

Finance and Synchronization

A. Cesa-Bianchi¹ J. Imbs² J. Saleheen³

¹Bank of England & Centre for Macroeconomics

²Paris School of Economics (CNRS), and CEPR

³Bank of England

ECB/BoE/IMF Workshop

“Global Spillovers: How Much Do We Really Know?”

ECB – April 26, 2016

Disclaimer

The views expressed in this paper are solely those of the authors and should not be taken to represent those of the Bank of England.

This paper

- ▶ We revisit a classic question in international macro: **what is the impact of financial integration on business cycle synchronization?**
- ▶ International synchronization of cycles first-order question
 - Propagation of shocks, external constraint on macro policy, coordination,...
- ▶ Old literature: openness in general main culprit
 - Historically, trade openness [Frankel and Rose (1998); Baxter and Kouparitsas (2005)]
- ▶ Recent events shifted the focus on financial openness. Heuristically, it seems financial linkages helped propagate the great recession

But Theory Ambiguous

- ▶ International Real Business Cycle model [Backus, Kehoe, and Kydland (1992, JPE)]
 - Idiosyncratic productivity shock leads to cross-country MPK differential
 - Because of efficient finance, resources shift where MPK is higher ⇒ **Negatively correlated cycles**
- ▶ IRBC model with credit frictions and integrated financial markets [Allen and Gale (2000), Devereux and Yetman (2010), Dedola and Lombardo (2010)]
 - Idiosyncratic shock (not necessarily to productivity) affects tightness of the constraint at home
 - Because of financial integration, credit constraints are interdependent across countries ⇒ **Positively correlated cycles**
- ▶ Key ingredient is idiosyncratic (ie, country-specific) shock

Common Shocks

- ▶ Shocks common to two or more countries (but with country-specific loadings) constitute a key driver of business cycles
 - Large role of world and regional factor in developed countries (60% for US, 72% for Canada, 72% for France, 56% for Germany) [Kose, Otrok, and Whiteman (2003, 2008), Crucini, Kose, and Otrok (2011)]
- ▶ Empirically important to distinguish country-specific shocks from common shocks with country-specific loadings
- ▶ We will show this is especially important in the literature on cycle synchronization

This paper: results

- ▶ Shows that financial linkages lowers synchronization, conditional on common shocks
- ▶ Shows that financial linkages do not lower synchronization, conditional on idiosyncratic shocks
- ▶ Explains theoretically why common shocks can create such a reversal
- ▶ Since theory builds from idiosyncratic shocks, evidence suggests credit constraints may be relevant empirically

Plan

- ▶ Estimation & Data
- ▶ Results
- ▶ Conclude

Empirics

- ▶ Frankel and Rose (1998, EJ), Imbs (2006, JIE):

$$\rho_{ij} = \alpha + \beta K_{ij} + \delta T_{ij} + \eta_{ij,t}$$

where K_{ij} is a measure of bilateral financial linkages, ρ_{ij} Pearson correlation coefficient

- ▶ Results: $\beta > 0, \delta > 0 \Rightarrow$ **Positively correlated cycles.**
- ▶ But if the true model is

$$\rho_{ij,t} = \alpha_{ij} + \beta K_{ij,t} + \delta T_{ij,t} + \eta_{ij,t}$$

then the between result is fallacious. Need time series to check.
Pearson correlation not adapted (Forbes-Rigobon, 2002)

Empirics (cont'd)

- ▶ Giannone, Lenza and Reichlin (2008), Morgan, Rime and Strahan (2004):

$$\begin{aligned} S_{ij,t} &= -|y_{it} - y_{jt}| \\ S_{ij,t}^e &= -|e_{it} - e_{jt}| \quad \text{where} \quad y_{it} = \alpha_i + \gamma_t + e_{it} \end{aligned}$$

- ▶ Then can estimate

$$S_{ij,t} = \alpha_{ij} + \gamma_t + \beta \cdot K_{ij,t} + \delta \cdot Z_{ij,t} + \eta_{ij,t}$$

- ▶ Kalemli-Ozcan, Papaioannou and Peydro (2013, JF): 18 OECD countries over 1978-2006 $\beta < 0 \Rightarrow$ **Negative correlated cycles**
- ▶ But common shocks pollute this estimation – if they have country-specific loadings

Why common shocks matter

- ▶ Suppose true model is:

$$y_{i,t} = a_i^y + b_i^y \mathcal{F}_t^y + \varepsilon_{i,t}^y$$

where \mathcal{F}_t is a vector of common factors

- ▶ Then \mathcal{S}_{ijt} embeds heterogeneous responses to common shocks:

$$\mathcal{S}_{ij,t} = - \left| a_i^y - a_j^y + (b_i^y - b_j^y) \mathcal{F}_t^y + \varepsilon_{i,t}^y - \varepsilon_{j,t}^y \right|$$

- ▶ True of both $\mathcal{S}_{ij,t}$ and $\mathcal{S}_{ij,t}^e$.

