial \left(\int_{z_t^*}^{\infty} \frac{\partial \omega_t(z)}{\partial t} dz - \omega_t(z_t^*) \frac{\partial z_t^*}{\partial t} \right) / \partial \left(\frac{\gamma\varphi_t}{q_t} \right) \\
&= \int_{z_t^*}^{\infty} \frac{\partial \tilde{\Phi}_t(z)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)} \omega_t(z) dz - \omega_t(z_t^*) \frac{\partial \left(\frac{\partial z_t^*}{\partial t} \right)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)}. \\
&= \int_{z_t^*}^{\infty} (z - z_t^*) \omega_t(z) dz - \omega_t(z_t^*) \frac{\partial \left(\frac{\partial z_t^*}{\partial t} \right)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)}. \\
&= (\mathbb{E}[z \mid z > z_t^*] - z_t^*) (1 - \Omega(z_t^*)) - \omega_t(z_t^*) \frac{\partial \left(\frac{\partial z_t^*}{\partial t} \right)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)}
\end{aligned}$$

Thus $\frac{\partial \left(\frac{\partial z_t^*}{\partial t} \right)}{\partial \left(\frac{\gamma\varphi_t}{q_t} \right)} > 0$, iff

$$\left(\frac{\partial^2 \frac{D_t}{A_t}}{\partial t \partial \frac{\gamma\varphi_t}{q_t}}\right) < \frac{(\mathbb{E}[z \mid z > z_t^*] - z_t^*) (1 - \Omega(z_t^*))}{\omega_t(z_t^*)} \equiv \hat{\Delta}_t$$

where $\hat{\Delta}_t > 0$. In words, if an increase in the slope of the excess investment function $\frac{\gamma\varphi_t}{q_t}$ does not cause too much growth in $\frac{D_t}{A_t}$, then it an increase in $\frac{\gamma\varphi_t}{q_t}$ is associated to an increase in the growth rate of the threshold z_t^* .

B.9 Baseline vs complete markets

In this appendix we want to highlight the differences between the model presented in this paper and the standard representative agent New Keynesian model with capital (complete markets). Note first that the baseline economy collapses to the standard complete market economy if the collateral constraint is made infinitely slack (assuming that the support of entrepreneurs productivity distribution is bounded above). In that case entrepreneurial net worth becomes irrelevant and only the entrepreneur with the highest level of productivity z_t produces, since she can frictionlessly lend all the capital in the economy. Her productivity determines aggregate productivity $Z_t = (z_t^{max})^\alpha$. In contrast, in the baseline model with incomplete markets, entrepreneurs' firms can only

use capital up to a multiple γ of their net worth, i.e. $\gamma a_t \leq k_t$. Thus entrepreneurs need to accumulate net worth (in units of capital) to alleviate these financial frictions. Hence, in the baseline model, the distribution of aggregate capital across entrepreneurs and the representative household matters and aggregate productivity depends on the expected productivity of constrained firms, $Z = (\mathbb{E}[z \mid z > z_t^*])^\alpha$. The rest of the agents (retailers, final good producers, capital producers) are identical in both economies.

Below we report the equilibrium conditions in the complete markets economy. Comparing them with those of the baseline economy reveals that they are identical up to the fact that in the baseline Z_t is endogenous (and determined by a bunch of extra equations) and up to a term in the condition equating the cost of capital R_t with the marginal return on capital.

The competitive equilibrium of the complete market model with capital consists of the following 15 equations for the 15 variables $\{w, r, q, \varphi, K, L, C, D, \iota, \pi, m, \tilde{m}, i, Y, T\}$:

$$\tilde{m}_t = m_t(1 - \tau)$$

$$w_t = (1 - \alpha)m_t Z_t K_t^\alpha L_t^{-\alpha}$$

$$R_t = \alpha m_t Z_t K_t^{\alpha-1} L_t^{1-\alpha}$$

$$K_t = D_t$$

$$\dot{K}_t = (\iota_t - \delta)K_t$$

$$\frac{\dot{C}_t}{C_t} = \frac{r_t - \rho_t^h}{\eta}$$

$$w_t = \frac{\Upsilon L_t^\vartheta}{C_t^{1-\eta}}$$

$$\dot{D}_t = [(R_t - \delta q_t) D_t + w_t L_t - C_t + T_t] / q_t$$

$$r_t = i_t - \pi_t$$

$$r_t = \frac{R_t - \delta q_t + \dot{q}_t}{q_t}$$

$$(q_t - 1 - \Phi'(\iota_t)) (r_t - (\iota_t - \delta)) = \dot{q}_t - \Phi''(\iota_t) i_t - (q_t \iota_t - \iota_t - \Phi(\iota_t))$$

$$\left(r_t - \frac{\dot{Y}_t}{Y_t} \right) \pi_t = \frac{\varepsilon}{\theta} (\tilde{m}_t - m^*) + \dot{\pi}_t, \quad m^* = \frac{\varepsilon - 1}{\varepsilon}$$

$$Y_t = Z_t K_t^\alpha L_t^{1-\alpha}$$

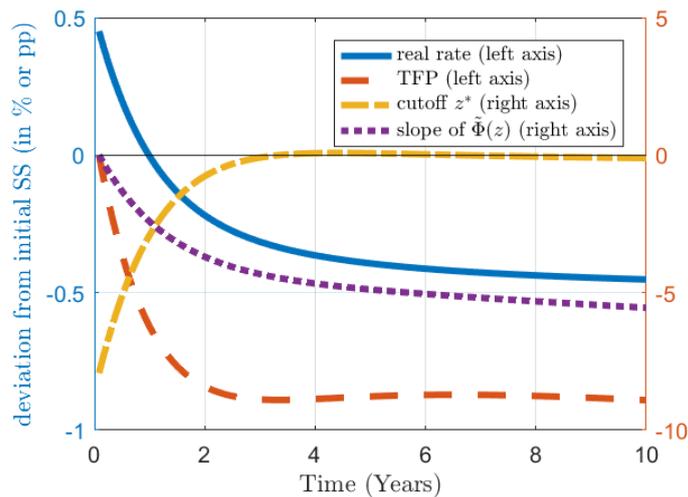
$$T_t = (1 - m_t) Y_t - \frac{\theta}{2} \pi_t^2 Y_t + \left[\iota_t q_t - \iota_t - \frac{\phi^k}{2} (\iota_t - \delta)^2 \right] K_t$$

$$di = -v \left(i_t - (\rho_t^h + \phi (\pi_t - \bar{\pi}) + \bar{\pi}) \right) dt.$$

Dynamics after permanent real interest rate declines

Figure 12 displays the impulse responses to a permanent decline in the household's discount factor ρ^h , from 1% to 0.5%. It shows that the decline in real rates (solid blue line) is accompanied by a decline in TFP (dashed orange line). This is both a consequence of the decline in the threshold (dashed-dotted yellow line) and the lower slope of the excess investment rate (dotted purple line), which increase the share of low-MRPK firms in production. The initial increase in real rates is a consequence of the nominal rigidities and the Taylor rule: as nominal rates do not decrease as fast as the natural rate on impact, it initially produces a fall in inflation that mechanically increases the real rate. As nominal rates progressively adjust, this effect disappears after one year.⁴⁵

Figure 12: Transition to a low-real-rate steady state.

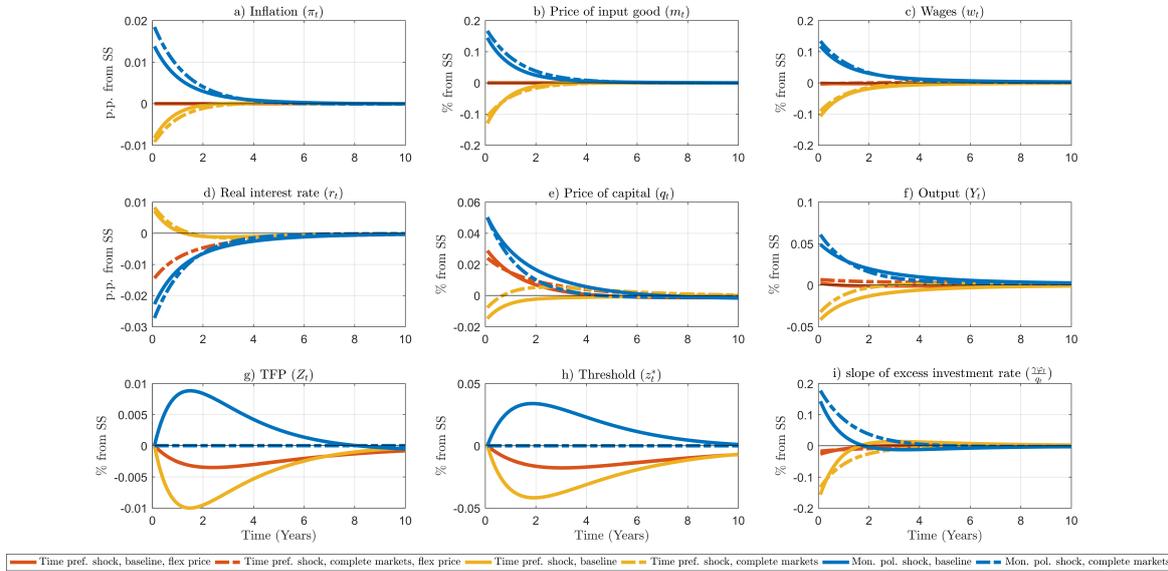


Notes: The figure shows the paths after an unexpected and permanent decline in the household's discount factor from 1% to 0.5% expressed in deviations from the initial steady state. The lines depict real rates r (solid blue), TFP Z (dashed orange), the threshold z^* (dashed-dotted yellow) and the slope of the excess investment function $\tilde{\Phi}(z)$ (dotted purple line).

⁴⁵In an economy without nominal rigidities, real rates would always be below the initial value, and TFP would also fall (not shown) through the same channels.

B.10 HANK vs RANK

Figure 13: Impulse responses.



Notes: The figure shows the dynamics of key variables in deviations from steady state of the economy. The solid orange line is the response of the heterogeneous firms economy with flexible prices to the time preference shock, as in Figure 4.2. The dashed orange line is the corresponding response in the complete markets textbook representative agent New Keynesian model. The solid blue line is the response of the baseline economy to monetary policy shock, as in Figure 4.2. The dashed blue line is the corresponding response in the complete markets economy. The paths for the time preference shock with sticky prices are shown in yellow. The orange solid line in panel d) is overlaid by the blue solid line.

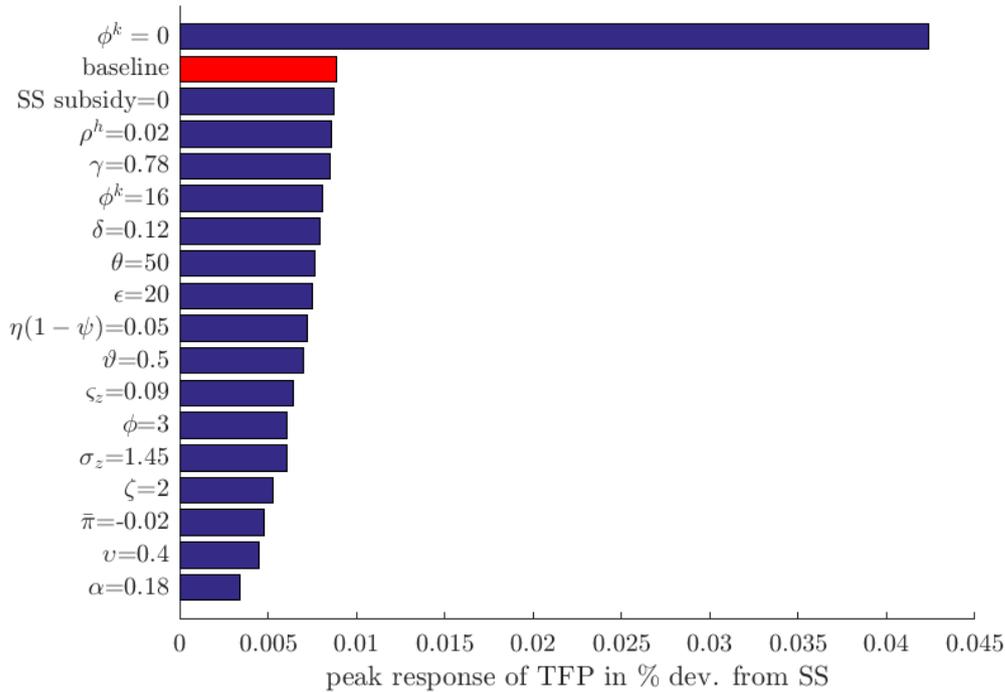
B.11 Robustness of the model with respect to parameters

Here we show how our parameter choices affect our key result, that expansionary monetary policy increases TFP. One at a time we vary all the parameters of the model. We vary most parameters by multiplying their baseline value (Table 3 in the main text) either by 2 or 1/2 (exceptions being $\bar{\pi}$ which we set from 0% to 2%, and the subsidy on input goods τ which we set to 0). The direction of the variation of the parameter is always chosen so as to reduce the impact of the monetary policy shock on TFP, relative to baseline. Deviating from the baseline in the opposite direction would increase the impact. Furthermore, we report of removing capital adjustment costs (first bar).

We report the peak of the TFP response, which is typically observed between towards the end of the second year after the shock. The response of TFP to a monetary policy expansion is positive regardless of the variation we consider. Unsurprisingly, a lower *capital* share is particularly effective at reducing the impact of monetary policy

on TFP through *capital* misallocation.

Figure 14: Robustness of the positive TFP response



Notes: Peak TFP response to a 1 b.p. expansionary monetary policy shock, conditional on parameters.

