# The Long Shadows of the Great Inflation Evidence from Residential Mortgages 

Matthew J. Botsch
Bowdoin College

Ulrike Malmendier
UC Berkeley

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## Mortgage Choice

- Major puzzle in mortgage markets (and financial contracting more broadly): consumers' aversion to adjustable-rate loans.
- Empirical contract mix in US: 80\% fixed-rate.
- Inconsistent with standard life-cycle consumption models (e.g., Campbell and Cocco 2003, 2015).
- Especially at high price, about 17obp above comparable variable-rate mortgages.
- Our own calculations (below): far more households choose FRMs than the standard economic model predicts, esp. in the wake of the Great Inflation: Baby Boomers should have taken out 1m fewer FRMs in the late 1980s, and 0.5 m fewer in the late 1990 s .


## Mortgage Choice

- Puzzling because: Cost of these deviations large. Given expected refinancing behavior and mobility, Baby Boomers overpayed $>\$ 14$ billion on their FRMS in the late 1980s, and almost $\$ 9$ billion in the late 1990 s.
- Puzzling because: Home purchase and financing one of the biggest financial decisions for many households.
- ARM-type contracts have high market shares in other countries (Australia, Belgium, Chile, Estonia, Finland, Greece, Hungary, Ireland, Israel, Korea, Luxembourg, Mexico, Norway, Poland, Portugal, Slovakia, Slovenia, Spain, Switzerland, Turkey).
- Idea here: Role of "experience effects" in past inflation.


## Idea

- Individuals overweight prior lifetime experiences when forming expectations.
- Cf. availability bias (Tversky and Kahneman 1974): "more memorable events are processed as more likely events" when forming beliefs.
- Underpinning: synaptic tagging (cf. Laudenbach, Niessen-Ruenzi, Malmendier AEA P\&P 2018) - personal experiences rewire our "hardware", especially experiences that are anchored more strongly due to emotions.


## Many applications

1. Political attitudes: Alesina \& Fuchs-Schundeln (2007)
2. Medical diagnoses: Weber et al. (1993); Hertwig et al. (2004)
3. Climate change: Deryugina (2013)
4. Stock-market participation: Malmendier \& Nagel (2011)
5. Consumption behavior: Malmendier \& Shen (2015)
6. Expected inflation: Malmendier \& Nagel (2016)

## Inflation experiences \& expectations

From Malmendier \& Nagel (2016) / Mich. Survey:


## Inflation experiences \& expectations

From Malmendier \& Nagel (2016) / Mich. Survey:


## This Paper

- Overweighting lifetime inflation experiences generates differences of opinion about the value of future inflation rates and hence nominal interest rates (Fisher equation).
$>$ Those with higher lifetime experiences of inflation will expect higher nominal interest rates.
- Overweighting lifetime inflation experiences generates differences of opinion about the value of fixed-rate assets (relative to variable/real-rate assets).
> Those with higher lifetime experiences of inflation will overvalue and overpay for fixed-rate mortgage contracts, relative to the full-information optimum.
- We assess the implications of experience-based beliefs for mortgage choice, and we provide quantitative estimates of the costs.


## Preview of Results

1. Individuals' inflation experiences significantly affect beliefs about future nominal interests.
2. Individuals' inflation experiences significantly affect mortgage choice.

- Individuals with high experienced inflation are more likely to choose FRMs (within year).
- 1 in 6 HHs choose FRMs over ARMs because of $\pi^{e}$

3. The costs of overweighting are large.

- Ex ante: individuals pay 6-14 basis points for every additional pp. of $\pi^{e}$
- Ex post: switching HHs overpay by $\$ 8,000-16,000$ (over expected tenure, in after-tax PV)
> Concentrated among Baby Boomers: overpaid in aggregate by $\$ 14$ billion on FRMs in 1980 s , $\$ 9$ billion in 1990 .