Data: Sample

- ▶ Data is extension of Kalemli-Ozcan, Papaioannou, Peydro (JF, 2013)
- ▶ KPP data set covers 18 advanced economies
 - Australia, Austria, Belgium, Canada, Switzerland, Germany, Denmark, Spain, Finland, France, UK, Ireland, Italy, Japan, Netherlands, Portugal, Sweden, and US
 - 153 country pairs
- ▶ Annual data from 1980 to 2012

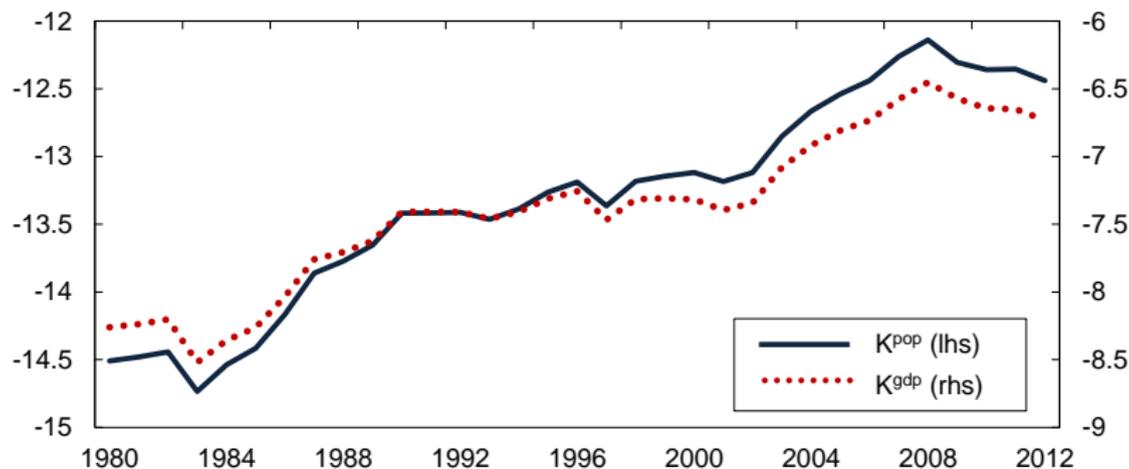
Banking integration measures

- ▶ Virtually non existent time varying measures of international capital but for bank assets and liabilities
- ▶ “International Locational Banking Statistics Database” provided by the BIS
 - Asset (A_{ij}) and liability (L_{ij}) of banks located in i (the “reporting area”) held in country j (the “vis-a-vis area”)
- ▶ Two measures: normalized by population or by GDP

$$BANKINT1_{ij,t} = \frac{1}{4} \left[\ln \left(\frac{A_{ij,t}}{P_i + P_j} \right) + \ln \left(\frac{L_{ij,t}}{P_i + P_j} \right) + \ln \left(\frac{A_{ji,t}}{P_i + P_j} \right) + \ln \left(\frac{L_{ji,t}}{P_i + P_j} \right) \right]$$

$$BANKINT2_{ij,t} = \frac{1}{4} \left[\ln \left(\frac{A_{ij,t}}{Y_i + Y_j} \right) + \ln \left(\frac{L_{ij,t}}{Y_i + Y_j} \right) + \ln \left(\frac{A_{ji,t}}{Y_i + Y_j} \right) + \ln \left(\frac{L_{ji,t}}{Y_i + Y_j} \right) \right]$$

Banking integration measures



The solid and dotted lines plot the evolution over time of the average value of $K_{ij,t}^{pop}$ and $K_{ij,t}^{gdp}$ for the 1980-2012 period. The average is computed across 153 country pairs (our sample spans 18 countries) for each year.