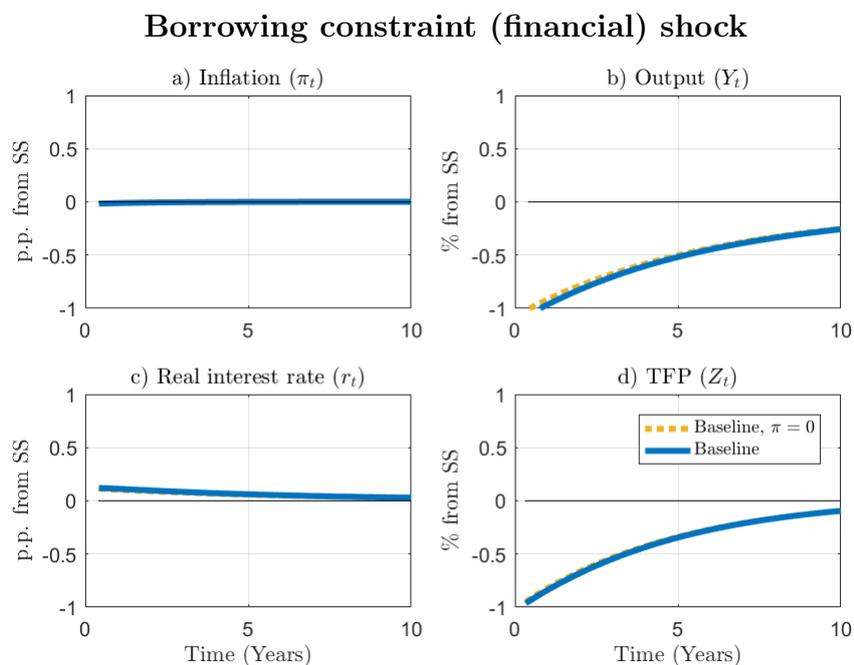
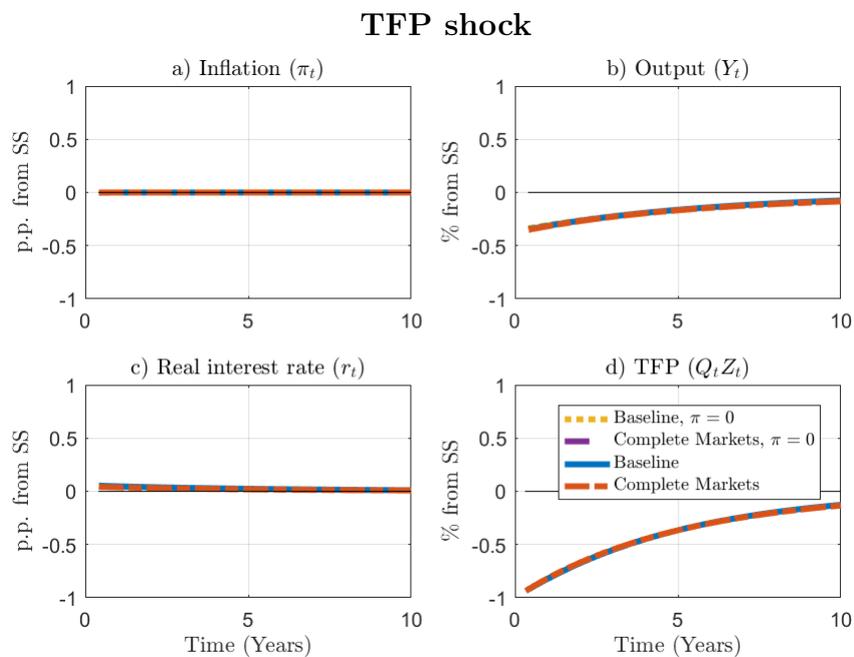
B.12 Optimal policy with other shocks

The divine coincidence also holds in the optimal response to other shocks, such as a shock to TFP or a financial shock, as we show in Appendix B.12. The TFP shock assumes that the individual production function is now $y_t = \Gamma_t f_t(z_t, k_t, l_t) = \Gamma_t (z_t k_t)^\alpha (l_t)^{1-\alpha}$, where Γ_t is an exogenous TFP shock. The aggregate TFP is then $Z_t = \Gamma_t (\mathbb{E}_{\omega_t(\cdot)} [z | z > z_t^*])^\alpha$.

The financial shock temporarily changes the borrowing limit γ_t in equation (4): $q_t k_t \leq \gamma_t q_t a_t$. This second shock is particularly interesting. As discussed by Reis (2022), changes in borrowing constraints feed into capital misallocation, and may be helpful to understand the widening gap between the average return on private capital investment, which in our model is given by $\varphi_t \mathbb{E}_{\omega_t(\cdot)} [z | z > z_t^*]$, and the real interest rate on public bonds r_t . Our results show that, in response to this shock, the central bank should steer

the real rate to follow the path of the natural rate, which may lead to a endogenous change in the gap between private and public returns. In other words, the central bank should still aim to preserve price stability, even if that leads endogenous misallocation to change the gap.

Figure 15: Timeless Ramsey policy in response to other shocks



Notes: Regardless of the shock, the divine coincidence holds in all cases (approximately): The real rate follows the natural rate (i.e. the real rate under the $\pi = 0$ policy), output follows the natural level of output, and inflation is very close to 0. Lines that are not visible (in particular those referring to the $\pi = 0$ policy) are overlaid by the corresponding optimal policy lines. In the first 4 panels, TFP drops by 1 p.p.. In the lower 4 panels, the drop in γ is set to 4.2% such that TFP drops by 1 p.p..

B.13 Computing the response of firm level capital growth

Figure 8 displays the average effect of an 1 b.p. expansionary monetary policy shock on the growth rate of the capital stock in the first year in p.p. – $100 * (\log k_{j,1} - \log k_{j,0})$ – as a function of the firms log MRPK before the shock $\log(MRPK_{j,0})$. We start the simulation at the steady state. To isolate the effect of monetary policy from the effect of idiosyncratic shocks, we subtract the evolution of this variable in the steady state. That is we plot $100 * ((\log k_{j,1;MP} - \log k_{j,0}) - (\log k_{j,1;SS} - \log k_{j,0}))$. The relationship is calculated analytically given equilibrium price paths, aggregate net-worth and the Kolmogorov forward equation. In particular we proceed as follows:

Be $A_{t,t+1}$ the operator of the law of motion of net worth share distribution that moves it forward from period t to $t + 1$. This operator depends on factor prices. Be $f_{t,s}(z_t, z_s)$ is the distribution of net worth shares across entrepreneurs that have productivity z_s in period s and had productivity z_t in period $t < s$. Hence $f_{t,t}(z_t, z_t) = \omega_t(z)$. Then:

$$f_{t,t+1}(z_t, z_{t+1}) = A_{t,t+1}f_{t,t}(z_t, z_t).$$

Define $\bar{f}_{t,s}(z_t)$ as the distribution of *net worth shares* in period s across entrepreneurs that had productivity z_t in period t .

$$\bar{f}_{t,t+1}(z_t) = \int_0^{\infty} f_{t,t+1}(z_t, z_{t+1}) dz_{t+1}$$

Define $\tilde{f}_{t,s}(z_t)$ as the distribution of *capital* employed in period s by firms that had productivity z_t in period t .

$$\tilde{f}_{t,t+1}(z_t) = \int_{z_{t+1}^*}^{\infty} \gamma A_t f_{t,t+1}(z_t, z_{t+1}) dz_{t+1}$$

Figure 8 thus plots $(\log(\tilde{f}_{0,1;MP}(z_0)) - \log(\tilde{f}_{0,0}(z_0))) - (\log(\tilde{f}_{0,1;SS}(z_0)) - \log(\tilde{f}_{0,0}(z_0))) = \log(\tilde{f}_{0,1;MP}(z_0)) - \log(\tilde{f}_{0,1;SS}(z_0))$, but converting the units on the x-axis from z to MRPK.⁴⁶ That is, the figure shows how much the capital stock of firms which had MRPK z_0 prior to the shock has changed on average 1 year after the monetary policy shock due to the shock. The slope of this function is the exact (nonlinear) counterpart

⁴⁶Note that the distribution \tilde{f} changes even in steady state, hence $\tilde{f}_{0,1;SS}(z_0) = \tilde{f}_{0,0}(z_0)$.

of the β_2 coefficient regression equation (41).

C Numerical Appendix

We discretize the model using a finite difference approach and compute non-linearly the responses to temporary change in parameters (an "MIT shock") using a Newton algorithm. Instead of time iterations over guesses for aggregate sequences, as is common in the literature, we use a global relaxation algorithm. This approach has been made popular in discrete-time models by [Juillard et al. \(1998\)](#) thanks to Dynare, but it is somewhat less common in continuous-time models (e.g. [Trimborn et al., 2008](#)). This approach helps to overcome the curse of dimensionality since in the sequence space the complexity of the problem grows only linearly in the number of aggregate variables, whereas the complexity of the state-space solution grows exponentially in the number of state variables. Recently [Auclert et al. \(2021\)](#) have exploited a particularly efficient variant of this approach in the context of heterogeneous-agent models.⁴⁷ We build on these contributions when we compute the optimal transition path. Again we make use of Dynare. We use its nonlinear Newton solver to compute both the steady state of the Ramsey problem and the optimal transition path under perfect foresight. To find the steady state, we provide Dynare with the steady state of the private equilibrium conditions as a function of the policy instrument.

C.1 Finite difference approximation of the Kolmogorov Forward equation

The KF equation is solved by a finite difference scheme following [Achdou et al. \(2021\)](#). It approximates the density $\omega_t(z)$ on a finite grid $z \in \{z_1, \dots, z_J\}$, $t \in \{t_1, \dots, t_N\}$ with steps Δz and time steps Δt . We use the notation $\omega_j^n := \omega_{n\Delta t}(z_j)$, $j = 1, \dots, J$, $n = 0, \dots, N$. The KF equation is then approximated as

$$\frac{\omega_j^n - \omega_j^{n-1}}{\Delta t} = \left(s_n(z_j) - \frac{\dot{A}_n}{A_n} - (1 - \psi)\eta \right) \omega_n(z_j) - \frac{\omega_j^n \mu(z_j) - \omega_{j-1}^n \mu(z_{j-1})}{\Delta z} + \frac{\omega_{j+1}^n \tilde{\sigma}^2(z_{j+1}) + \omega_{j-1}^n \tilde{\sigma}^2(z_{j-1}) - 2\omega_j^n \tilde{\sigma}^2(z_j)}{2(\Delta z)^2},$$

⁴⁷Compared to [Auclert et al. \(2020\)](#), who break the solution procedure into two steps, first solving for the idiosyncratic variables given the aggregate variables, we solve for the path of all aggregate and idiosyncratic variables at once. Note that, besides the nonlinear perfect foresight method we refer to here (see their Section 6), they also propose a linear method.

which, grouping, results in

$$\frac{\omega_j^n - \omega_j^{n-1}}{\Delta t} = \underbrace{\left[\left(s_n(z_j) - \frac{\dot{A}_n}{A_n} - (1 - \psi)\eta \right) - \frac{\mu(z_j)}{\Delta z} - \frac{\tilde{\sigma}^2(z_j)}{(\Delta z)^2} \right]}_{\beta_j^n} \omega_j^n + \underbrace{\left[\frac{\mu(z_{j-1})}{\Delta z} + \frac{\tilde{\sigma}^2(z_{j-1})}{2(\Delta z)^2} \right]}_{\varrho_{j-1}^n} \omega_{j-1}^n + \underbrace{\left[\frac{\tilde{\sigma}^2(z_{j+1})}{2(\Delta z)^2} \right]}_{\chi_{j+1}^n} \omega_{j+1}^n.$$

The boundary conditions are the ones associated with a reflected process z at the boundaries:⁴⁸

$$\begin{aligned} \frac{\omega_1^n - \omega_1^{n-1}}{\Delta t} &= (\beta_1^n + \chi_1^n) \omega_1^n + \chi_2^n \omega_{j+1}^n, \\ \frac{\omega_J^n - \omega_J^{n-1}}{\Delta t} &= (\beta_J^n + \varrho_J^n) \omega_J^n + \varrho_{J-1}^n \omega_{J-1}^n. \end{aligned}$$

If we define matrix

$$\mathbf{B}^n = \begin{bmatrix} \beta_1^n + \chi_1^n & \chi_2^n & 0 & 0 & \cdots & 0 & 0 & 0 & 0 \\ \varrho_1^n & \beta_2^n & \chi_3^n & 0 & \cdots & 0 & 0 & 0 & 0 \\ 0 & \varrho_2^n & \beta_3^n & \chi_4^n & \cdots & 0 & 0 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & \varrho_{J-2}^n & \beta_{J-1}^n & \chi_J^n & 0 \\ 0 & 0 & 0 & 0 & \cdots & 0 & \varrho_{J-1}^n & \beta_J^n + \varrho_J^n & 0 \end{bmatrix},$$

then we can express the KF equation as

$$\frac{\boldsymbol{\omega}^n - \boldsymbol{\omega}^{n-1}}{\Delta t} = \mathbf{B}^{n-1} \boldsymbol{\omega}^n,$$

or

$$\boldsymbol{\omega}^n = (\mathbf{I} - \Delta t \mathbf{B}^{n-1})^{-1} \boldsymbol{\omega}^{n-1}, \quad (81)$$

where $\boldsymbol{\omega}^n = \left[\omega_1^n \ \omega_2^n \ \dots \ \omega_{J-1}^n \ \omega_J^n \right]^T$, and \mathbf{I} is the identity matrix of dimension J .

⁴⁸It is easy to check that this formulation preserves the fact that matrix \mathbf{B}^n below is the transpose of the matrix associated with the infinitesimal generator of the process.