## Data and Methodology

## Learning from Experiences

Experience effect hypothesis: individuals learn from lifetime experiences
Prior empirical evidence suggests that HH n's experience-based inflation forecast at time $t$ is approximately:

$$
\pi_{n, t}^{e} \propto \sum_{j=0}^{\text {age }_{n}}\left(\frac{\operatorname{age}_{n}-j}{\operatorname{age}_{n}}\right) \pi_{t-j}
$$

where lag $j$ runs from today $(j=0)$ to birth $\left(j=\right.$ age $\left._{n}\right)$.
Empirical content: cross-sectional heterogeneity of forecasts (by householder age).

Inflation Experiences in 1985


Inflation Experiences in 1985


Inflation Experiences in 2000


Inflation Experiences in 2000


## Three Steps

1. Overweighting lifetime inflation experiences and nominal interest rates.
$>$ Those with higher lifetime experiences of inflation will expect higher nominal interest rates.
> SCF data
2. Overweighting lifetime inflation experiences choice of fixed-rate mortgages.
$>$ Those with higher lifetime experiences of inflation will overvalue and overpay for fixed-rate mortgage contracts, relative to the full-information optimum.
$>$ RFS data (and BLS, PMMS)
3. Quantitative estimates of the costs.

## Inflation experiences \& Interest Rate expectations




## Mortgage Data

Residential Finance Survey: decennial Census Bureau survey of households, cross-referenced with servicers, in 1991 and 2001.

- Microdata on outstanding mortgages linked to 1-4 unit, owner-occupied properties:
$\square$ FRM/ARM status
Loan terms \& property value
HH income \& demographics
$\square$ Census region
- Missing recent movers
- We subset on mortgages originated $\leq 6$ years prior.

| Table 1: Summary Statistics |  |  |  |
| :--- | :---: | :---: | :---: |
|  | FRM | ARM | FRM - ARM |
|  | 12,416 | 2,245 |  |
| Contract Characteristics |  |  |  |
| Current rate (bps) | 972.7 | 924.5 | $48.2^{*}$ |
| Initial rate (bps) | $"$ | 876.2 | $96.4^{*}$ |
| Margin (bps) | n.a. | 282.7 | $\mathrm{n} . \mathrm{a}$. |
| Term (years) | 23.2 | 26.1 | $-2.9^{*}$ |
| Loan Amount (2000 \$k) | 102.0 | 140.3 | $-38.3^{*}$ |
| Borrower Characteristics |  |  |  |
| Primary owner age | 41.4 | 41.8 | -0.4 |
| Non-white | 0.136 | 0.099 | $0.037^{*}$ |
| First-time owner | 0.413 | 0.348 | $0.065^{*}$ |
| Total income (2000 \$) | 75,177 | 84,165 | $-8,989^{*}$ |
| Other Loan Characteristics |  |  |  |
| Junior mortgage | 0.129 | 0.086 | $0.043^{*}$ |
| Non-conventional | 0.211 | 0.061 | $0.150^{*}$ |
| Refi | 0.256 | 0.244 | 0.012 |
| Loan / income | 1.73 | 2.04 | $-0.31^{*}$ |
| Loan / value $\times$ 100 | 81.7 | 90.0 | $-8.3^{*}$ |
| Jumbo loan? | 0.043 | 0.127 | $-0.084^{*}$ |
| Notes. Sample of mortgages $<=6$ years old at time of 1991 and 2001 |  |  |  |
| Residential Finance Surveys of homeowner properties. Statistics are |  |  |  |
| based on available cases. * p<0.05. |  |  |  |
| Botsch \& Malmendier (2019): Long Shadows ofthe Great Inflation |  |  |  |

## Other Data

Inflation: log changes in CPI-U (from BLS)
Primary Mortgage Market Survey index rates for FRM and ARM (from Freddie Mac)

- Representative, nationwide survey of mortgage originators
- Quotes interest rates on first-lien, prime, conventional, conforming, 30-year loans with LTV $=0.8$
$>$ FRM and $1 / 1$ ARM
- Reweight from 5 Freddie Mac regions to 4 Census regions using 1990 Census state housing counts.
- Annual average of weekly data


## Path of PMMS Interest Rates



## Identification

Identification from cross-sectional differences in inflation experiences + their evolution over time (time series). This rules out:

1. Time-specific effects unrelated to learning from experiences.
$>$ Time dummies capture the effect of all individuals learning from the full historical inflation data, including current inflation.
2. Life-cycle effects
$>$ Age is separately identified in repeated XS data
$\beta_{\pi, F R M}$ picks up influence of remaining, individual heterogeneity in inflation experiences on expectations.
$>$ If experiences don't matter, coefficient should be zero.