Synchronization measures we consider

- ▶ Consider the following measures

$$S_{ij,t} = - \left| a_i^y - a_j^y + (b_i^y - b_j^y) \mathcal{F}_t^y + \varepsilon_{i,t}^y - \varepsilon_{j,t}^y \right|$$

$$S_{ij,t}^{\mathcal{F}} = - \left| (b_i^y - b_j^y) \mathcal{F}_t^y \right|$$

$$S_{ij,t}^{\varepsilon} = - \left| \varepsilon_{i,t}^y - \varepsilon_{j,t}^y \right|$$

- ▶ $S_{ij,t}^{\mathcal{F}}$ and $S_{ij,t}^{\varepsilon}$ are the components of $S_{ij,t}$ associated with common and idiosyncratic shocks, respectively
- ▶ Use either measure in conventional panel regression

$$S_{ij,t} = \alpha_{ij} + \gamma_t + \beta \cdot K_{ij,t} + \delta \cdot Z_{ij,t} + \eta_{ij,t}$$

OLS between, then OLS within. With or without trade controls.

How to proxy for unobserved common factors?

- ▶ Objective: compute *country-specific* decompositions of the type

$$y_{i,t} = a_i^y + b_{1,i}^y \mathcal{F}_{1,t}^y + \dots + b_{n,i}^y \mathcal{F}_{n,t}^y + \nu_{it}^y,$$

- ▶ Simple methodology: extract the first n principal components (\mathcal{F}_t^n) from the panel (28 years \times 18 countries) of GDP growth rates
- ▶ How many principal components? Retain principal components as long as their associated eigenvalue is > 1

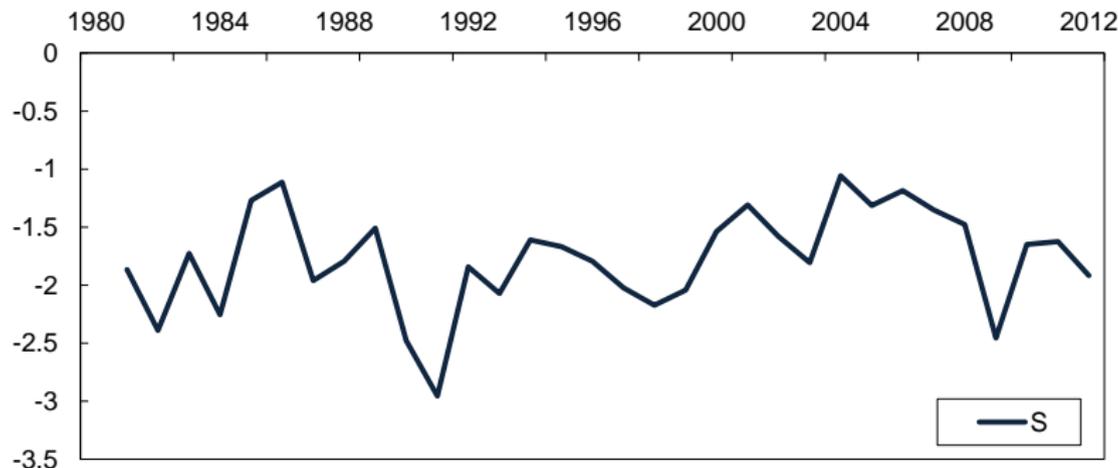
How to proxy for unobserved common factors?

Factor estimates

	Eigenvalues		Share of variance		Cum. share of variance	
	y_{it}	$K_{ij,t}$	y_{it}	$K_{ij,t}$	y_{it}	$K_{ij,t}$
\mathcal{F}_1	10.67	13.15	59%	73%	59%	73%
\mathcal{F}_2	2.21	2.89	12%	16%	72%	89%
\mathcal{F}_3	1.02	0.79	6%	4%	77%	93%
\mathcal{F}_4	0.89	0.66	5%	4%	82%	97%
\mathcal{F}_5	0.83	0.27	5%	2%	87%	99%