Extension to non-homogeneous grids Our model can be solved using a homogeneous grid. However, we use a non-homogeneous grid for the state z to economize on grid points. This is useful for two reasons: First, it allows us to concentrate grid points around z_t^* , which is convenient since z_t^* does not live on the grid, which introduces additional approximation error. Second, numerical error may pile up at the lower end of the grid. We could not find a universally applicable way to implement non-homogeneous grids in the economics literature, so we propose the following discretization scheme.⁴⁹

Be $z = [z_1, z_2, \dots, z_{J-1}, z_J]$ the grid. Define $\Delta z_{a,b} = z_b - z_a$ and let $\Delta z = \frac{1}{2} [\Delta z_{1,2}, \Delta z_{1,3}, \Delta z_{2,4}, \dots, \Delta z_{J-2,J}, \Delta z_{J-1,J}]$. We approximate the KFE (26) using central difference for both the first derivative and the second derivative.

$$\begin{aligned} \frac{\omega_j^n - \omega_j^{n-1}}{\Delta t} &= \left(s_n(z) - (1 - \psi)\eta - \frac{\dot{A}_n}{A_n} \right) \omega^n(z_j) - \left[\frac{\mu(z_{j+1})\omega^n(z_{j+1}) - \mu(z_{j-1})\omega^n(z_{j-1})}{\Delta z_{j-1,j+1}} \right] \\ &+ \frac{1}{2} \frac{\Delta z_{j-1,j} \sigma^2(z_{j+1}) \omega^n(z_{j+1}) + \Delta z_{j,j+1} \sigma^2(z_{j-1}) \omega^n(z_{j-1}) - \Delta z_{j-1,j+1} \sigma^2(z_j) \omega^n(z_j)}{\frac{1}{2} (\Delta z_{j-1,j+1}) \Delta z_{j,j+1} \Delta z_{j-1,j}} \end{aligned}$$

which, grouping, results in

$$\begin{aligned} \frac{\omega_j^n - \omega_j^{n-1}}{\Delta t} &= \underbrace{\left[\left(s_n(z) - (1 - \psi)\eta - \frac{\dot{A}_n}{A_n} \right) - \frac{\sigma^2(z_j)}{\Delta z_{j,j+1} \Delta z_{j-1,j}} \right]}_{\beta_j^n} \omega^n(z_j) \\ &+ \underbrace{\left[\frac{\mu(z_{j-1})}{\Delta z_{j-1,j+1}} + \frac{\sigma^2(z_{j-1})}{(\Delta z_{j-1,j+1}) \Delta z_{j,j+1}} \right]}_{\rho_{j-1}^n} \omega^n(z_{j-1}) \\ &+ \underbrace{\left[-\frac{\mu(z_{j+1})}{\Delta z_{j-1,j+1}} + \frac{\sigma^2(z_{j+1})}{(\Delta z_{j-1,j+1}) \Delta z_{j,j+1}} \right]}_{\chi_{j+1}^n} \omega^n(z_{j+1}). \end{aligned}$$

The boundary conditions are the ones associated with a reflected process z at the boundaries:

$$\frac{\omega_1^n - \omega_1^{n-1}}{\Delta t} = (\beta_1^n + \chi_1^n) \omega_n(z_1) + \chi_2^n \omega_{j+1}^n,$$

⁴⁹Our approach builds on the one in the appendix to [Achdou et al., 2021](#). It differs from theirs in two ways. First, it can be derived as a finite difference scheme over the KFE. Second, it relies on central differences for the first order derivative, and hence it is not an upwind scheme.

$$\frac{\omega_J^n - \omega_J^{n-1}}{\Delta t} = (\beta_J^n + \varrho_J^n) \omega_n(z_J) + \varrho_{J-1}^n \omega_{j-1}^n.$$

where we define $\Delta z_{0,1} \equiv \Delta z_{1,2}$ and $\Delta z_{J,J+1} \equiv \Delta z_{J-1,J}$.

The law of motion of ω can equivalently be written in matrix form

$$\frac{\omega^n - \omega^{n-1}}{\Delta t} = \mathbf{B}^{n-1} \omega^n$$

where

$$\mathbf{B}^n = \begin{bmatrix} \beta_1^n + \chi_1^n & \chi_2^n & 0 & 0 & \cdots & 0 & 0 & 0 \\ \varrho_1^n & \beta_2^n & \chi_3^n & 0 & \cdots & 0 & 0 & 0 \\ 0 & \varrho_2^n & \beta_3^n & \chi_4^n & \cdots & 0 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & \varrho_{J-2}^n & \beta_{J-1}^n & \chi_J^n \\ 0 & 0 & 0 & 0 & \cdots & 0 & \varrho_{J-1}^n & \beta_J^n + \varrho_J^n \end{bmatrix},$$

Abstracting for brevity from the term $\left(s_n(z) - (1 - \psi)\eta - \frac{\dot{A}_n}{A_n} \right)$, which is independent of the grid, and spelling out \mathbf{B}^n we have

$$\frac{\omega^n - \omega^{n-1}}{\Delta t} = \begin{bmatrix} \frac{\sigma^2(z_1)}{\Delta z_{1,2} \Delta z_{1,2}} - \frac{\mu(z_1)}{\Delta z_{1,2}} - \frac{2\sigma^2(z_1)}{\Delta z_{1,2} \Delta z_{1,2}} & -\frac{\mu(z_2)}{\Delta z_{1,2}} + \frac{\sigma^2(z_2)}{\Delta z_{1,2} \Delta z_{1,2}} & 0 & \cdots \\ \frac{\mu(z_1)}{\Delta z_{1,3}} + \frac{\sigma^2(z_1)}{\Delta z_{1,3} \Delta z_{1,2}} & -\frac{\sigma^2(z_2)}{\Delta z_{1,2} \Delta z_{2,3}} & -\frac{\mu(z_3)}{\Delta z_{1,3}} + \frac{\sigma^2(z_3)}{\Delta z_{1,3} \Delta z_{2,3}} & \cdots \\ 0 & \frac{\mu(z_2)}{\Delta z_{2,4}} + \frac{\sigma^2(z_2)}{\Delta z_{2,4} \Delta z_{2,3}} & -\frac{\sigma^2(z_3)}{\Delta z_{2,3} \Delta z_{3,4}} & \cdots \\ 0 & 0 & \frac{\mu(z_3)}{\Delta z_{3,5}} + \frac{\sigma^2(z_3)}{\Delta z_{3,4} \Delta z_{3,5}} & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix} \omega^n.$$

We can rewrite this as follows

$$\frac{\omega^n - \omega^{n-1}}{\Delta t} = \begin{bmatrix} -\frac{\mu(z_1)}{\Delta z_{1,2}} - \frac{\sigma^2(z_1)}{\Delta z_{1,2} \Delta z_{1,2}} & -\frac{\mu(z_2)}{\Delta z_{1,2}} + \frac{\Delta z_{2,3} \sigma^2(z_2)}{\Delta z_{2,3} (\Delta z_{1,2} \Delta z_{1,2})} & 0 & \cdots \\ \frac{\mu(z_1)}{\Delta z_{1,3}} + \frac{\sigma^2(z_1)}{\Delta z_{1,3} \Delta z_{1,2}} & -\frac{(\Delta z_{1,2} + \Delta z_{2,3}) \sigma^2(z_2)}{\Delta z_{1,3} (\Delta z_{1,2} \Delta z_{2,3})} & -\frac{\mu(z_3)}{\Delta z_{1,3}} + \frac{\Delta z_{3,4} \sigma^2(z_3)}{\Delta z_{3,4} (\Delta z_{1,3} \Delta z_{2,3})} & \cdots \\ 0 & \frac{\mu(z_2)}{\Delta z_{2,4}} + \frac{\Delta z_{1,2} \sigma^2(z_2)}{\Delta z_{1,2} (\Delta z_{2,4} \Delta z_{2,3})} & -\frac{(\Delta z_{2,3} + \Delta z_{3,4}) \sigma^2(z_3)}{\Delta z_{2,4} (\Delta z_{2,3} \Delta z_{3,4})} & \cdots \\ 0 & 0 & \frac{\mu(z_3)}{\Delta z_{3,5}} + \frac{\Delta z_{2,3} \sigma^2(z_3)}{\Delta z_{2,3} (\Delta z_{3,4} \Delta z_{3,5})} & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix} \omega^n.$$

Note that the bold terms in row i are equal to $1/\Delta z_i$, where Δz_i is the i -th element of Δz . Furthermore note that, up to the bold terms, the columns sum up to 0. Thus $\Delta z \mathbf{B}^n$ yields a vector of ones and the operation is mass preserving, in the sense that

the above relationship guarantees that

$$\sum_j \omega_j^n \Delta z_j = \sum_j \omega_j^{n-1} \Delta z_j = 1$$

where $\sum_j \omega_j^n \Delta z_j$ is a trapezoid approximation of the integral $\int \omega^n(z) dz$.

C.2 Finite difference approximation of the integrals

To approximate the integrals in $\int_0^z \omega_t(z) dz$ and $\int_{z_t^*}^\infty z \omega_t(z) dz$ we use the trapezoid rule. I.e. if $f(z)$ is either $\omega_t(z)$ or $z \omega_t(z)$ and $z_j \leq \bar{z} \leq z_{j+1}$ then the integral from the closest lower gridpoint is given by

$$\int_{z_j}^{\bar{z}} f(z) dz = \left[f(z_j) + \frac{1}{2} [f(z_{j+1}) - f(z_j)] \frac{\bar{z} - z_j}{z_{j+1} - z_j} \right] (\bar{z} - z_j)$$

We use this formula to construct the integrals over a larger range piecewise. For example:

$$\int_{z_1}^{z_N} f(z) dz = \begin{bmatrix} \frac{1}{2} & 1 & 1 & \dots & 1 & \frac{1}{2} \end{bmatrix} \begin{bmatrix} f(z_1) \\ f(z_2) \\ \vdots \\ f(z_N) \end{bmatrix}$$

and

$$\begin{aligned} \int_{z_1}^{z^*} f(z) dz &= \begin{bmatrix} \frac{1}{2} & 1 & 1 & \dots & 1 & \frac{1}{2} \end{bmatrix} \begin{bmatrix} f(z_1) \\ f(z_2) \\ \vdots \\ f(z_{j^*-1}) \end{bmatrix} \\ &+ \left[f(z_{j^*-1}) + \frac{1}{2} [f(z_{j^*}) - f(z_{j^*-1})] \frac{z^* - z_{j^*-1}}{z_{j^*} - z_{j^*-1}} \right] (z^* - z_{j^*-1}) \\ &\text{where } j^* = \arg \min_j \{j \leq J | z_{j^*} > z^*\} \end{aligned}$$

C.3 Algorithm to solve for the SS

Here we present how to solve for the SS of the private equilibrium, that is for the SS when the central bank sets a certain level of the nominal interest rate in SS i^{ss} .

We know that in SS consumption does not grow, hence from (14)

$$r^{ss} = \rho^h. \quad (82)$$

We also know that in SS, the investment rate is equal to the depreciation,

$$l^{ss} = \delta. \quad (83)$$

This means that, from equation (17) and the functional form we assumed for the capital adjustment costs

$$(q_t - 1 - \Phi'(l_t))(r_t - (l_t - \delta)) = \dot{q}_t - \Phi''(l_t) l_t - (q_t l_t - l_t - \Phi(l_t)) \quad (84)$$

$$(q^{ss} - 1 - \phi^k(l^{ss} - \delta))(\rho^{hh} - (l^{ss} - \delta)) = 0 - \phi^k * 0 - (q^{ss} l^{ss} - l^{ss} - \phi^k(l^{ss} - \delta))$$

$$\rho^{hh}(q^{ss} - 1) = \delta(1 - q^{ss})$$

.From here we can solve for the steady state value of q^{ss} , which is given by

$$q^{ss} = 1. \quad (85)$$

Furthermore, combining (82) with the fisher equation and the fact that the planner sets a certain nominal rate i^{ss} we get that

$$\pi^{ss} = i^{ss} - \rho^h. \quad (86)$$

In SS, $\dot{\pi}_t = 0$ and $\dot{Y}_t = 0$. Hence, from equation (21) we obtain

$$m^{ss} = \left(m^* + \rho^h \pi^{ss} \frac{\theta}{\varepsilon} \right). \quad (87)$$

Using equation (34) and (82),

$$\rho^h = \frac{1}{q^{ss}} \left(\alpha m_t Z_t A_t^{\alpha-1} L^{1-\alpha} \frac{z_t^*}{\gamma X_t} \right) - \delta \quad (88)$$

From equation (35) and (82),

$$\frac{\dot{A}_t}{A_t} = 0 = \frac{1}{q_t}(\alpha m_t Z_t A_t^{\alpha-1} L_t^{1-\alpha} - R_t(1 - \Omega(z_t^*)) + R_t - \delta q_t - q_t(1 - \psi)\eta). \quad (89)$$

Plugging the latter equation into the former, using $q^{SS} = 1$ and using the definition of r_t we obtain:

$$\rho^h + \delta = [(\rho^h + \delta)(\gamma(1 - \Omega(z_t^*)) - 1) + (1 - \psi)\eta + \delta] \frac{z^*}{\gamma X^*}. \quad (90)$$

In the algorithm, we use a non-linear equation solver to obtain z^* from this equation. *The Algorithm.*

- Get $r^{ss} = \rho^h$, $\pi^{ss} = \bar{\pi}$ and $i^{ss} = \rho^h + \pi^{ss}$ and $R^{ss} = q^{ss}(\rho^h + \delta)$ and $m^{ss} = m^* + \rho^h \pi^{ss} \frac{\theta}{\epsilon}$.
- Given that our calibration target for $L^{ss} = 1$, we “guess” $L^{ss} = 1$
- Let n now denote the iteration counter. Make an initial guess for the net worth distribution ω^0
 1. Use a non-linear equation solver on equation (90) to obtain z^* from equation (90).
 2. Obtain $Z_n = (\gamma_n X_n^*)^\alpha$.
 3. Find A from equation (33),

$$A^n = \left[\frac{q^{ss} \rho^h + \delta q^{ss}}{\alpha m_n Z_n L_n^{1-\alpha} \frac{z_t^*}{\gamma X_t}} \right]^{\frac{1}{\alpha-1}}.$$

4. Find the stocks $K_n = \gamma(1 - \Omega^n(z^*))A^n$, $D_n = K_n - A_n$.
5. Compute $w_n = (1 - \alpha)m^{ss}Z_n A_n^\alpha L_n^{-\alpha}$, $\varphi_n = \alpha \left(\frac{1-\alpha}{w_n} \right)^{(1-\alpha)/\alpha} m^{ss \frac{1}{\alpha}}$.
6. Get aggregate output $Y = Z_n A_n^\alpha L_n^{1-\alpha}$, transfers $T_n = (1 - m^{ss})Y_n - \frac{\theta}{2}(\pi^{ss})^2 Y_n + (1 - \psi)\eta A_t$, and consumption $C_n = w_n L_m + r^{ss} D_n + T_n$.
7. Update $\hat{s}_j^n = \frac{1}{q^{ss}}(\gamma \max\{z\varphi_n - R_n, 0\} + R_n - \delta q^{ss})$ and employ it to construct matrix \mathbf{B}^{n-1} .