FRM Share and Experienced Inflation by Age Group


## Econometric Model

McFadden (1974): Indirect utility of $\mathrm{HH} n$ considering alternative $i$ (FRM or ARM):

$$
U_{n i}=\alpha_{i t}+\beta_{R} \text { Rate }_{n i}+\beta_{\pi, i} \pi_{n}^{e}+x_{n}^{\prime} \delta_{i}+\varepsilon_{n i}
$$

Alternative $i$ is chosen iff $U_{n i}>U_{n j} \forall j \neq i$.

## Predictions:

1. $\beta_{R}<0$ - price elasticity of demand is negative
2. $\beta_{\pi, F R M}>0$ - inflation histories raise the FRM share (learning-from-experiences effect)

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$$

Alternative $i$ is chosen iff $U_{n i}>U_{n j} \forall j \neq i$.

## Missing Data Problem:

We only observe the rate of the chosen alternative:

$$
\text { Rate }_{n}=D_{n} \text { Rate }_{n, F R M}+\left(1-D_{n}\right) \text { Rate }_{n, A R M}
$$

## Three-Step Estimation Procedure

1. Reduced-form mortgage choice model using Freddie Mac FRM \& ARM index rates (by year-region)

$$
U_{n i}=\alpha_{i t}+\tilde{\beta}_{R} P M M S \text { Rate }_{n i}+\beta_{\pi, i} \pi_{n}^{e}+x_{n}^{\prime} \delta_{i}+\tilde{\varepsilon}_{n i}
$$

2. Mortgage pricing equations, correcting for any selection bias using choice probabilities from Step 1-Heckman (1979), Powell (1984), Newey (2009)

$$
\text { Rate }_{n i}=\gamma_{R} \text { PMMSRate }_{n i}+x_{n}^{\prime} \gamma_{i}+v_{n i}
$$

3. Structural mortgage choice model using individual-level predicted interest rates for each alternative from Step 2

$$
U_{n i}=\alpha_{i t}+\beta_{R} \widehat{\operatorname{Rate}_{n i}}+\beta_{\pi, i} \pi_{n}^{e}+x_{n}^{\prime} \delta_{i}+\varepsilon_{n i}
$$

$n$ : household; $i$ : FRM or ARM.

## Results

# Table 2: Reduced-Form 

 Mortgage Choice Model|  | (4) | (5) |
| :--- | :---: | :---: |
| FRM Alternative-Specific Characteristics |  |  |
| Freddie Mac PMMS FRM | $-3.33^{* * *}$ | $-3.59^{* * *}$ |
| index rate (\%) | $(0.575)$ | $(0.816)$ |
| Experienced inflation (\%) | $0.254^{* * *}$ | $0.187^{*}$ |
|  | $(0.086)$ | $(0.098)$ |
| Log(Income) | $0.0276^{* *}$ | $0.0278^{* *}$ |
|  | $(0.012)$ | $(0.012)$ |

ARM Alternative-Specific Characteristics

| Freddie Mac PMMS ARM | $-0.768^{* * *}$ | $-0.844^{* * *}$ |
| :--- | :---: | :---: |
| initial rate index (\%) | $(0.250)$ | $(0.314)$ |
| Alternative-specific constants | YES | YES |
| Origination year FE | YES | YES |
| Other controls | YES | YES |
| Number of Choice Situations | 15,051 | 14,337 |
| Number of Alternatives | 3 | 2 |