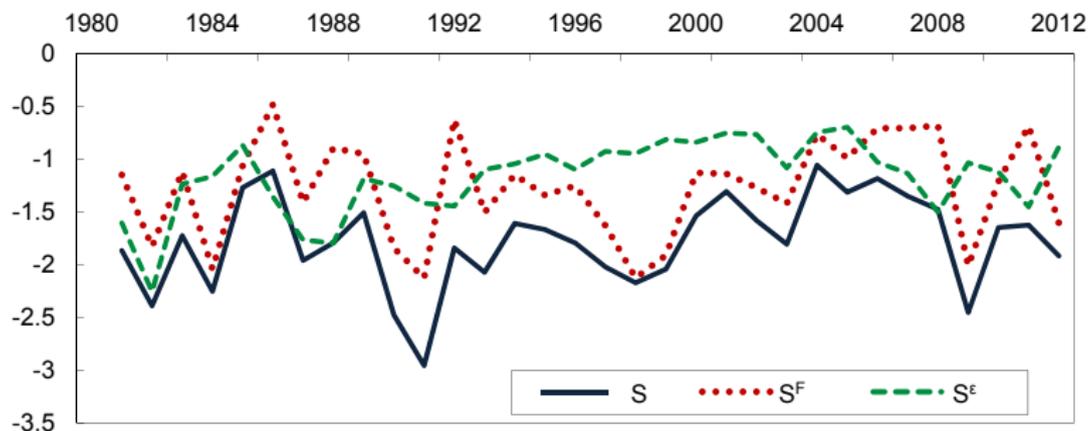
- ▶ Compute $S_{ij,t}^{\mathcal{F}}$ using fitted values $\hat{a}_i + \hat{b}_{1,i}^y \hat{\mathcal{F}}_{1,t}^y + \hat{b}_{2,i}^y \hat{\mathcal{F}}_{2,t}^y + \hat{b}_{3,i}^y \hat{\mathcal{F}}_{3,t}^y$
- ▶ Compute $S_{ij,t}^{\varepsilon}$ using residuals ν_{it}^y

How do our synchronization measures look like?



The solid line plots the evolution over time of the average value of $S_{ij,t}$ for the 1980-2012 period. The average is computed across 153 country pairs (our sample spans 18 countries) for each year. The chart also reports the cross-sectional averages of the idiosyncratic component (dashed line) and the common component (dotted line) of $S_{ij,t}$. F_t has been proxied by the first 3 principal components on the full panel of GDP growth rates. The averages are computed across 153 country pairs for each year over the 1980-2012 period.

How do our synchronization measures look like?



The solid line plots the evolution over time of the average value of $S_{ij,t}$ for the 1980-2012 period. The average is computed across 153 country pairs (our sample spans 18 countries) for each year. The chart also reports the cross-sectional averages of the idiosyncratic component (dashed line) and the common component (dotted line) of $S_{ij,t}$. F_t has been proxied by the first 3 principal components on the full panel of GDP growth rates. The averages are computed across 153 country pairs for each year over the 1980-2012 period.

Plan

- ▶ Estimation & Data
- ▶ **Results**
- ▶ Conclude

OLS “between” estimates (1980–2012)

	S	$S^{\mathcal{F}}$	S^{ε}	S	$S^{\mathcal{F}}$	S^{ε}
	(1)	(2)	(3)	(4)	(5)	(6)
Banking / Pop. (K^{pop})	0.095 (0.011) [8.70]	0.106 (0.008) [12.85]	0.038 (0.007) [5.43]			
Banking / GDP (K^{gdp})				0.091 (0.010) [9.52]	0.082 (0.007) [11.31]	0.049 (0.006) [7.98]
Observations	4863	4863	4863	4863	4863	4863
R^2	0.092	0.176	0.121	0.095	0.170	0.127
Country Pairs	153	153	153	153	153	153

All regression specifications include a vector of year fixed effects. Estimation is performed over the 1980-2012 period.

OLS “within” estimates (1980–2012)

	S	$S^{\mathcal{F}}$	S^{ε}	S	$S^{\mathcal{F}}$	S^{ε}
	(1)	(2)	(3)	(4)	(5)	(6)
Banking / Pop. (K^{pop})	-0.144 (0.040) [-3.63]	-0.154 (0.030) [-5.05]	0.075 (0.021) [3.54]			
Banking / GDP (K^{gdp})				-0.148 (0.042) [-3.56]	-0.159 (0.032) [-4.98]	0.072 (0.022) [3.28]
Observations	4863	4863	4863	4863	4863	4863
R^2	0.099	0.222	0.133	0.099	0.222	0.133
Country Pairs	153	153	153	153	153	153

All regression specifications include a vector of country-pair fixed effects and a vector of year fixed effects. Estimation is performed over the 1980-2012 period. Standard errors are adjusted for country-pair-level heteroskedasticity and autocorrelation.