8. Update ω^{n+1} using equation $\frac{\omega^{n+1}-\omega^n}{\Delta t} = \mathbf{B}^n \omega^{n+1}$.
 9. If the net worth distribution do not coincide with the guess, set $n = n + 1$ and return to point 1
- Set $\Upsilon = (w_{L=1} C_{L=1}^{-\eta})$ to ensure our “guess” for L^{ss} is correct.

D Computing optimal policies in heterogeneous-agent models

D.1 General algorithm

Solving for the optimal policy in models with heterogeneous agents poses a certain challenge since the state in such a model contains a distribution, which is an infinite-dimensional object. In this Section, we explain how such models can be solved in a relatively straightforward manner. Our approach relies on three main conceptual ingredients: (i) finite difference approximation of continuous time and continuous idiosyncratic states, (ii) symbolic derivation of the planner's first-order conditions, and (iii) use of a Newton algorithm to solve the optimal policy problem non-linearly in the sequence space. Here we present a general overview which goes beyond the particular model presented in the paper.

(i) Finite difference approximation A continuous-time, continuous-space heterogeneous-agent model discretized using an upwind finite-difference method becomes a discrete-time, discrete-space model. In this discretized model the dynamics of the (now finite-dimensional) distribution $\boldsymbol{\mu}_t$ at period t are given by

$$\boldsymbol{\mu}_t = (\mathbf{I} - \Delta t \mathbf{A}_t^T)^{-1} \boldsymbol{\mu}_{t-1}, \quad (91)$$

where Δt is the time step between periods and \mathbf{A}_t is a matrix whose entries depend *nonlinearly* and in *closed form* on the idiosyncratic and aggregate variables in period t .⁵⁰ Similarly, the HJB equation is approximated as⁵¹

$$\rho \mathbf{v}_{t+1} = \mathbf{u}_{t+1} + \mathbf{A}_{t+1} \mathbf{v}_{t+1} - (\mathbf{v}_{t+1} - \mathbf{v}_t) / \Delta t. \quad (92)$$

Together with additional static equations, such as market clearing conditions or budget constraints, and aggregate dynamic equations, including the Euler equations of representative agents (if any) and the dynamics of aggregate states, they define the discretized model.

⁵⁰Technically, this matrix results from the discretization of the *infinitesimal generator* of the idiosyncratic states. In the example of Section 2, $\boldsymbol{\mu}_t = \boldsymbol{\omega}_t$ and $\mathbf{A}_t = \mathbf{B}_t$.

⁵¹In the model presented in this paper the HJB can be solved analytically and hence there is no need to solve it computationally.

Though we have ended up with a discrete-time approximation, casting the original model in continuous time is central to our method. The discretized dynamics of the distribution (91) and Bellman equation (92) present two advantages compared to their counterparts in the discrete-time continuous-state formulation typically employed in the literature. First, the analytical tractability of the original continuous-time model implies that the agents' optimal choices in the discretized version are always "on the grid", avoiding the need for interpolation, and are "one step at a time" making the matrix Π_t sparse.⁵² Second, the private agent's FOCs hold with equality even at the exogenous boundaries (see Achdou et al. (2021) for a detailed discussion of these advantages).

(ii) Symbolic derivation of planner's FOCs Once we have a finite-dimensional discrete-time discrete-space model, we can derive the planner's FOCs by *symbolic differentiation* using standard software packages. For convenience, we rely on Dynare's toolbox for Ramsey optimal policy to do this task for us. To this end, we simply provide the discretized version of our model's private equilibrium conditions to Dynare (the discretized counterpart to the equations in Appendix B.7), making use of loops for the heterogeneous-agent block, as in Winberry (2018). We furthermore provide the discretized objective function, and Dynare then takes symbolic derivatives to construct the set of optimality conditions of the planner for us.

A natural question at this stage is under which conditions the optimal policies of the discrete-time, discrete-space problem coincide with those of the original problem. The following proposition shows that, if the time interval is small enough (the standard condition when approximating continuous-time models), then the two solutions coincide.

Proposition E.1 : *Provided that all the Lagrange multipliers associated to the equilibrium conditions are continuous for $t > 0$, the solution of the "discretize-optimize" and the "optimize-discretize" algorithms converge to each other as the time step Δt goes towards 0.*

Proof: See Appendix D.3.

The proposition guarantees that both strategies coincide when Δt goes towards zero. This proposition is quite general, as most continuous-time, perfect-foresight, general equilibrium models do not feature discontinuities for $t > 0$.

⁵²The introduction of Poisson shocks would not change the sparsity of matrix Π_t .

The model presented in Section 2 is arguably simpler than the general heterogeneous-agent model covered by Proposition 1, as it features an analytic solution for the HJB equation. To get an idea of the performance of our method in a case in which the HJB is also a constraint in the planner's problem, as well as to showcase its generality in dealing with different problems, we compute the optimal monetary policy in the HANK model of Nuño and Thomas (2022) using our method in Dynare (see Appendix D.4). We compare our results with those using their "optimize-discretize" algorithm at monthly frequency $\Delta t = 1/12$. We conclude that both approaches essentially coincide.

(iii) Newton algorithm to solve the optimal policy problem non-linearly in the sequence space Finally, we use the discretized optimality conditions of the planner to compute non-linearly the *optimal responses* to a temporary change in parameters (an "MIT shock") using a Newton algorithm. Instead of time iterations over guesses for aggregate sequences, as is common in the literature, we use a global relaxation algorithm. This approach has been made popular in discrete-time models by Juillard et al. (1998) thanks to Dynare, but it is somewhat less common in continuous-time models (e.g. Trimborn et al., 2008). This approach helps to overcome the curse of dimensionality since in the sequence space the complexity of the problem grows only linearly in the number of aggregate variables, whereas the complexity of the state-space solution grows exponentially in the number of state variables. Recently Auclert et al. (2021) have exploited a particularly efficient variant of this approach in the context of heterogeneous-agent models.⁵³ We build on these contributions when we compute the optimal transition path. Again we make use of Dynare. We use its nonlinear Newton solver to compute both the steady state of the Ramsey problem and the optimal transition path under perfect foresight.⁵⁴ Our hope is that the convenience of using Dynare will make optimal policy problems in heterogeneous-agent models easily accessible to a large audience of researchers.

The solution to the perfect foresight problem can be easily adapted to the case with aggregate shocks. As Boppart et al. (2018) show, the perfect-foresight transitional

⁵³Compared to Auclert et al. (2020), who break the solution procedure into two steps, first solving for the idiosyncratic variables given the aggregate variables, we solve for the path of all aggregate and idiosyncratic variables at once. Note that, besides the nonlinear perfect foresight method we refer to here (see their Section 6), they also propose a linear method.

⁵⁴To find the steady state, we provide Dynare with the steady state of the private equilibrium conditions as a function of the policy instrument.

dynamics to an "MIT shock" coincides with the solution of the model with aggregate uncertainty using a first-order perturbation approach. We follow this approach to analyze the optimal response to a cost-push shock below.

As discussed in Section 3, our solution approach is different from the one in [Winberry \(2018\)](#) or [Ahn et al. \(2018\)](#). These papers expand the seminal contribution by [Reiter \(2009\)](#), based on a two-stage algorithm that (i) first finds the nonlinear solution of the steady state of the model and (ii) then applies perturbation techniques to produce a linear system of equations describing the dynamics around the steady state. These methods, however, were not created to deal with the problem of finding the optimal policies, the focus of our algorithm, as the first stage requires the computation of the steady state, which in our case is the steady state of the problem under optimal policies. Our algorithm finds the steady state of the planner's problem, including the Lagrange multipliers. Naturally, this steady does not need to coincide with the steady state that can be found by looking for the value of the planner's policy that maximizes steady-state welfare.

D.2 Comparison to other solution methods

We are aware of 4 methods to solve Ramsey problems in general heterogeneous agents models. We compare them in this table. The purpose is not to argue that one approach dominates the others, but to show similarities and differences. An advantage of our method is that it allows for an arbitrary path of the distribution while allowing for occasionally binding constraints and being highly automatized. Grouping 4 papers into as one conceptual method glosses over some differences.

The next appendix shows that methods (1) and (4) lead to the same result in the limit as the time step converges to 0, and the subsequent appendix shows that this is indeed the case even for a normal calibration, i.e. outside of this limit.

	(1)	(2)	(3)	(4)
	Niño and Thomas (2022), Bigio and Sannikov (2021), Smirnov (2022) and Dávila and Schaab (2022)	Le Grand et al. (2022)	Bhandari et al. (2021)	This paper
time	continuous	discrete	discrete	continuous
how to cope with infinite dimensional distribution in planners problem	infinite-dimensional calculus	discrete number of truncated histories N	discretize distribution by considering large number N of individuals	discretize HJB, KFE and other equations containing the distribution with N finite elements
planner optimizes	continuous state problem	discrete state problem	discrete state problem	discrete state problem
how to find planner's FOCs	by hand	by hand	custom built code	automated in Dynare
sequence or state space method	sequence space	sequence space	state space	sequence space
representation of distribution as	finite difference approximation of distribution function with N elements	population of N histories	population of N agents	finite difference approximation of distribution function with N elements
how to tame curse of dimensionality	perfect foresight	perfect foresight	local approximation	perfect foresight
solving dynamic equilibrium conditions under optimal policy	custom build iterative procedure where the heterogeneous agents block is solved conditional on prices and the prices are solved for conditional on the heterogeneous agents block until convergence	Newton method, automated in Dynare	custom built code	Newton method, automated in Dynare
occasionally binding constraints and non-linearities	possible	possible	not possible	possible

D.3 Proof of proposition D.1

Proof: The proof has the following structure. First, we set up a generic planner's problem in a continuous-time heterogeneous-agent economy without aggregate uncertainty. Second, we derive the continuous time optimality conditions of the planner's problem and discretize them. Third, we discretize the planners problem and the derive the optimality conditions. Fourth, we compare the two sets of discretized optimality conditions.

1. The generic problem The planner's problem in an economy with heterogeneity among one agent type (e.g. households or firms) can be written as

$$\max_{Z_t, u_t(x), \mu_t(x), v_t(x)} \int_0^{\infty} \exp(-\rho t) f_0(Z_t) dt \quad (93)$$

s.t. $\forall t$

$$\dot{X}_t = f_1(Z_t) \quad (94)$$

$$\dot{U}_t = f_2(Z_t) \quad (95)$$

$$0 = f_3(Z_t) \quad (96)$$

$$\tilde{U}_t = \int f_4(x, u_t(x), Z_t) \mu_t(x) dx \quad (97)$$

$$\begin{aligned} \rho v_t(x) = & \dot{v}_t(x) + f_5(x, u_t(x), Z_t) \\ & + \sum_{i=1}^I b_i(x, u_t(x), Z_t) \frac{\partial v_t(x)}{\partial x_i} + \sum_{i=1}^I \sum_{k=1}^I \frac{(\sigma(x)\sigma(x)^\top)_{i,k}}{2} \frac{\partial^2 v_t(x)}{\partial x_i \partial x_k}, \forall x \end{aligned} \quad (98)$$

$$0 = \frac{\partial f_5}{\partial u_{j,t}} + \sum_{i=1}^I \frac{\partial b_i}{\partial u_{j,t}} \frac{\partial v_t(x)}{\partial x_i}, \quad j = 1, \dots, J, \forall x. \quad (99)$$

$$\begin{aligned} \dot{\mu}_t(x) = & - \sum_{i=1}^I \frac{\partial}{\partial x_i} [b_i(x, u_t(x), Z_t) \mu_t(x)] \\ & + \frac{1}{2} \sum_{i=1}^I \sum_{k=1}^I \frac{\partial^2}{\partial x_i \partial x_k} [(\sigma(x)\sigma(x)^\top)_{i,k} \mu_t(x)], \forall x \end{aligned} \quad (100)$$

$$X_0 = \bar{X}_0 \quad (101)$$

$$\mu_0(x) = \bar{\mu}_0(x) \quad (102)$$

$$\lim_{t \rightarrow \infty} U = \bar{U}_\infty \quad (103)$$

$$\lim_{t \rightarrow \infty} v(x) = \bar{v}(x)_\infty \quad (104)$$

where we have adopted the following notation:

- Variables (capitals are reserved for aggregate variables):
 - x individual state vector with I elements
 - u individual control vector with J elements
 - v individual value function vector with 1 element
 - $u(x)$ control vector as function of individual state
 - $\mu(x)$ distribution of agents across states
 - $v(x)$ value function as function of individual state
 - X aggregate state vector (other than μ)
 - \hat{U} aggregate control vector of purely contemporaneous variables
 - U aggregate control vector of intertemporal variables
 - \tilde{U} control vector of aggregator variables
 - $Z_t = \{\tilde{U}_t, U_t, \bar{U}_t, X_t\}$ vector of all aggregate variables
- Functions
 - b function that determines the drift of x
 - f_0 welfare function
 - f_1, f_2, f_3 aggregate equilibrium conditions
 - f_4 aggregator function
 - f_5 individual utility function

Line (93) is the planner's objective function.⁵⁵ Equations (94)-(96) are the aggregate equilibrium conditions for aggregate states, jump variables and contemporaneous variables. In our model, examples for each of these three types of equations are the law of motion of aggregate capital, the household's Euler equation and the household's labor supply condition, respectively. Equation (97) links aggregate and individual variables, such as the definition of aggregate TFP in our model. Equations (98) and (99) are the individual agent's value function and first order conditions, which must hold across the

⁵⁵Notice that the planner's discount factor, ϱ , can be different to that of individual agents, ρ .

whole individual state vector x . In our model we do not have these two types of equations since we can analytically solve the individual optimal choice. The Kolmogorov Forward equation (24) determines the evolution of the distribution of agents. Finally (101)-(104) are the initial and terminal conditions for the aggregate and individual state and dynamic control variables. In our model these are the initial capital stock and firm distribution and the terminal conditions for variables such as consumption.