Notes. Multinomial logit coefficients shown (robust SEs).
*** $\mathrm{p}<0.01$, ** $\mathrm{p}<0.05$, * $\mathrm{p}<0.1$

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ARM Alternative-Specific Characteristics

## WTP calculation:

$\beta_{\pi^{e}}=0.254$ $\beta_{\text {Rate }, \text { FRM }}=-3.33$
$\Rightarrow W T P=-\frac{\beta_{\pi^{e}}}{\beta_{\text {Rate,FRM }}}$
$=7.6 \mathrm{bps}$
$(S E=2.9 \mathrm{bps})$

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## Actual and Counterfactual FRM Shares



Table 4: Structural Logit Model of Mortgage Choice

|  | (1) | (2) | (3) | (4) | (5) | (6) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Step 2 Selection Correction? | No | Yes | No | Yes | No | Yes |
| FRM Rate Offered <br> Initial ARM Rate Offered | $\begin{aligned} & \hline 0.764 \\ & (0.74) \\ & -0.368 \\ & (0.62) \\ & \hline \end{aligned}$ | $\begin{gathered} \hline \hline-1.474^{* *} \\ (0.58) \\ 1.280^{* *} \\ (0.54) \end{gathered}$ |  |  |  |  |
| ARM Margin Offered |  |  |  |  |  |  |
| Experienced inflation (\%) | $\begin{gathered} 0.237 * * \\ (0.09) \end{gathered}$ | $\begin{aligned} & 0.181^{*} \\ & (0.10) \end{aligned}$ |  |  |  |  |
| Log(Income) | $\begin{gathered} 0.00221 \\ (0.02) \end{gathered}$ | $\begin{gathered} -0.00875 \\ (0.03) \end{gathered}$ |  |  |  |  |
| Age | $\begin{aligned} & -0.015 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.02) \end{aligned}$ |  |  |  |  |
| Age ${ }^{2} / 100$ | $\begin{aligned} & 0.018 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & -0.005 \\ & (0.02) \end{aligned}$ |  |  |  |  |
| Joint owners | $\begin{aligned} & 0.144 \\ & (0.12) \end{aligned}$ | $\begin{gathered} -0.074 \\ (0.13) \end{gathered}$ |  |  |  |  |
| Rural county | $\begin{aligned} & -0.053 \\ & (0.32) \end{aligned}$ | $\begin{gathered} -0.776^{* *} \\ (0.35) \end{gathered}$ |  |  |  |  |
| Non-conventional |  |  |  |  |  |  |
| Origination year FE | YES | YES | YES | YES | YES | YES |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

Notes. Binomial logit coefficient estimates. Dependent variable is $1=$ FRM, $0=$ ARM. Bootstrapped standard errors in parentheses, adjusting for 1st- and 2nd-step estimation, from 200 repetitions. ${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$

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| Non-conventional |  |  |  |  |  |  |
| Origination year FE | YES | YES | YES | YES | YES | YES |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

Notes. Binomial logit coefficient estimates. Dependent variable is $1=$ FRM, $0=$ ARM. Bootstrapped standard errors in parentheses, adjusting for 1st- and 2nd-step estimation, from 200 repetitions. ${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$

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| Step 2 Selection Correction? | No | Yes | No | Yes | No | Yes |
| FRM Rate Offered | 0.764 | $-1.474^{* *}$ |  |  |  |  |
|  | $(0.74)$ | $(0.58)$ |  |  |  |  |
| Initial ARM Rate Offered | -0.368 | $1.280^{* *}$ |  |  |  |  |
|  | $(0.62)$ | $(0.54)$ |  |  |  |  |
|  |  |  |  |  |  |  |

ARM Margin Offered

| Experienced inflation (\%) | $0.237^{* *}$ <br> $(0.09)$ | $0.181^{*}$ <br> $(0.10)$ |
| :--- | :---: | :---: |
| Log(Income) | 0.00221 | -0.00875 |
|  | $(0.02)$ | $(0.03)$ |
| Age | -0.015 | 0.004 |
|  | $(0.02)$ | $(0.02)$ |
| Age $^{2} / 100$ | 0.018 | -0.005 |
|  | $(0.02)$ | $(0.02)$ |
| Joint owners | 0.144 | -0.074 |
|  | $(0.12)$ | $(0.13)$ |
| Rural county | -0.053 | $-0.776^{* *}$ |
|  | $(0.32)$ | $(0.35)$ |