OLS “within” estimates with controls (1980–2012)

	S	$S^{\mathcal{F}}$	S^{ε}	S	$S^{\mathcal{F}}$	S^{ε}
	(1)	(2)	(3)	(4)	(5)	(6)
Banking / Pop. (K^{pop})	-0.102 (0.040) [-2.57]	-0.132 (0.028) [-4.71]	0.060 (0.024) [2.55]			
Banking / GDP (K^{gdp})				-0.106 (0.041) [-2.55]	-0.137 (0.029) [-4.65]	0.056 (0.024) [2.32]
Trade	-0.382 (0.134) [-2.86]	-0.198 (0.114) [-1.75]	0.132 (0.078) [1.69]	-0.386 (0.133) [-2.90]	-0.203 (0.113) [-1.79]	0.141 (0.078) [1.81]
Observations	4859	4859	4859	4859	4859	4859
R^2	0.103	0.224	0.134	0.103	0.225	0.134
Country Pairs	153	153	153	153	153	153

All regression specifications include a vector of country-pair fixed effects and a vector of year fixed effects. Estimation is performed over the 1980-2012 period. Standard errors are adjusted for country-pair-level heteroskedasticity and autocorrelation.

Summing up

- ▶ Sign of β is negative for $\mathcal{S}_{ij,t}$ and $\mathcal{S}_{ij,t}^{\mathcal{F}}$, positive for $\mathcal{S}_{ij,t}^{\mathcal{E}}$
 - Show that reversal is due to permanent country features
- ▶ Robustness
 1. Focus on tranquil times (1980-2006)
 - ▶ β less significant but still positive
 2. Deal with endogeneity: bilateral, time-varying IV using the instrument of KPP:
 - ▶ β positive and significant
 3. Pairwise correlations
 - ▶ β is not statistically different from zero
 4. Time-varying factor loadings
 - ▶ β positive and significant

This paper: results

- ▶ Shows that financial linkages lowers synchronization, conditional on common shocks
- ▶ Shows that financial linkages do not lower synchronization, conditional on idiosyncratic shocks
- ▶ Explains theoretically why common shocks can create such a reversal
- ▶ Since theory builds from idiosyncratic shocks, evidence suggests credit constraints may be relevant empirically

Appendix

Why common factors matter? Analytical results

- ▶ Assume $K_{ij,t}$ also responds to both common and idiosyncratic shocks, i.e.:

$$K_{ij,t} = a_{ij}^K + b_{ij}^K \mathcal{F}_t^K + \varepsilon_{ij,t}^K$$

- Forbes and Warnock (2012), Rey (2013), Bruno and Shin (2014)
- Common shock term is general. Accounts for global cycles in financial linkages, or a potential trend in $K_{ij,t}$

- ▶ Consider our baseline regression with $\mathcal{S}_{ij,t}^{\mathcal{F}}$

$$- \left| (b_i^y - b_j^y) \mathcal{F}_t^y \right| = \alpha_{ij} + \gamma_t + \beta^{\mathcal{F}} \cdot [a_{ij}^K + b_{ij}^K \mathcal{F}_t^K + \varepsilon_{ij,t}^K] + \delta \cdot Z_{ij,t} + \eta_{ij,t}^{\mathcal{F}}$$

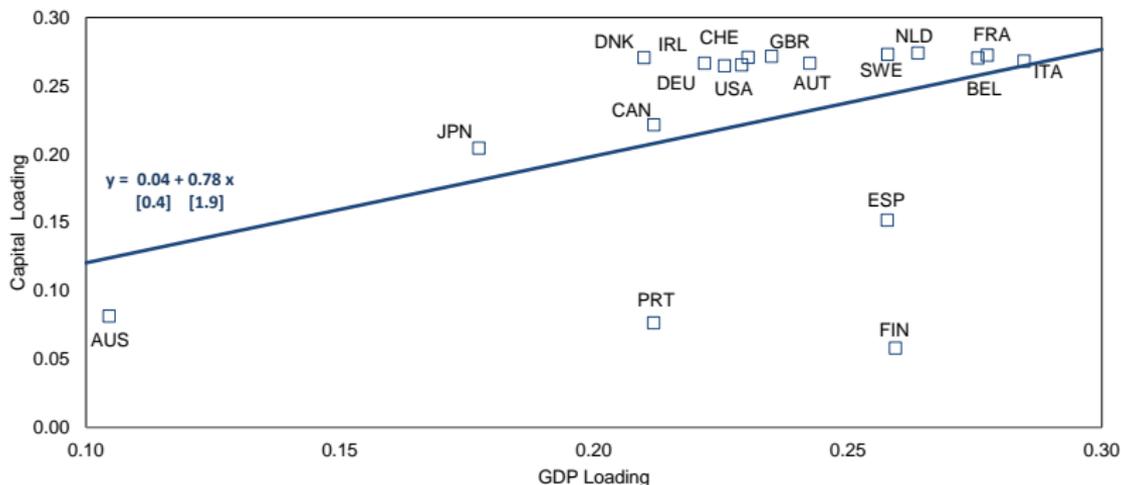
- ▶ Sign of $\beta^{\mathcal{F}}$ is given by:

$$- \left| b_i^y - b_j^y \right| \cdot b_{ij}^K \text{Cov} [|\mathcal{F}_t^y|, \mathcal{F}_t^K]$$

Permanent Features?