2. Optimize, then discretize First we consider the approach introduced in Nuño and Thomas (2022), namely to compute the first order conditions using calculus of variations and then to discretize the problem using an upwind finite difference scheme.

2.a The Lagrangian The Lagrangian for this problem is given by:⁵⁶

$$\begin{aligned}
\mathcal{L} = & \int_0^\infty \left\{ e^{-\rho t} f_0(Z_t) \right. \\
& + \lambda_{1,t} \left(\dot{X}_t - f_1(Z_t) \right) \\
& + \lambda_{2,t} \left(\dot{U}_t - f_2(Z_t) \right) \\
& + \lambda_{3,t} (f_3(Z_t)) \\
& + \lambda_{4,t} \left(\tilde{U}_t - \int f_4(x, u_t(x), Z_t) \mu_t(x) dx \right) \\
& + \int \left[\lambda_{5,t}(x) \left(-\rho v_t(x) + \dot{v}_t(x) + f_5(x, u_t(x), Z_t) + \sum_{i=1}^I b_i(x, u_t(x), Z_t) \frac{\partial v_t(x)}{\partial x_i} + \sum_{i=1}^I \frac{\sigma_i^2(x)}{2} \frac{\partial^2 v_t(x)}{\partial^2 x_i} \right) \right] dx \\
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial f_5}{\partial u_{j,t}} + \sum_{i=1}^I \frac{\partial b_i}{\partial u_{j,t}} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \left. \int \left[\lambda_{7,t}(x) \left(-\dot{\mu}_t(x) + \left(-\sum_{i=1}^I \frac{\partial}{\partial x_i} [b_i(x, u_t(x), Z_t) \mu_t(x)] + \frac{1}{2} \sum_{i=1}^I \frac{\partial^2}{\partial^2 x_i} [\sigma_i^2(x) \mu_t(x)] \right) \right) \right] dx \right\} dt
\end{aligned}$$

where λ_1 to λ_7 denote the multipliers on the respective constraints. For convenience, we write the time derivatives in a separate line at the end. The Lagrangian becomes:

$$\begin{aligned}
\mathcal{L} = & \int_0^\infty \left\{ e^{-\rho t} f_0(Z_t) \right. \\
& + \lambda_{1,t} (-f_1(Z_t)) \\
& + \lambda_{2,t} (-f_2(Z_t))
\end{aligned}$$

⁵⁶For simplicity, we assume that the Wiener processes driving the dynamics of the state x are independent, though the proof can be trivially extended to that case, at the cost of a more cumbersome notation.

$$\begin{aligned}
& + \lambda_{3,t}(-f_3(Z_t)) \\
& + \lambda_{4,t} \left(\tilde{U}_t - \int f_4(x, u_t(x), Z_t) \mu_t(x) dx \right) \\
& + \int \left[\lambda_{5,t}(x) \left(-\rho v_t(x) + f_5(x, u_t(x), Z_t) + \sum_{i=1}^I b_i(x, u_t(x), Z_t) \frac{\partial v_t(x)}{\partial x_i} + \sum_{i=1}^I \frac{\sigma_i^2(x)}{2} \frac{\partial^2 v_t(x)}{\partial^2 x_i} \right) \right] dx \\
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial f_{5,t}}{\partial u_{j,t}} + \sum_{i=1}^I \frac{\partial b_i}{\partial u_{j,t}} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \int \left[\lambda_{7,t}(x) \left(-\sum_{i=1}^I \frac{\partial}{\partial x_i} [b_i(x, u_t(x), Z_t) \mu_t(x)] + \frac{1}{2} \sum_{i=1}^I \frac{\partial^2}{\partial^2 x_i} [\sigma_i^2(x) \mu_t(x)] \right) \right] dx \Big\} dt \\
& + \int_0^\infty \left\{ e^{-\rho t} \lambda_{1,t} \dot{X}_t + \lambda_{2,t} \dot{U}_t + \int [\lambda_{5,t} \dot{v}_t(x)] dx - \int [\lambda_{7,t} \dot{\mu}_t(x)] dx \right\} dt.
\end{aligned}$$

We have ignored the terminal and initial conditions but we will account for them later on. Now we manipulate the Lagrangian using integration by parts in order to bring it into a more convenient form. We start with the last line. Switching the order of integration, the last line becomes

$$\begin{aligned}
& \int_0^\infty e^{-\rho t} \lambda_{1,t} \dot{X}_t dt + \int_0^\infty e^{-\rho t} \lambda_{2,t} \dot{U}_t dt + \int \int_0^\infty [e^{-\rho t} \lambda_{5,t}(x) \dot{v}_t(x)] dt dx \\
& \quad - \int \int_0^\infty [e^{-\rho t} \lambda_{7,t}(x) \dot{\mu}_t(x)] dt dx
\end{aligned}$$

Now we integrate this expression by parts with respect to time t , using

$$\begin{aligned}
\int_0^\infty e^{-\rho t} a_t \dot{b}_t dt &= [e^{-\rho t} a_t b_t]_0^\infty - \int_0^\infty e^{-\rho t} (\dot{a}_t - \rho a_t) b_t dt \\
&= \lim_{t \rightarrow \infty} e^{-\rho t} a_t b_t - a_0 b_0 - \int_0^\infty e^{-\rho t} (\dot{a}_t - \rho a_t) b_t dt
\end{aligned}$$

to get

$$\begin{aligned}
& \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{1,t} X_t - \lambda_{1,0} X_0 - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{1,t} - \rho \lambda_{1,t}) X_t dt + \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{2,t} U_t - \lambda_{2,0} U_0 \\
& \quad - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{2,t} - \rho \lambda_{2,t}) U_t dt \\
& + \int \left(\lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{5,t}(x) v_t(x) - \lambda_{5,0}(x) v_0(x) \right) dx - \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{5,t}(x) - \rho \lambda_{5,t}(x)) v_t(x) dt dx \\
& - \int \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{7,t}(x) \mu_t(x) - \lambda_{7,0}(x) \mu_0(x) dx + \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{7,t}(x) - \rho \lambda_{7,t}(x)) \mu_t(x) dt dx
\end{aligned}$$

Now we use the initial and terminal conditions to drop some $\lim_{t \rightarrow \infty}$ and $t = 0$ terms,

$$\begin{aligned}
& + \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{1,t} X_t - \lambda_{2,0} U_0 - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{1,t} - \rho \lambda_{1,t}) X_t dt - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{2,t} - \rho \lambda_{2,t}) U_t dt \\
& - \int \lambda_{5,0}(x) v_0(x) dx + \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{5,t}(x) - \rho \lambda_{5,t}(x)) v_t(x) dt dx \\
& - \int \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{7,t}(x) \mu_t(x) dx + \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{7,t}(x) - \rho \lambda_{7,t}(x)) \mu_t(x) dt dx
\end{aligned}$$

Next we integrate lines 6 to 8 by parts with respect to x . This yields:

$$\begin{aligned}
& + \int \left\{ \left[\left(-\rho \lambda_{5,t}(x) v_t(x) + f_5(x, u_t(x), Z_t) - \sum_{i=1}^I \frac{\partial b_i(x, u_t(x), Z_t) \lambda_{5,t}(x)}{\partial x_i} v_t(x) \right) \right] dx \right. \\
& \quad + \int \left[\left(+\frac{1}{2} \sum_{i=1}^I \frac{\partial^2}{\partial^2 x_i} [\sigma_i^2(x) \lambda_{5,t}(x)] v_t(x) \right) \right] dx \\
& \quad + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \frac{\partial f_{5,t}}{\partial u_{j,t}} - \sum_{i=1}^I \frac{\partial [\lambda_{6,j,t}(x) \frac{\partial b_i}{\partial u_{j,t}}]}{\partial x_i} v_t(x) \right] dx \\
& \quad \left. + \int \left[\left(\sum_{i=1}^I \frac{\partial \lambda_{7,t}(x)}{\partial x_i} [b_i(x, u_t(x), Z_t) \mu_t(x)] + \sum_{i=1}^I \frac{\partial^2 \lambda_{7,t}(x) \sigma_i^2(x)}{\partial^2 x_i} \frac{\mu_t(x)}{2} \right) \right] dx \right\} dt
\end{aligned}$$

Putting this all together the Lagrangian has become:

$$\begin{aligned}
\mathcal{L} & = \int_0^\infty \{ e^{-\rho t} f_0(Z_t) \\
& + \lambda_{1,t} (-f_1(Z_t)) \\
& + \lambda_{2,t} (-f_2(Z_t)) \\
& + \lambda_{3,t} (-f_3(Z_t)) \\
& + \lambda_{4,t} \left(\tilde{U}_t - \int f_4(x, u_t(x), Z_t) \mu_t(x) dx \right) \\
& + \int \left(-\rho \lambda_{5,t}(x) v_t(x) + \lambda_{5,t}(x) f_5(x, u_t(x), Z_t) - \sum_{i=1}^I \frac{\partial [b_i(x, u_t(x), Z_t) \lambda_{5,t}(x)]}{\partial x_i} v_t(x) \right) dx \\
& + \int \left(\frac{1}{2} \sum_{i=1}^I \frac{\partial^2}{\partial^2 x_i} [\sigma_i^2(x) \lambda_{5,t}(x)] v_t(x) \right) dx
\end{aligned}$$

$$\begin{aligned}
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \frac{\partial f_{5,t}}{\partial u_{j,t}} - \sum_{i=1}^I \frac{\partial \left[\lambda_{6,j,t}(x) \frac{\partial b_i}{\partial u_{j,t}} \right]}{\partial x_i} v_t(x) \right] dx \\
& + \int_0^\infty \left[\left(\sum_{i=1}^I \frac{\partial \lambda_{7,t}(x)}{\partial x_i} [b_i(x, u_t(x), Z_t) \mu_t(x)] + \sum_{i=1}^I \frac{\partial^2 \lambda_{7,t}(x) \sigma_i^2(x)}{\partial^2 x_i} \frac{\mu_t(x)}{2} \right) \right] dx \Big\} dt \\
& + \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{1,t} X_t - \lambda_{2,0} U_0 - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{1,t} - \rho \lambda_{1,t}) X_t dt - \int_0^\infty e^{-\rho t} (\dot{\lambda}_{2,t} - \rho \lambda_{2,t}) U_t dt \\
& + \int -\lambda_{5,0}(x) v_0(x) dx + \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{5,t}(x) - \rho \lambda_{5,t}(x)) v_t(x) dt dx \\
& - \int \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{7,t}(x) \mu_t(x) dx + \int \int_0^\infty e^{-\rho t} (\dot{\lambda}_{7,t}(x) - \rho \lambda_{7,t}(x)) \mu_t(x) dt dx.
\end{aligned}$$

2.b Optimality conditions in the continuous state space We take the Gateaux derivatives in direction $h_t(x)$ for each endogenous variable x . These derivatives have to be equal to zero for any $h_t(x)$ in the optimum. This implies the following optimality conditions:

Aggregate variables:

$$U_t : 0 = -(\dot{\lambda}_{2,t} - \rho \lambda_{2,t}) \quad (105)$$

$$+ \frac{\partial f_{0,t}}{\partial U_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial U_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial U_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial U_t} - \lambda_{4,t} \int \frac{\partial f_{4,t}}{\partial U_t} \mu_t(x) dx \quad (106)$$

$$+ \int \left[\lambda_{5,t}(x) \left(\frac{\partial f_{5,t}}{\partial U_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial U_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \quad (107)$$

$$+ \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial^2 f_{5,t}}{\partial u_{j,t} \partial U_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial u_{j,t} \partial U_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \quad (108)$$

$$+ \int \left[\lambda_{7,t}(x) \left(- \sum_{i=1}^I \frac{\partial}{\partial x_i} \left[\frac{\partial b_{i,t}}{\partial U_t} \mu_t(x) \right] \right) \right] dx, \quad (109)$$

$$\forall t > 0, \quad (110)$$

$$0 = \lambda_{2,0}. \quad (111)$$

$$X_t : 0 = -(\dot{\lambda}_{1,t} - \rho \lambda_{1,t})$$

$$\begin{aligned}
& + \frac{\partial f_{0,t}}{\partial X_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial X_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial X_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial X_t} - \lambda_{4,t} \int \frac{\partial f_{4,t}}{\partial X_t} \mu_t(x) dx \\
& + \int \left[\lambda_{5,t}(x) \left(\frac{\partial f_{5,t}}{\partial X_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial X_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial^2 f_{5,t}}{\partial u_{j,t} \partial X_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial u_{j,t} \partial X_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \int \left[\lambda_{7,t}(x) \left(- \sum_{i=1}^I \frac{\partial}{\partial x_i} \left[\frac{\partial b_{i,t}}{\partial X_t} \mu_t(x) \right] \right) \right] dx, \\
& \forall t \geq 0,
\end{aligned}$$

$$0 = \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{1,t}(x).$$

$$\begin{aligned}
\hat{U}_t : 0 = & 0 \\
& + \frac{\partial f_{0,t}}{\partial \hat{U}_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial \hat{U}_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial \hat{U}_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial \hat{U}_t} - \lambda_{4,t} \int \frac{\partial f_{4,t}}{\partial \hat{U}_t} \mu_t(x) dx \\
& + \int \left[\lambda_{5,t}(x) \left(\frac{\partial f_{5,t}}{\partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial \hat{U}_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial^2 f_{5,t}}{\partial u_{j,t} \partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial u_{j,t} \partial \hat{U}_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \int \left[\lambda_{7,t}(x) \left(- \sum_{i=1}^I \frac{\partial}{\partial x_i} \left[\frac{\partial b_{i,t}}{\partial \hat{U}_t} \mu_t(x) \right] \right) \right] dx, \\
& \forall t \geq 0.
\end{aligned}$$