Non-conventional

| Origination year FE | YES | YES | YES | YES | YES | YES |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

Notes. Binomial logit coefficient estimates. Dependent variable is $1=$ FRM, $0=$ ARM. Bootstrapped standard errors in parentheses, adjusting for 1st- and 2nd-step estimation, from 200 repetitions. ${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$

Table 4: Structural Logit Model of Mortgage Choice


Non-conventional

| Origination year FE | YES | YES | YES | YES | YES | YES |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

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| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Step 2 Selection Correction? | No | Yes | No | Yes | No | Yes |
| FRM Rate Offered | $\begin{aligned} & \hline 0.764 \\ & (0.74) \end{aligned}$ | $\begin{gathered} \hline \hline-1.474^{* *} \\ (0.58) \end{gathered}$ | $\begin{aligned} & \hline \hline-0.127 \\ & (0.60) \end{aligned}$ | $\begin{gathered} \hline \hline-1.272 * * * \\ (0.45) \end{gathered}$ |  |  |
| Initial ARM Rate Offered | $\begin{aligned} & -0.368 \\ & (0.62) \end{aligned}$ | $\begin{gathered} 1.280^{* *} \\ (0.54) \end{gathered}$ | $\begin{aligned} & 0.838 \\ & (0.55) \end{aligned}$ | $\begin{gathered} 1.196 * * * \\ (0.38) \end{gathered}$ |  |  |
| ARM Margin Offered |  |  | $\begin{gathered} -2.364^{* * *} \\ (0.55) \\ \hline \end{gathered}$ | $\begin{aligned} & -0.302 \\ & (0.47) \end{aligned}$ |  |  |
| Experienced inflation (\%) | $\begin{gathered} \hline 0.237^{* *} \\ (0.09) \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 0.181^{*} \\ & (0.10) \\ & \hline \end{aligned}$ | $\begin{gathered} \hline 0.222^{* *} \\ (0.10) \\ \hline \end{gathered}$ | $\begin{aligned} & 0.180^{*} \\ & (0.10) \\ & \hline \end{aligned}$ |  |  |
| Log(Income) | $\begin{gathered} 0.00221 \\ (0.02) \end{gathered}$ | $\begin{gathered} -0.00875 \\ (0.03) \end{gathered}$ | $\begin{gathered} \hline-0.0572 \\ (0.04) \end{gathered}$ | $\begin{gathered} \hline-0.0171 \\ (0.04) \end{gathered}$ |  |  |
| Age | $\begin{aligned} & -0.015 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.02) \end{aligned}$ | $\begin{gathered} -0.007 \\ (0.02) \end{gathered}$ | $\begin{aligned} & 0.004 \\ & (0.02) \end{aligned}$ |  |  |
| Age ${ }^{2} / 100$ | $\begin{aligned} & 0.018 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & -0.005 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & 0.010 \\ & (0.02) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.02) \end{aligned}$ |  |  |
| Joint owners | $\begin{aligned} & 0.144 \\ & (0.12) \end{aligned}$ | $\begin{gathered} -0.074 \\ (0.13) \end{gathered}$ | $\begin{aligned} & 0.035 \\ & (0.15) \end{aligned}$ | $\begin{aligned} & -0.062 \\ & (0.12) \end{aligned}$ |  |  |
| Rural county | $\begin{aligned} & -0.053 \\ & (0.32) \end{aligned}$ | $\begin{gathered} -0.776 * * \\ (0.35) \end{gathered}$ | $\begin{gathered} -0.860^{* *} \\ (0.36) \end{gathered}$ | $\begin{gathered} -0.761^{* * *} \\ (0.28) \end{gathered}$ |  |  |
| Non-conventional |  |  |  |  |  |  |
| Origination year FE | YES | YES | YES | YES | YES | YES |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