- ▶ Suppose now a systematic positive correlation exists between $(b_i^y - b_j^y)$ and b_{ij}^K
 - E.g., in response to common shocks, capital intensity between i and j is larger in country pairs with large differences in GDP elasticity (high $b_i^y - b_j^y$)
- ▶ Then a negative correlation exists between $S_{ij,t}^{\mathcal{F}}$ and $K_{ij,t}$, and it is driven by permanent features of GDP and capital flows
- ▶ Empirical question: do high b_i countries also display high b_{ij}^K ?
- ▶ Plot estimates of $\hat{b}_{1,i}$ against $\hat{b}_{1,i}^K$, where $\hat{b}_{1,i}^K = \frac{1}{J} \sum_j \hat{b}_{1,ij}^K$

Permanent Features?



On the horizontal axis is the loading on GDP ($\hat{b}_{1,i}^y$). On the vertical axis is the loading on capital $\hat{b}_{1,i}^K$, where $\hat{b}_{1,i}^K = \frac{1}{J} \sum_j \hat{b}_{1,i,j}^K$ is the average capital loading in country i . The slope and the constant of the fitted line are reported together with t-Statistics in square brackets.

Permanent Features?

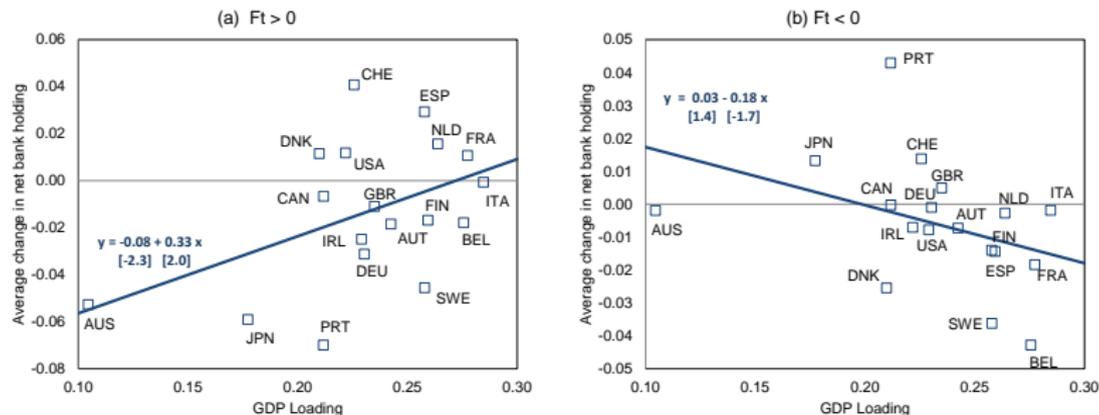
- ▶ Also implies that capital ($\hat{K}_{ij,t} = \hat{a}_{ij}^K + \hat{b}_{ij}^K \mathcal{F}_t$) should go TO countries with elastic GDP in periods of global (or regional) booms ($\mathcal{F}_t > 0$), but FROM them in years of global recession ($\mathcal{F}_t < 0$).
- ▶ Define the average change in net bank holdings, computed for positive or negative values of \mathcal{F}_t :

$$KNET_i^+ = \sum_{\mathcal{F}_t > 0} \Delta_t \left[\sum_j \ln(A_{ji,t} + L_{ij,t}) - \ln(A_{ij,t} + L_{ji,t}) \right],$$

and:

$$KNET_i^- = \sum_{\mathcal{F}_t < 0} \Delta_t \left[\sum_j \ln(A_{ji,t} + L_{ij,t}) - \ln(A_{ij,t} + L_{ji,t}) \right]$$

Permanent Features?



On the horizontal axis is the loading on GDP ($\hat{b}_{1,i}^y$). On the vertical axis is the change in net bank holdings averaged over periods when $F_t > 0$ ($KNET_i^+$), in panel (a); and when $F_t < 0$ ($KNET_i^-$), in panel (b). The slope and the constant of the fitted line are reported together with t-Statistics in square brackets.