$$\begin{aligned}
\tilde{U}_t : 0 = & \lambda_{4,t} \\
& + \frac{\partial f_{0,t}}{\partial \tilde{U}_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial \tilde{U}_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial \tilde{U}_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial \tilde{U}_t} - \lambda_{4,t} \int \frac{\partial f_{4,t}}{\partial \tilde{U}_t} \mu_t(x) dx \\
& + \int \left[\lambda_{5,t}(x) \left(\frac{\partial f_{5,t}}{\partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial \tilde{U}_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx \\
& + \sum_{j=1}^J \int \left[\lambda_{6,j,t}(x) \left(\frac{\partial^2 f_{5,t}}{\partial u_{j,t} \partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}}{\partial u_{j,t} \partial \tilde{U}_t} \frac{\partial v_t(x)}{\partial x_i} \right) \right] dx
\end{aligned}$$

$$+ \int \left[\lambda_{7,t}(x) \left(- \sum_{i=1}^I \frac{\partial}{\partial x_i} \left[\frac{\partial b_{i,t}}{\partial \tilde{U}_t} \mu_t(x) \right] \right) \right] dx,$$

$$\forall t \geq 0.$$

Value function, distribution and policy functions

$$v_t(x) : 0 = \left(-\lambda_{5,t}(x)\rho - \sum_{i=1}^I \frac{\partial [\lambda_{5,t}(x)b_i(x, u_t(x), Z_t)]}{\partial x_i} + \frac{1}{2} \sum_{i=1}^I \frac{\partial^2}{\partial^2 x_i} [\sigma_i^2(x)\lambda_{5,t}(x)] \right)$$

$$- \sum_{j=1}^J \sum_{i=1}^I \frac{\partial}{\partial x_i} \left(\lambda_{6,j,t}(x) \frac{\partial b_i(x, u_t(x), Z_t)}{\partial u_{j,t}} \right)$$

$$- (\dot{\lambda}_{5,t}(x) - \rho \lambda_{5,t}(x)),$$

$$\forall t > 0,$$

$$0 = \lambda_{5,0}(x).$$

$$\mu_t(x) : 0 = -\lambda_{4,t} f_4(x, u_t(x), Z_t)$$

$$+ \lambda_{7,t}(x) \left(\sum_{i=1}^I \frac{\partial \lambda_{7,t}(x)}{\partial x_i} b_i(x, u_t(x), Z_t) + \sum_{i=1}^I \frac{\partial^2 \lambda_{7,t}(x)}{\partial^2 x_i} \frac{\sigma_i^2(x)}{2} \right)$$

$$+ (\dot{\lambda}_{7,t}(x) - \rho \lambda_{7,t}(x)),$$

$$\forall t \geq 0,$$

$$0 = \lim_{t \rightarrow \infty} e^{-\rho t} \lambda_{7,t}(x).$$

$$u_{l,t}(x) : 0 = -\lambda_{4,t} \frac{\partial f_4}{\partial u_{l,t}} \mu_t(x)$$

$$+ \overbrace{\left[\lambda_{5,t}(x) \left(\frac{\partial f_5}{\partial u_{l,t}} + \sum_{i=1}^I \frac{\partial b_i}{\partial u_{l,t}} \frac{\partial v_t(x)}{\partial x_i} \right) \right]}{=0}$$

$$\begin{aligned}
& + \sum_{j=1}^J \lambda_{6,k,t}(x) \left(\frac{\partial^2 f_5}{\partial u_{l,t} \partial u_{j,t}} + \sum_{i=1}^I \frac{\partial^2 b_i}{\partial u_{l,t} \partial u_{j,t}} \frac{\partial v_t(x)}{\partial x_i} \right) \\
& - \left(\sum_{i=1}^I \frac{\partial \lambda_{7,t}(x)}{\partial x_i} \frac{\partial b_{i,t}}{\partial u_{l,t}} \mu_t(x) \right).
\end{aligned}$$

2.c Discretized optimality conditions Now we discretize these conditions with respect to time and idiosyncratic states.

The idiosyncratic state is discretized by a evenly-spaced grid of size $[N_1, \dots, N_I]$ where $1, \dots, I$ are the dimensions of the state x . We assume that in each dimension there is no mass of agents outside the compact domain $[x_{i,1}, x_{i,N_i}]$. The state step size is Δx_i . We define $x^n \equiv (x_{1,n_1}, \dots, x_{i,n_i}, \dots, x_{I,n_I})$, where $n_1 \in \{1, N_1\}, \dots, n_I \in \{1, N_I\}$. We are assuming that, due to state constraints and/or reflecting boundaries, the dynamics of idiosyncratic states are constrained to the compact set $[x_{1,1}, x_{1,N_1}] \times [x_{2,1}, x_{2,N_2}] \times \dots \times [x_{I,1}, x_{I,N_I}]$. We also define $x^{n_i+1} \equiv (x_{1,n_1}, \dots, x_{i,n_i+1}, \dots, x_{I,n_I})$, $x^{n_i-1} \equiv (x_{1,n_1}, \dots, x_{i,n_i-1}, \dots, x_{I,n_I})$ $f_t^n \equiv f(x^n, u_t^n, Z_t)$, $f_t^{n_i-1} \equiv f(x^{n_i-1}, u_t^n, Z_t)$ and $f_t^{n_i+1} \equiv f(x^{n_i+1}, u_t^n, Z_t)$. I.e. the superscript n indicates a particular grid point and the superscript $n_i + 1$ and $n_i - 1$ indicate neighboring grid points along dimension i .

To discretize the problem we now replace (i) time derivatives of multipliers by backward derivatives, (ii) integrals by sums (iii) derivatives with respect to x by the upwind derivatives ∇ or $\hat{\nabla}$:

$$\begin{aligned}
\nabla_i [v_t^n] & \equiv \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{v_t^{n_i+1} - v_t^n}{\Delta x_i} + \mathbb{I}_{b_{i,t}^n < 0} \frac{v_t^n - v_t^{n_i-1}}{\Delta x_i} \right], \\
\hat{\nabla}_i [\mu_t^n] & \equiv \left[\frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0} \mu_t^{n_i+1} - \mathbb{I}_{b_{i,t}^n < 0} \mu_t^n}{\Delta x_i} + \frac{\mathbb{I}_{b_{i,t}^n > 0} \mu_t^n - \mathbb{I}_{b_{i,t}^{n_i-1} > 0} \mu_t^{n_i-1}}{\Delta x_i} \right],
\end{aligned}$$

for any discretized functions v_t^n, μ_t^n . We simplify the notation for sums $\sum_n \equiv \sum_{n_1 \in \{1, \dots, N_1\}, \dots, n_I \in \{1, \dots, N_I\}}$.

We maintain the subscript t even if it refers now to discrete time with a step Δt , that is, X_{t+1} is the shortcut for $X_{t+\Delta t}$. The second-order derivative is approximated as

$$\Delta_i [v_t^n] \equiv \left[\frac{(v_t^{n_i+1}) + (v_t^{n_i-1}) - 2(v_t^n)}{(\Delta x_i)^2} \right].$$

We start with the optimality condition for U_t

$$U_t : 0 = - \left(\frac{\lambda_{2,t} - \lambda_{2,t-1}}{\Delta t} - \varrho \lambda_{2,t} \right) \quad (112)$$

$$+ \frac{\partial f_0}{\partial U_t} - \lambda_{1,t} \frac{\partial f_1}{\partial U_t} - \lambda_{2,t} \frac{\partial f_2}{\partial U_t} - \lambda_{3,t} \frac{\partial f_3}{\partial U_t} - \lambda_{4,t} \sum_{n=1}^N \frac{\partial f_4^n}{\partial U_t} \mu_t^n \quad (113)$$

$$+ \sum_n \left[\lambda_{5,t}^n \left(\frac{\partial f_5^n}{\partial U_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial U_t} \nabla_i [v_t^n] \right) \right]$$

$$+ \sum_{j=1}^J \sum_n \left[\lambda_{6,j,t}^n \left(\frac{\partial^2 f_5^n}{\partial u_j \partial U_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial u_j \partial U_t} \nabla_i [v_t^n] \right) \right]$$

$$+ \sum_n \left[-\lambda_{7,t}^n \sum_{i=1}^I \hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial U_t} \mu_t^n \right] \right] \quad (114)$$

$$\forall t \geq 0.$$

The optimality conditions for the other aggregate variables look very much alike:

$$X_t : 0 = - \left(\frac{\lambda_{1,t} - \lambda_{1,t-1}}{\Delta} - \varrho \lambda_{1,t} \right)$$

$$+ \frac{\partial f_0}{\partial X_t} - \lambda_{1,t} \frac{\partial f_1}{\partial X_t} - \lambda_{2,t} \frac{\partial f_2}{\partial X_t} - \lambda_{3,t} \frac{\partial f_3}{\partial X_t} - \lambda_{4,t} \sum_n \frac{\partial f_4^n}{\partial X_t} \mu_t^n$$

$$+ \sum_n \left[\lambda_{5,t}^n \left(\frac{\partial f_5^n}{\partial X_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial X_t} \nabla_i [v_t^n] \right) \right]$$

$$+ \sum_{j=1}^J \sum_n \left[\lambda_{6,j,t}^n \left(\frac{\partial^2 f_5^n}{\partial u_j \partial X_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial u_j \partial X_t} \nabla_i [v_t^n] \right) \right]$$

$$+ \sum_n \left[-\lambda_{7,t}^n \sum_{i=1}^I \hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial X_t} \mu_t^n \right] \right]$$

$$\forall t > 0.$$

$$\hat{U}_t : 0 = 0$$

$$+ \frac{\partial f_0}{\partial \hat{U}_t} - \lambda_{1,t} \frac{\partial f_1}{\partial \hat{U}_t} - \lambda_{2,t} \frac{\partial f_2}{\partial \hat{U}_t} - \lambda_{3,t} \frac{\partial f_3}{\partial \hat{U}_t} - \lambda_{4,t} \sum_n \frac{\partial f_4^n}{\partial \hat{U}_t} \mu_t^n$$

$$\begin{aligned}
& + \sum_n \left[\lambda_{5,t}^n \left(\frac{\partial f_5^n}{\partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial \hat{U}_t} \nabla_i [v_t^n] \right) \right] \\
& + \sum_{j=1}^J \sum_n \left[\lambda_{6,j,t}^n \left(\frac{\partial^2 f_5^n}{\partial u_j \partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial u_j \partial \hat{U}_t} \nabla_i [v_t^n] \right) \right] \\
& + \sum_n \left[-\lambda_{7,t}^n \sum_{i=1}^I \hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial \hat{U}_t} \mu_t^n \right] \right] \\
& \forall t \geq 0.
\end{aligned}$$

$$\begin{aligned}
\tilde{U}_t : 0 = & \lambda_{4,t} \\
& + \frac{\partial f_0}{\partial \tilde{U}_t} - \lambda_{1,t} \frac{\partial f_1}{\partial \tilde{U}_t} - \lambda_{2,t} \frac{\partial f_2}{\partial \tilde{U}_t} - \lambda_{3,t} \frac{\partial f_3}{\partial \tilde{U}_t} - \lambda_{4,t} \sum_{n=1}^N \frac{\partial f_4^n}{\partial \tilde{U}_t} \mu_t^n \\
& + \sum_n \left[\lambda_{5,t}^n \left(\frac{\partial f_5^n}{\partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial \tilde{U}_t} \nabla_i [v_t^n] \right) \right] \\
& + \sum_{j=1}^J \sum_n \left[\lambda_{6,j,t}^n \left(\frac{\partial^2 f_5^n}{\partial u_j \partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial u_j \partial \tilde{U}_t} \nabla_i [v_t^n] \right) \right] \\
& + \sum_n \left[-\lambda_{7,t}^n \sum_{i=1}^I \hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial \tilde{U}_t} \mu_t^n \right] \right] \\
& \forall t \geq 0.
\end{aligned}$$

The discretized optimality condition with respect to the value function $v_t(x)$, the distribution $\mu_t(x)$ and the individual jump variable $u_{j,t}(x)$ are.

$$\begin{aligned}
v_t(x) : 0 = & -\lambda_{5,t}^n \rho - \sum_{i=1}^I \hat{\nabla}_i [\lambda_{5,t}^n b_{i,t}^n] \\
& + \frac{1}{2} \sum_{i=1}^I \sum_{k=1}^I \nabla_i [\sigma_{i,k}^n \lambda_{5,t}^n] \\
& - \sum_{j=1}^J \sum_{i=1}^I \left(\hat{\nabla}_i \left[\lambda_{6,j,t}^n \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} \right] \right) \\
& - \left(\frac{\lambda_{5,t}^n - \lambda_{5,t-1}^n}{\Delta t} - \varrho \lambda_{5,t}^n \right).
\end{aligned} \tag{115}$$

$$\begin{aligned}
\mu_t(x) : 0 = & -\lambda_{4,t} f_{4,t}^n & (116) \\
& + \lambda_{7,t}(x) \left(\sum_{i=1}^I b_i(x, u_t(x), Z_t) \nabla_i [\lambda_{7,t}^n] + \frac{1}{2} \sum_{i=1}^I (\sigma_i^2)^n \Delta_i^2 [\lambda_{7,t}^n] \right) \\
& + \frac{\lambda_{7,t}^n - \lambda_{7,t-1}^n}{\Delta t} - \varrho \lambda_{7,t}^n
\end{aligned}$$

$$\begin{aligned}
u_{l,t}(x) : 0 = & -\lambda_{4,t} \frac{\partial f_4}{\partial u_{l,t}} \mu_t^n & (117) \\
& + \sum_{j=1}^J \lambda_{6,k,t}^n \left(\frac{\partial^2 f_{5,t}^n}{\partial u_{l,t}^n \partial u_{j,t}^n} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{l,t}^n \partial u_{j,t}^n} \nabla_i [v_t^n] \right) \\
& - \sum_{i=1}^I \nabla_i [\lambda_{7,t}^n] \frac{\partial b_{i,t}^n}{\partial u_{l,t}} \mu_t^n
\end{aligned}$$