Notes. Binomial logit coefficient estimates. Dependent variable is $1=$ FRM, $0=$ ARM. Bootstrapped standard errors in parentheses, adjusting for 1st- and 2nd-step estimation, from 200 repetitions. ${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$

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|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Step 2 Selection Correction? | No | Yes | No | Yes | No | Yes |
|  | FRM Rate Offered | 0.764 | $-1.474^{* *}$ | -0.127 | $-1.272^{* * *}$ | -0.575 |
|  | $(0.74)$ | $(0.58)$ | $(0.60)$ | $(0.45)$ | $(0.45)$ | $\left(0.692^{*}\right.$ |
| Initial ARM Rate Offered | -0.368 | $1.280^{* *}$ | 0.838 | $1.196^{* * *}$ | 0.184 | 0.593 |
|  | $(0.62)$ | $(0.54)$ | $(0.55)$ | $(0.38)$ | $(0.35)$ | $(0.39)$ |
| ARM Margin Offered |  |  | $-2.364^{* * *}$ | -0.302 | $3.738^{* * *}$ | $2.600^{* *}$ |
|  |  |  | $(0.55)$ | $(0.47)$ | $(1.03)$ | $(1.22)$ |
| Experienced inflation (\%) | $0.237^{* *}$ | $0.181^{*}$ | $0.222^{* *}$ | $0.180^{*}$ | $0.181^{*}$ | $0.192^{* *}$ |
|  | $(0.09)$ | $(0.10)$ | $(0.10)$ | $(0.10)$ | $(0.10)$ | $(0.10)$ |
| Log(Income) | 0.00221 | -0.00875 | -0.0572 | -0.0171 | $0.0798^{*}$ | 0.0916 |
|  | $(0.02)$ | $(0.03)$ | $(0.04)$ | $(0.04)$ | $(0.05)$ | $(0.06)$ |
| Age | -0.015 | 0.004 | -0.007 | 0.004 | 0.007 | 0.015 |
|  | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ |
| Age ${ }^{2} /$ 100 | 0.018 | -0.005 | 0.010 | -0.004 | -0.006 | -0.014 |
|  | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ | $(0.02)$ |
| Joint owners | 0.144 | -0.074 | 0.035 | -0.062 | 0.101 | 0.183 |
|  | $(0.12)$ | $(0.13)$ | $(0.15)$ | $(0.12)$ | $(0.16)$ | $(0.20)$ |
| Rural county | -0.053 | $-0.776^{* *}$ | $-0.860^{* *}$ | $-0.761^{* * *}$ | 0.106 | -0.375 |
|  | $(0.32)$ | $(0.35)$ | $(0.36)$ | $(0.28)$ | $(0.33)$ | $(0.40)$ |
| Non-conventional |  |  |  |  | $3.744^{* * *}$ | $4.736^{* *}$ |
|  |  |  |  |  | $(0.59)$ | $(2.16)$ |
| Origination year FE | YES | YES | YES | YES | YES |  |
| Number of Choice Situations | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 | 14,337 |

Notes. Binomial logit coefficient estimates. Dependent variable is $1=$ FRM, $0=$ ARM. Bootstrapped standard errors in parentheses, adjusting for 1st- and 2nd-step estimation, from 200 repetitions. *** $\mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05$, * $\mathrm{p}<0.1$

## Simulation of ex post Costs

## Simulation Details

We simulate the ex post payments each household would make.
Standard contract types:

- 30-year term
- Self-amortizing, level payment FRM
- 1/1 ARM indexed to 1-yearTreasury
- No early payments or defaults
- Predicted interest rates $\left(\widehat{\text { Rate }}_{n, F R M}, \widehat{\text { Rate }}_{n, A R M}\right)$

Time horizon:

- Survey year
- If held to 5, 10, 15 years


## Simulation Details

## How to model refinancing behavior?

1. No Refinancing: borrower holds FRM until maturity
2. Expected Refinancing: empirical model from Andersen, Campbell, Meisner-Nielsen, Ramadorai (2014): $P\left(\right.$ refi $\left.\mid i_{0}\right)$

$$
=\Phi\left\{-1.921+\exp \left(-1.033 \times\left(O T-\left(i-i_{0}\right)\right)\right)\right\}
$$

3. Optimal Refinancing: follow Agarwal, Driscoll, Laibson (2013) square-root rule for Optimal Threshold:

$$
\text { Refinance iff } i-i_{0}<O T
$$

$$
O T \approx-\sqrt{\frac{\sigma \kappa}{M(1-\tau)} \sqrt{2(\rho+\lambda)}}
$$

## FRM Rate for Mortgage ID 500



## Table 6: Interest Payments for a Sample Household

PV of interest payments, discounted at $8 \%$ per year. The loan was for $\$ 204,844$ in constant year 2000 \$. Marginal tax rate $=0.25 ; \mathrm{T}=25$ years.

|  | FRM (\$) |  |  | ARM (\$) |
| :---: | :---: | :---: | :---: | :---: |
|  | No Refi | Expected Refi | Optimal Refi |  |
| PDV | 235,498 | 199,637 | 193,659 | 163,074 |
| - Int. Deduct. | -58,874 | -49,909 | -48,415 | -40,768 |
| + Refi Cost | 0 | 4,633 | 3,895 | 0 |
| Total | 176,623 | 154,361 | 149,139 | 122,305 |

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## Simulated dollar cost of choosing an FRM:

- No Refi: \$176K - \$122K = \$54,000
- Expected Refi: \$154K - \$122K = \$32,000
- Optimal Refi: \$149K - \$122K = \$27,000


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Welfare-Relevant TE: weighted average over all households (using switching probabilities)

## Aggregate Cost Calculations

What is the ex post cost of choosing an FRM for marginal households?

- We calculate a "welfare-relevant treatment effect":

$$
\underset{\text { potential outcomes }}{\mathbb{E}[\underbrace{Y_{n, F R M}-Y_{n, A R M}}_{\text {potential treatments }} \mid \underbrace{D_{n}\left(\beta_{\pi}\right)=1, D_{n}(0)=0}_{n}]}
$$

- Intuition: difference in actual and counterfactual mortgage payments for the subpopulation of nearlyindifferent HHs


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\quad \int \Delta y \cdot f\left(\Delta y \mid D_{n}\left(\beta_{\pi}\right)=1, D_{n}(0)=0\right) d \Delta y
\end{gathered}
$$

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= \\
\int \Delta y \cdot \frac{h\left(D_{n}\left(\beta_{\pi}\right)=1, D_{n}(0)=0 \mid \Delta y\right) \cdot f(\Delta y)}{g\left(D_{n}\left(\beta_{\pi}\right)=1, D_{n}(0)=0\right)} d \Delta y
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\end{gathered}
$$

can be estimated by

$$
\propto \frac{1}{N} \sum_{n=1}^{N}\left(\hat{Y}_{n, F R M}-\hat{Y}_{n, A R M}\right) \cdot\left(\hat{P}\left(D_{n}\left(\hat{\beta}_{\pi}\right)=1\right)-\hat{P}\left(D_{n}(0)=0\right)\right)
$$

Table 7: Additional Interest Paid Due to Inflation Experiences
Scenario 2: Risk-adjusted rates, seniority-adjusted ARM margins

| Time Horizon: | Survey Year | 5 years | 10 years | 15 years |
| :---: | :---: | :---: | :---: | :---: |
| After-tax PDV: (all in \$) |  |  |  |  |
| No Refi | 5,674 | 10,124 | 19,126 | 27,345 |
| Expected Refi | - | 10,056 | 15,886 | 20,505 |
| Optimal Refi | - | 9,455 | 14,460 | 18,639 |
| \% switching households | 13.5 | 13.5 | 13.5 | 13.5 |