3. Discretize, then optimize We follow here the reverse approach, discretizing first and optimizing next. 3.a The discretized planner's problem

Now first discretize the optimization problem with respect to time (time step Δt) and the idiosyncratic state (N grid points, grid step Δx_i). We define the discount factor $\beta \equiv (1 + \rho\Delta t)^{-1}$.

$$\begin{aligned} \max_{Z_t, u_t^n, \mu_t^n, v_t^n} \quad & \sum_t \beta^t f_0(Z_t) \\ \text{s.t. } \forall t \quad & \\ \frac{X_{t+1} - X_t}{\Delta t} = \quad & f_1(Z_t) \end{aligned} \tag{118}$$

$$\frac{U_{t+1} - U_t}{\Delta t} = f_2(Z_t) \tag{119}$$

$$0 = f_3(Z_t) \tag{120}$$

$$\tilde{U}_t = \sum_{n=1}^N f_4(x^n, u_t^n, Z_t) \mu_t^n \tag{121}$$

$$\rho v_t^n = \frac{v_{t+1}^n - v_t^n}{\Delta t} + f_5(x^n, u_t^n, Z_t) + \sum_{i=1}^I b_i(x^n, u_t^n, Z_t) \nabla_i [v_t^n] \tag{122}$$

$$+ \frac{1}{2} \sum_{i=1}^I (\sigma_i^2)^n \Delta_i^2 [v_t^n], \quad \forall n$$

$$0 = \frac{\partial f_{5,t}^n}{\partial u_{j,t}^n} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} \nabla_i [v_t^n], \quad \forall j, n. \tag{123}$$

$$\frac{\mu_{t+1}^n - \mu_t^n}{\Delta t} = - \sum_{i=1}^I \hat{\nabla}_i [b_{i,t}^n \mu_t^n] \tag{124}$$

$$+ \frac{1}{2} \sum_{i=1}^I \Delta_i [\sigma_i^2 \mu_t^n] \tag{125}$$

$$X_0 = \bar{X}_0 \tag{126}$$

$$\mu_0^n = \bar{\mu}_0^n \tag{127}$$

3.b The Lagrangian

The Lagrangian is

$$\begin{aligned}
L = & \sum_t \beta^t f_0(Z_t) \\
& + \sum_t \beta^t \lambda_{1,t} \left\{ \frac{X_{t+1} - X_t}{\Delta t} - f_1(Z_t) \right\} \\
& + \sum_t \beta^t \lambda_{2,t} \left\{ \frac{U_{t+1} - U_t}{\Delta t} - f_2(Z_t) \right\} \\
& + \sum_t \beta^t \lambda_{3,t} \{-f_3(Z_t)\} \\
& + \sum_t \beta^t \lambda_{4,t} \left\{ \tilde{U}_t - \sum_n f_4(x^n, u_t^n, Z_t) \mu_t^n \right\} \\
& + \sum_t \sum_n \beta^t \lambda_{5,t}^n \left\{ -\rho v_t^n + \frac{v_{t+1}^n - v_t^n}{\Delta t} + f_5(x^n, u_t^n, Z_t) + \sum_{i=1}^I b_i(x^n, u_t^n, Z_t) \nabla_i [v_t^n] \right. \\
& \qquad \qquad \qquad \left. + \sum_{i=1}^I \Delta_i^2 [v_t^n] \right\} \\
& + \sum_t \sum_n \sum_{j=1}^J \beta^t \lambda_{6,j,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial u_{j,t}^n} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} \nabla_i [v_t^n] \right\} \\
& + \sum_t \sum_n \beta^t \lambda_{7,t}^n \left\{ -\frac{\mu_{t+1}^n - \mu_t^n}{\Delta t} - \sum_{i=1}^I \hat{\nabla}_i [b_{i,t}^n \mu_t^n] \right. \\
& \qquad \qquad \qquad \left. + \frac{1}{2} \sum_{i=1}^I \Delta_i [\sigma_i^2 \mu_t^n] \right\}
\end{aligned}$$

3.c The optimality conditions

The FOCs are

$$\begin{aligned}
\frac{\partial L}{\partial U_t} : 0 = & \frac{\partial f_{0,t}}{\partial U_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial U_t} + \lambda_{2,t} \left\{ -\frac{1}{\Delta t} - \frac{\partial f_{2,t}}{\partial U_t} \right\} + \beta^{-1} \lambda_{2,t-1} \frac{1}{\Delta t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial U_t} - \lambda_{4,t} \sum_n \frac{\partial f_{4,t}^n}{\partial U_t} \\
& + \sum_n \lambda_{5,t}^n \left\{ + \frac{\partial f_{5,t}^n}{\partial U_t} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \sum_{j=1}^J \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial U_t} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left\{ \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial U_t} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial U_t} \frac{\mu_t^n}{\Delta x_i} \right] \right\} \\
& \forall t \geq 0
\end{aligned}$$

$$\frac{\partial L}{\partial X_t} : 0 = \frac{\partial f_{0,t}}{\partial X_t} - \lambda_{1,t} \left\{ \frac{1}{\Delta t} + \frac{\partial f_{1,t}}{\partial X_t} \right\} + \beta^{-1} \lambda_{1,t-1} \frac{1}{\Delta t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial X_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial X_t} - \lambda_{4,t} \sum_n \frac{\partial f_{4,t}^n}{\partial X_t} \mu_t^n$$

$$\begin{aligned}
& + \sum_n \lambda_{5,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial X_t} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial X_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \sum_j \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial X_t} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial X_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left\{ \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial X_t} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial X_t} \frac{\mu_t^n}{\Delta x_i} \right] \right\} \\
& \forall t > 0
\end{aligned}$$

$$\begin{aligned}
\frac{\partial L}{\partial \tilde{U}_t} : 0 & = \frac{\partial f_{0,t}}{\partial \tilde{U}_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial \tilde{U}_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial \tilde{U}_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial \tilde{U}_t} - \lambda_{4,t} \sum_n \frac{\partial f_{4,t}^n}{\partial \tilde{U}_t} \mu_t^n \\
& + \sum_n \lambda_{5,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial \tilde{U}_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \sum_j \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial \tilde{U}_t} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial \tilde{U}_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left\{ \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial \tilde{U}_t} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial \tilde{U}_t} \frac{\mu_t^n}{\Delta x_i} \right] \right\} \\
& \forall t \geq 0
\end{aligned}$$

$$\begin{aligned}
\frac{\partial L}{\partial \hat{U}_t} : 0 & = \frac{\partial f_{0,t}}{\partial \hat{U}_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial \hat{U}_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial \hat{U}_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial \hat{U}_t} - \lambda_{4,t} \sum_n \frac{\partial f_{4,t}^n}{\partial \hat{U}_t} \mu_t^n \\
& + \sum_n \lambda_{5,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial \hat{U}_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \sum_j \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial \hat{U}_t} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial \hat{U}_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left\{ \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial \hat{U}_t} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial \hat{U}_t} \frac{\mu_t^n}{\Delta x_i} \right] \right\} \\
& \forall t \geq 0
\end{aligned}$$

$$\begin{aligned}
\frac{\partial L}{\partial v_t^n} : 0 & = \lambda_{5,t}^n \left\{ -\rho - \frac{1}{\Delta t} + \sum_{i=1}^I b_{i,t}^n \frac{\mathbb{I}_{b_{i,t}^n < 0} - \mathbb{I}_{b_{i,t}^n > 0}}{\Delta x_i} - \sum_{i=1}^I \frac{2(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \lambda_{5,t-1}^n \beta^{-1} \frac{1}{\Delta t} \\
& + \sum_{i=1}^I \lambda_{5,t}^{n_i-1} b_{i,t}^{n_i-1} \frac{\mathbb{I}_{b_{i,t}^{n_i-1} > 0}}{\Delta x_i} + \sum_{i=1}^I \lambda_{5,t}^{n_i-1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2}
\end{aligned} \tag{129}$$

$$\begin{aligned}
& - \sum_{i=1}^I \lambda_{5,t}^{n_i+1} b_{i,t}^{n_i+1} \frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0}}{\Delta x_i} + \sum_{i=1}^I \lambda_{5,t}^{n_i+1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \\
& + \sum_{j=1}^J \sum_{i=1}^I \left\{ \lambda_{6,j,t}^n \left\{ \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} \frac{\mathbb{I}_{b_{i,t}^n < 0} - \mathbb{I}_{b_{i,t}^n > 0}}{\Delta x_i} \right\} + \lambda_{6,j,t}^{n_i-1} \left\{ \frac{\partial b_{i,t}^{n_i-1}}{\partial u_{j,t}^{n_i-1}} \frac{\mathbb{I}_{b_{i,t}^{n_i-1} > 0}}{\Delta x_i} \right\} \right\} \\
& - \sum_{j=1}^J \sum_{i=1}^I \lambda_{6,j,t}^{n_i+1} \left\{ \frac{\partial b_{i,t}^{n_i+1}}{\partial u_{j,t}^{n_i+1}} \frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0}}{\Delta x_i} \right\} \\
& \forall t \geq 0
\end{aligned} \tag{130}$$

$$\begin{aligned}
\frac{\partial L}{\partial \mu_t^n} : 0 &= -\lambda_{4,t} f_{4,t}^n \\
& + \lambda_{7,t}^n \left\{ \frac{1}{\Delta t} - \sum_{i=1}^I \left[\left(\mathbb{I}_{b_{i,t}^n > 0} - \mathbb{I}_{b_{i,t}^n < 0} \right) \frac{b_{i,t}^n}{\Delta x_i} \right] - \sum_{i=1}^I \frac{-2(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \left\{ - \sum_{i=1}^I \lambda_{7,t}^{n_i-1} \left[\frac{\mathbb{I}_{b_{i,t}^n < 0} b_{i,t}^n}{\Delta x_i} \right] + \sum_{i=1}^I \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \left\{ - \sum_{i=1}^I \lambda_{7,t}^{n_i+1} \left[\frac{-\mathbb{I}_{b_{i,t}^n > 0} b_{i,t}^n}{\Delta x_i} \right] + \sum_{i=1}^I \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \beta^{-1} \lambda_{7,t-1}^n \left\{ -\frac{1}{\Delta t} \right\} \\
& \forall t > 0
\end{aligned} \tag{131}$$

$$\begin{aligned}
\frac{\partial L}{\partial u_{l,t}^n} : 0 &= -\lambda_{4,t} \frac{\partial f_{4,t}^n}{\partial u_{l,t}^n} \mu_t^n \\
& + \beta^t \lambda_{5,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial u_{l,t}^n} + \sum_{i=1}^I \frac{\partial b_{i,t}^n}{\partial u_{l,t}^n} \nabla_i [v_t^n] \right\} \\
& + \sum_j \lambda_{6,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} \nabla_i [v_t^n] \right\} \\
& + \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n_i-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial u_{l,t}^n} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n_i+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial u_{l,t}^n} \frac{\mu_t^n}{\Delta x_i} \right] \\
& \forall t \geq 0
\end{aligned} \tag{132}$$

By the individual agents' optimality condition, line 2 of this expression is equal to 0.

4. Compare Finally, by comparing the respective discretized optimality conditions, we show that the two procedures yield the same equilibrium conditions in the limit. Consider first the condition for U_t . The optimize-discretize condition is given by (112),

which we reproduce here

$$\begin{aligned}
U_t : 0 = & - \left(\frac{\lambda_{2,t} - \lambda_{2,t-1}}{\Delta} - \varrho \lambda_{2,t} \right) \\
& + \frac{\partial f_0}{\partial U_t} - \lambda_{1,t} \frac{\partial f_1}{\partial U_t} - \lambda_{2,t} \frac{\partial f_2}{\partial U_t} - \lambda_{3,t} \frac{\partial f_3}{\partial U_t} - \lambda_{4,t} \sum_{n=1}^N \frac{\partial f_4^n}{\partial U_t} \mu_t^n \\
& + \sum_n \lambda_{5,t}^n \left\{ \frac{\partial f_5^n}{\partial U_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \sum_{j=1}^J \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial U_t} + \sum_{i=1}^I \frac{\partial^2 b_t^n}{\partial u_{j,t}^n \partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left[-\lambda_{7,t}^n \sum_{i=1}^I \hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial U_t} \mu_t^n \right] \right] \\
& \forall t \geq 0
\end{aligned}$$

The discretize-optimize condition (128), rearranges to

$$\begin{aligned}
\frac{\partial L}{\partial U_t} : 0 = & - \left(\frac{\lambda_{2,t} - \lambda_{2,t-1}}{\Delta t} - \frac{\beta^{-1} - 1}{\Delta t} \lambda_{2,t-1} \right) \\
& + \frac{\partial f_{0,t}}{\partial U_t} - \lambda_{1,t} \frac{\partial f_{1,t}}{\partial U_t} - \lambda_{2,t} \frac{\partial f_{2,t}}{\partial U_t} - \lambda_{3,t} \frac{\partial f_{3,t}}{\partial U_t} - \lambda_{4,t} \sum_{n=1}^N \frac{\partial f_{4,t}^n}{\partial U_t} \mu_t^n \\
& + \sum_{n=1}^N \lambda_{5,t}^n \left\{ \frac{\partial f_{5,t}^n}{\partial U_t} + \sum_{i=1}^I \frac{\partial b_i^n}{\partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_{n=1}^N \sum_{j=1}^J \lambda_{6,j,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial U_t} + \frac{\partial^2 b_t^n}{\partial u_{j,t}^n \partial U_t} \nabla_i [v_t^n] \right\} \\
& + \sum_n \left\{ \sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n_i-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial U_t} \frac{\mu_t^n}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n_i+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial U_t} \frac{\mu_t^n}{\Delta x_i} \right] \right\} \\
& \forall t \geq 0
\end{aligned}$$

The second to fourth lines are evidently identical. The last lines also coincide once

we take into account the definition of $\hat{\nabla}_i \left[\frac{\partial b_{i,t}^n}{\partial U_t} \mu_t^n \right] = \frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0} \frac{\partial b_{i,t}^{n_i+1}}{\partial U_t} \mu_t^{n_i+1} - \mathbb{I}_{b_{i,t}^n < 0} \frac{\partial b_{i,t}^n}{\partial U_t} \mu_t^n}{\Delta x_i} + \frac{\mathbb{I}_{b_{i,t}^n > 0} \frac{\partial b_{i,t}^n}{\partial U_t} \mu_t^n - \mathbb{I}_{b_{i,t}^{n_i-1} > 0} \frac{\partial b_{i,t}^{n_i-1}}{\partial U_t} \mu_t^{n_i-1}}{\Delta x_i}$.

Finally compare the first lines. Since $\beta \equiv (1 + \varrho \Delta t)^{-1}$ we have that $\frac{\beta^{-1} - 1}{\Delta t} = \varrho$. The difference between these two equations hence is $\|\varrho (\lambda_{2,t} - \lambda_{2,t-1})\|$. In the limit as

$\Delta t \rightarrow 0$, and provided that $\lambda_{2,t}$ features no jumps for $t > 0$, this difference converges to zero. The same argument applies to the optimality conditions with respect to X_t with the difference now proportional to $\|\varrho(\lambda_{1,t} - \lambda_{1,t-1})\|$. The optimality conditions with respect to \hat{U}_t and \tilde{U}_t are identical, that is, there is no difference.

Next consider the two discretized optimality conditions with respect to v_t^n (115) and (129). After some rearranging they are given by

$$\begin{aligned}
v_t(x) : 0 = & - \sum_{i=1}^I \left(\frac{\mathbb{I}_{b_{i,t}^n > 0} \lambda_{5,j,t}^n b_{i,t}^n - \mathbb{I}_{b_{i,t}^{n-1} > 0} \lambda_{5,j,t}^{n-1} b_{i,t}^{n-1}}{\Delta x_i} + \frac{\mathbb{I}_{b_{i,t}^{n+1} < 0} \lambda_{5,j,t}^{n+1} b_{i,t}^{n+1} - \mathbb{I}_{b_{i,t}^n < 0} \lambda_{5,j,t}^n b_{i,t}^n}{\Delta x_i} \right) \\
& + \frac{1}{2} \sum_{i=1}^I \frac{(\sigma_i^2)^{n_i+1} \lambda_{5,t}^{n_i+1} + (\sigma_i^2)^{n_i-1} \lambda_{5,t}^{n_i-1} - 2(\sigma_i^2)^n \lambda_{5,t}^n}{(\Delta x_i)^2} \\
& - \sum_{j=1}^J \sum_{i=1}^I \left(\frac{\mathbb{I}_{b_{i,t}^n > 0} \lambda_{6,j,t}^n \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} - \mathbb{I}_{b_{i,t}^{n-1} > 0} \lambda_{6,j,t}^{n-1} \frac{\partial b_{i,t}^{n-1}}{\partial u_{j,t}^{n-1}}}{\Delta x_i} + \frac{\mathbb{I}_{b_{i,t}^{n+1} < 0} \lambda_{6,j,t}^{n+1} \frac{\partial b_{i,t}^{n+1}}{\partial u_{j,t}^{n+1}} - \mathbb{I}_{b_{i,t}^n < 0} \lambda_{6,j,t}^n \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n}}{\Delta x_i} \right) \\
& - \lambda_{5,t}^n \rho - \left(\frac{\lambda_{5,t}^n - \lambda_{5,t-1}^n}{\Delta t} - \varrho \lambda_{5,t}^n \right)
\end{aligned}$$

and

$$\begin{aligned}
\frac{\partial L}{\partial v_t^n} : 0 = & \lambda_{5,t}^n \left\{ \sum_{i=1}^I b_{i,t}^n \frac{\mathbb{I}_{b_{i,t}^n < 0} - \mathbb{I}_{b_{i,t}^n > 0}}{\Delta x_i} - \sum_{i=1}^I \frac{(\sigma_i^2)^n}{(\Delta x_i)^2} \right\} \\
& + \left\{ \sum_{i=1}^I \lambda_{5,t}^{n_i-1} b_{i,t}^{n_i-1} \frac{\mathbb{I}_{b_{i,t}^{n_i-1} > 0}}{\Delta x_i} + \sum_{i=1}^I \lambda_{5,t}^{n_i-1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \left\{ - \sum_{i=1}^I \lambda_{5,t}^{n_i+1} b_{i,t}^{n_i+1} \frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0}}{\Delta x_i} + \sum_{i=1}^I \lambda_{5,t}^{n_i+1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& + \sum_{j=1}^J \sum_{i=1}^I \left(\lambda_{6,j,t}^n \frac{\partial b_{i,t}^n}{\partial u_{j,t}^n} \frac{\mathbb{I}_{b_{i,t}^n < 0} - \mathbb{I}_{b_{i,t}^n > 0}}{\Delta x_i} + \lambda_{6,j,t}^{n_i-1} \frac{\partial b_{i,t}^{n_i-1}}{\partial u_{j,t}^{n_i-1}} \frac{\mathbb{I}_{b_{i,t}^{n_i-1} > 0}}{\Delta x_i} - \lambda_{6,j,t}^{n_i+1} \frac{\partial b_{i,t}^{n_i+1}}{\partial u_{j,t}^{n_i+1}} \frac{\mathbb{I}_{b_{i,t}^{n_i+1} < 0}}{\Delta x_i} \right) \\
& - \rho \lambda_{5,t}^n - \left(\frac{\lambda_{5,t}^n - \lambda_{5,t-1}^n}{\Delta t} - \frac{\beta^{-1} - 1}{\Delta t} \lambda_{5,t-1}^n \right) \tag{133}
\end{aligned}$$

Again these, two expressions are identical up to the last time index in the last line (λ_5^n), and thus the difference is $\|\varrho(\lambda_{5,t} - \lambda_{5,t-1})\|$.

Next, consider the two discretized optimality conditions with respect to μ_t^n (116) and (131). After some rearranging they are given by

$$\begin{aligned}
\mu_t(x) : 0 = & -\lambda_{4,t} f_{4,t}^n \tag{134} \\
& + \sum_{i=1}^I b_{i,t}^n \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\lambda_{7,t}^{n_i+1} - \lambda_{7,t}^n}{\Delta x_i} + \mathbb{I}_{b_{i,t}^n < 0} \frac{\lambda_{7,t}^n - \lambda_{7,t}^{n_i-1}}{\Delta x_i} \right] + \frac{1}{2} \sum_{i=1}^I (\sigma_i^2)^n \frac{\lambda_{7,t}^{n_i+1} + \lambda_{7,t}^{n_i-1} - 2\lambda_{7,t}^n}{2(\Delta x_i)^2}
\end{aligned}$$

$$\begin{aligned}
& + \frac{\lambda_{7,t}^n - \lambda_{7,t-1}^n}{\Delta t} - \varrho \lambda_{7,t}^n \\
\frac{\partial L}{\partial \mu_t^n} : 0 & = -\lambda_{4,t} f_{4,t}^n \\
& + \lambda_{7,t}^n \left\{ - \sum_{i=1}^I \left[\left(\mathbb{I}_{b_{i,t}^n > 0} - \mathbb{I}_{b_{i,t}^n < 0} \right) \frac{b_{i,t}^n}{\Delta x_i} \right] - \sum_{i=1}^I \frac{-2(\sigma_i^2)^n}{2(\Delta x_i)^2} \right\} \\
& - \sum_{i=1}^I \left[\lambda_{7,t}^{n_i-1} \frac{\mathbb{I}_{b_{i,t}^n < 0} b_{i,t}^n}{\Delta x_i} \right] + \sum_{i=1}^I \lambda_{7,t}^{n_i-1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \\
& + \sum_{i=1}^I \left[\lambda_{7,t}^{n_i+1} \frac{\mathbb{I}_{b_{i,t}^n > 0} b_{i,t}^n}{\Delta x_i} \right] + \sum_{i=1}^I \lambda_{7,t}^{n_i+1} \frac{(\sigma_i^2)^n}{2(\Delta x_i)^2} \\
& + \frac{\lambda_{7,t}^n - \lambda_{7,t-1}^n}{\Delta t} - \frac{\beta^{-1} - 1}{\Delta t} \lambda_{7,t-1}^n,
\end{aligned} \tag{135}$$

which again differ in $\|\varrho(\lambda_{7,t} - \lambda_{7,t-1})\|$.

Finally, consider the two discretized optimality conditions with respect to $u_{l,t}^n(x)$, (117) and (132). After some rearranging they are given by

$$\begin{aligned}
u_{l,t}(x) : 0 & = -\lambda_{4,t} \frac{\partial f_4}{\partial u_{l,t}} \mu_t^n \\
& + \sum_{j=1}^J \lambda_{6,l,t}^n \left(\frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{v_t^{n_i+1} - v_t^n}{\Delta x_i} + \mathbb{I}_{b_{i,t}^n < 0} \frac{v_t^n - v_t^{n_i+1}}{\Delta x_i} \right] \right) \\
& - \sum_{i=1}^I \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{\lambda_{7,t}^{n_i+1} - \lambda_{7,t}^n}{\Delta x_i} + \mathbb{I}_{b_{i,t}^n < 0} \frac{\lambda_{7,t}^n - \lambda_{7,t}^{n_i-1}}{\Delta x_i} \right] \frac{\partial b_{i,t}^n}{\partial u_{l,t}} \mu_t^n
\end{aligned} \tag{136}$$

$$\begin{aligned}
\frac{\partial L}{\partial u_{l,t}^n} : 0 & = -\lambda_{4,t} \frac{\partial f_{4,t}^n}{\partial u_{l,t}^n} \mu_t^n \\
& + \sum_j \lambda_{6,t}^n \left\{ \frac{\partial^2 f_{5,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} + \sum_{i=1}^I \frac{\partial^2 b_{i,t}^n}{\partial u_{j,t}^n \partial u_{l,t}^n} \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{v_t^{n_i+1} - v_t^n}{\Delta x_i} + \mathbb{I}_{b_{i,t}^n < 0} \frac{v_t^n - v_t^{n_i+1}}{\Delta x_i} \right] \right\} \\
& + \left[\sum_{i=1}^I (\lambda_{7,t}^n - \lambda_{7,t}^{n_i-1}) \left[\mathbb{I}_{b_{i,t}^n < 0} \frac{1}{\Delta x_i} \right] + \sum_{i=1}^I (\lambda_{7,t}^{n_i+1} - \lambda_{7,t}^n) \left[\mathbb{I}_{b_{i,t}^n > 0} \frac{1}{\Delta x_i} \right] \right] \frac{\partial b_{i,t}^n}{\partial u_{l,t}} \mu_t^n,
\end{aligned}$$

which are identical. To summarize, whether one discretize the optimality conditions of the planner and then discretizes them, or one discretizes the planner's problem and then derives the optimality conditions, one arrives to a set of optimality conditions that coincide in everything but the timing of the multiplier in the term $\varrho \lambda_t$. Provided that multipliers experience no jumps, the difference between the two approaches goes to 0 as $\Delta t \rightarrow 0$. Note that this issue has nothing to do with heterogeneity.

D.4 Solving the Nuño and Thomas model using Dynare

Here we apply the “discretize-optimize” methodology outlined in Section D to the heterogeneous-agent model introduced in Nuño and Thomas (2022). This is a model *à la* Aiyagari-Bewley-Huggett with non-state-contingent long-term nominal debt contracts. Finding the optimal policy in this problem requires that the central bank takes into account not only the dynamics of the state distribution (given by the KF equation) but also the HJB equation. Figure 16 displays the time-0 optimal policy (inflation) in this case, compared to the one obtained through the “optimize-discretize” methodology employed in Nuño and Thomas (2022). Optimal inflation coincides in both cases, up to a numerical error that is reduced as we increase the number of grid points and we reduce the time step.

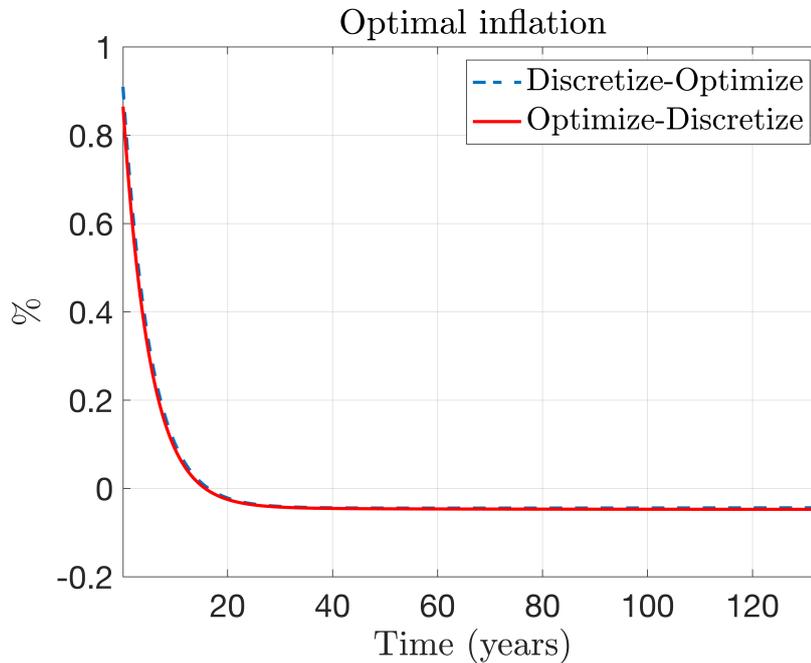


Figure 16: Time-0 optimal monetary policy using the two approaches.

Notes: The figure shows the optimal path of inflation in the Nuño and Thomas (2022) model using the “discretize-optimize” and “optimize-discretize” methods.

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