| Scenario 3: Risk-adjusted rates and ARM margins |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Time Horizon: | Survey Year | 5 years | 10 years | 15 years |
| After-tax PDV: (all in \$) |  |  |  |  |
| No Refi | 5,355 | 9,635 | 18,193 | 26,176 |
| Expected Refi | - | 9,556 | 14,915 | 19,261 |
| Optimal Refi | - | 8,947 | 13,474 | 17,374 |
| \% switching households | 14.3 | 14.3 | 14.3 | 14.3 |
| Notes. Table reports WRTEs, measured as the extra interest (after taxes) + refinancing costs paid by a household choosing an FRM instead of an ARM due to experienced inflation. Original loan amounts are in constant 2000 \$. |  |  |  |  |

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Householder Age and Mobility


Source: CPS ASEC 2005 \& 2010 / authors' calculations.
Fitted values calculated using fourth-order polynomial.

## Table 7: Additional Interest Paid Due to Inflation Experiences

 Scenario 2: Risk-adjusted rates, seniority-adjusted ARM marginsTime Horizon:
E[tenure | age]
After-tax PDV: (all in \$)

| No Refi | 20,819 |
| :--- | :--- |
| Expected Refi | 15,769 |

Optimal Refi 14,475
\% switching households
13.5

Scenario 3: Risk-adjusted rates and ARM margins
Time Horizon:
E[tenure | age]
After-tax PDV: (all in \$)

| No Refi | 19,964 |
| :--- | :--- |
| Expected Refi | 14,854 |
| Optimal Refi | 13,543 |


| Optimal Refi $\frac{13,543}{} 14.3$ |
| :--- |
| Notes. Table reports WRTEs, measured as the extra interest (after taxes) + |
| refinancing costs paid by a household choosing an FRM instead of an ARM due to |
| experienced inflation. Original loan amounts are in constant $2000 \$$. |

## Different Inflation Environments

Our ex post estimates reflect the actual realization of inflation over 1985-2013.

- Are results driven by the Great Moderation?
- Other realizations were possible given initial conditions.


## Simulation parameters

| Variable | Process | $\boldsymbol{\mu} \quad \boldsymbol{\sigma}$ | Source |
| :---: | :---: | :---: | :---: |
| $\pi$ | $\operatorname{AR}(1), \phi=0.811$ | $0.038 \quad 0.027$ | CPI-U, 1960-2013 |
| $r_{1}$ | Indep. WN | 0.020 .022 | Campbell-Cocco (2003) |
| Nominal rates | ST given by Fisher equation LT given by EH + TP | $\theta_{10}=0.01$ | Campbell-Cocco (2003) |
| Mortgage rates | $\begin{aligned} & y_{F R M}=y_{10}+\theta_{F R M} \\ & y_{A R M, 1}=y_{1}+\theta_{A R M, 1} \\ & y_{A R M, 2+}=y_{1}+\theta_{A R M, 2+} \end{aligned}$ | $\begin{aligned} & \theta_{F R M}=0.017 \\ & \theta_{A R M, 1}=0.015 \\ & \theta_{A R M, 2+}=0.0275 \end{aligned}$ | PMMS, 1971-2013 <br> PMMS, 1984-2013 <br> PMMS, 1987-2013 |

## Average Inflation and E[WRTE] in 100 Simulations



## Aggregate Implications

Lifetime experiences of macroeconomic outcomes influence HH decision-making in an economically-significant manner.

- Most effected cohorts: young in the 1980s - Boomers.
- Overweighting lifetime experiences significantly distorted mortgage decisions. Using structural choice estimates and accounting for E[tenure] and E[refi]:
$>$ In late 1980s: 1 million additional $\mathrm{FRMs} \Rightarrow$ ex post overpayment of $\$ 14$ billion.
$>$ Long shadows: in late 1990s, Boomers took out $1 / 2$ million additional $\mathrm{FRMs} \Rightarrow$ ex post overpayment of $\$ 9$ billion.

Welfare implications of overweighting are potentially large.

